

**WHO DO UNIONS TARGET? UNIONIZATION OVER THE LIFE-CYCLE OF U.S.
BUSINESSES***

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Abstract

What type of businesses do unions target for organizing and when? A dynamic model of the union organizing process is constructed to answer this question. A union monitors establishments in an industry to learn about their productivity, and decides which ones to organize and when. An establishment becomes unionized if the union targets it for organizing and wins the union certification election. The model predicts two main selection effects: unions target larger and more productive establishments early in their life-cycles, and among the establishments targeted, unions are more likely to win elections in smaller and less productive ones. These predictions find support in union certification elections data for 1977-2007 matched with data on establishment characteristics.

Keyword: Unionization, Union Organizing, Union Certification Election, Diffusion of Unionization, Bayesian Learning, Productivity.

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*Any opinions and conclusions expressed herein are those of the authors and do not necessarily represent the views of the U.S. Census Bureau. All results have been reviewed to ensure that no confidential information is disclosed.

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1 Introduction

Unions have been influential in the U.S. economy and politics for much of the 20th century. They continue to be so even though private sector unionization has declined persistently over the last four decades.¹ Despite the long tradition of union activity in the U.S., there is still little information about the timing and nature of union formation in a business and the diffusion of unionization in an industry. In particular, the union selection process is relatively unexplored. What type of businesses do unions target for organizing? Do they typically target big and profitable business establishments that can provide larger employment and benefits to the union? Such establishments may be harder for unions to organize, because they in general have higher wages, as well as greater resources and better management to resist unionization. Alternatively, unions may focus on smaller establishments that are easier to organize because of poor labor conditions and weak management, which may be correlated with lower profitability. These type of establishments also offer lower wages in general, potentially implying a higher demand for unionization. Furthermore, not much is known about the timing of union activity in an establishment. When in an establishment's life-cycle does a union try to organize it? Do unions organize young businesses or wait until a business is more established? A lack of comprehensive panel data on establishment-level union activity in the U.S. has precluded definitive answers to these questions. The few empirical studies on aspects of the unionization process usually focus on newly certified establishments in datasets that include unions alone, with little detail on establishment characteristics and life-cycles.² As a consequence, the establishment-level dynamics of union activity have remained largely undocumented.

This paper examines unionization at the establishment level for the period 1977-2007 by combining the entire National Labor Relations Board (NLRB) union election data with data on all private-sector establishments from the U.S. Census Bureau. This newly constructed establishment-level panel makes it possible to relate establishment characteristics, such as size, age, and productivity, to union activity over the life-cycle of an establishment in the form of certification and decertification elections, and to the outcomes of these elections. The data is

¹In 2012, only 6.6% of private sector workers in the U.S. were union members, compared to 24.2% in 1973 – see the *Union Membership and Coverage Database* at www.unionstats.com and the associated documentation in Hirsch and MacPherson (2003).

²See, e.g., Holmes (2006), Farber and Western (2001), Farber (2013), and Ferguson (2008).

analyzed from the perspective of a dynamic model of union organizing. The model highlights union learning about an establishment's underlying productivity as a potential mechanism in determining the type of establishments that are targeted by unions and the timing of such targeting. The unionization process is formulated in a variant of Jovanovic's (1982) model of industry dynamics. An establishment enters the industry with a prior about its unknown underlying productivity. After entry, each establishment experiences random shocks to its productivity over time. More productive establishments tend to be larger, and generate higher profits. There is also a union in the industry, whose aim is to maximize the expected life-time benefits from organizing labor. It monitors establishments to learn about their productivity in a Bayesian fashion over time. Unionization is a costly and uncertain process. The cost and uncertainty of unionization limit the number and type of establishments the union can organize. Which establishments in the industry should the union organize and when?

The model predicts that unions target large and productive establishments early on in their life-cycles. Among establishments of a given age, larger and more productive ones are more likely to be targeted. For any given size or productivity, younger establishments are more likely to experience a certification election, and the likelihood of an establishment being targeted for the first time by a union declines with age. Similar predictions apply to the likelihood of the event that an establishment is organized by a union successfully for the first time – successful organizing occurs when a union targets an establishment for organizing *and* wins the certification election. Furthermore, the probability of a union win in an election, conditional on an establishment being targeted for organizing, is lower for larger and more productive establishments. The model also suggests that, in a cross section of establishments, the unionized ones (those that have experienced in their life-time at least one certification election won by a union) are on average larger, older, and more productive than the rest.

The predictions of the model are taken to the data on union activity in the entire set of U.S. private-sector establishments born between 1977 and 2007.³ During this period, the number of union certification elections in a year declined from around 9,000 in 1977 to about

³In the U.S., union certification elections occur mainly in occupational groups at the establishment level. The empirical work focuses on establishments as the unit of analysis, as the theoretical model considers the first time a union targets an establishment and wins an election, regardless of the occupational group that is the subject of the organizing.

2,000 in 2007. Only about 0.03% of all establishments born during the sample period were targeted for the first time by a union for potential organizing. Unions tended to win around 47% of certification elections in a year, on average, throughout this period, though the win rate increased from about 50% to nearly 60% between 2000 and 2007. The likelihood of a union successfully organizing an establishment for the first time was around 0.015%. As of 2007, around 0.2% of all surviving establishments born between 1977 and 2007 were still unionized in the sense that they experienced in their life-time at least one certification election that resulted in union victory, and no subsequent decertification election.⁴

The empirical analysis proceeds in a flexible way without imposing the exact form and assumptions of the model on the data. Based on the predictions of the model, four probabilities of interest are explored: the probability of an establishment being targeted for the first time by a union for potential organizing, the probability of a union win in the first ever certification election held in an establishment, the probability that a union successfully organizes an establishment the first time, and the probability that an establishment is unionized. Each of these probabilities is formulated using a logistic distribution, where the probability is related to establishment size, productivity, age and other controls following the model's implications. The data supports the model's predictions that unions target large and productive establishments for organizing. For the entire private sector, the likelihood of an establishment with more than 500 employees being targeted by a union is about 10 times larger than an establishment with less than 10 employees, holding all else constant. This relative likelihood is as high as 23 when only the manufacturing sector is considered. Furthermore, unions do not wait too long to target a large and productive establishment after it is born. Conditional on size and other observables, the likelihood of an establishment being targeted by a union for the first time is highest around the time of its birth, and declines steadily until about 10 to 12 years after entry, remaining relatively flat thereafter. The youngest group of establishments (0-3 years old) are approximately twice as likely to be targeted as the oldest group (25+ years old).

⁴The union status of an establishment is not observed directly, only the outcomes of certification and decertification elections. Throughout the empirical analysis, an establishment is labelled as "unionized" if it has experienced at least one certification election won by a union and no subsequent decertification of that union. There can be instances where a certification election occurs but the establishment is not unionized or where there is a union without a prior certification election.

Within the set of establishments that are targeted, unions are less likely to win certification elections in larger and more productive ones, as the model suggests. The probability of a union win in a certification election declines as establishment size increases: all else equal, the probability of win is about 30% for the largest employment category (500+ employees), and about 60% for the smallest category (1-9 employees). Overall, successful union organizing (union targeting *and* union win) is more likely to occur in larger and more productive establishments. Establishments in the largest employment category are 12 times more likely to be successfully organized by a union compared with the smallest category. Moreover, at any point in time, union establishments tend to be more productive, larger, and older compared with non-union ones, as predicted by the model. Establishments in the largest employment class are about 11 times more likely to be unionized, compared with the ones in the smallest one. In general, size effects tend to be much higher in magnitude and statistically more significant than productivity effects. All effects appear to be more pronounced in the manufacturing sector than for the private sector as a whole.

Certain other characteristics of establishments that are observable by unions are also significantly associated with the likelihood of being targeted. For instance, being part of a multi-unit firm and having at least one sister establishment that is already unionized increase the likelihood than a union targets an establishment. These effects also give support to the hypothesized learning process behind unionization, as unions can use all of these additional signals to learn more about the eligibility of an establishment for successful organizing. In practice, unions indeed seem to rely on such signals. For instance, Figure 1 shows the information solicited by the United Automobile Workers (UAW) union in online petitions for union organizing by workers in an establishment. In addition to requesting a “best estimate” of the establishment’s total employment, the UAW is interested in obtaining information on the identity of the parent firm, whether the firm has multiple establishments, and the union status of other establishments within the same firm.

The findings are relevant for a large body of literature on the effects of unionization on firm-level outcomes including employment, wages, productivity, exit, and stock market value.⁵ Many of these studies do not consider the process by which establishments are selected for

⁵See, e.g., Freeman and Medoff (1984) for a comprehensive discussion of such studies.

union organizing.⁶ Most studies are based on relatively small samples of large or publicly traded firms that stay in business over a narrow window of time before and after unionization. With some exceptions, these studies are therefore subject to a combination of survival, union selection, size, and public-status bias. This study avoids such biases by focusing on the entire set of private sector establishments in the U.S., and by following the establishments from their birth onwards, not just over a period before and after unionization. The establishments that are targeted and successfully organized by unions differ systematically from the non-targeted. Two distinct selection effects emerge. The first is the selection of larger and more productive establishments by unions as targets for potential organizing. The second pertains to the result of a certification election. In establishments that experience a certification election for the first time in their life-cycle, unions are more likely to win the election among smaller and less productive ones.⁷ There may be many reasons for this effect. The model suggests that the payoff from organizing a large, productive establishment is higher. Thus, a union targets such an establishment even when the likelihood of a win in the certification election is lower. Furthermore, a large or productive establishment may be inherently less prone to losing an election due to better management and organizational practices, which may be correlated with high productivity or large size. However, the negative effect of size and productivity on union win likelihood does not in general overwhelm the positive relation between union targeting activity and establishment size or productivity. As a result, unions are more likely to successfully organize establishments that are on average larger and more productive. These selection biases need to be taken into account in assessing the effects of unionization on business outcomes. To the extent that post-unionization outcomes depend on establishment characteristics at the time of unionization that persist into future, the selection based on these characteristics (e.g. size, productivity and age) is important.

The age effects found in this study also offer further insight to the unionization process. Conditional on observable characteristics, establishments are most likely to experience their first

⁶Addison and Hirsch (1989) point to the selection issues involved in union studies and state (p. 83) that no comprehensive study has been done on these issues.

⁷Dinardo and Lee (2004) also discuss these two potential selection effects in union organizing. They do not study the unionization process, but rather the effects of unionization on outcomes such as wages and exit likelihood. Their approach is to look at cases where unions barely won certifications elections versus cases where they barely lost. Lee and Mas (2012) and Frandsen (2013) follow a similar approach.

certification election and a union victory within the first couple of years after their entry. The likelihood of an establishment being targeted by a union and a subsequent union victory taper off as new establishments age. While this pattern is consistent with the learning model proposed here, there may be alternative explanations, such as lack of managerial experience with union organizing tactics in young businesses. The age effects also suggest a cautionary note for future efforts in using NLRB union certification election data in conjunction with establishment-level data. The sample selection techniques used in many prior studies often de-emphasize or exclude from the analysis establishments that did not exist prior to the election, or those young establishments that existed only a few years before the election. Such establishments appear to be precisely the ones that are most likely to experience a certification election.

The selection effect in union organizing documented in this study have important implications for research on the effects of unions. Even if unions are able to extract some rents from businesses, the selection effect implies that unions may not have a discernible effect on certain outcomes, such as survival, in large and productive establishments. For example, Dinardo and Lee (2004) find little effect of unions on wages, employment, and survival in relatively large manufacturing establishments, a result that may be driven by the selection effects documented here. These large establishments may be able to better withstand any potential adverse effects of unions on profits and survival. Smaller or less productive establishments, however, may experience more significant effects. Future work can assess this potentially heterogeneous effect of unionization on establishment-level outcomes. In general, unions may influence the type of technologies and work practices that are implemented in establishments, as well as the flexibility of hiring and wage setting.⁸ These effects may take time to materialize and may not always show up in available measures of productivity. Because unions target more productive and larger establishments, the welfare costs of unions may be higher than the ones traditionally estimated by Rees (1963), which arise solely from high union wages.⁹ Recent work also argues that unions played an important role in the decline of manufacturing in the U.S. Rust Belt.¹⁰ This role is potentially amplified when unions are concentrated in larger and more productive manufactur-

⁸See, e.g., Schmitz (2005).

⁹Additionally, just the threat of unionization may effect a firm's behavior. To prevent unionization, firms may have to offer high wages and more benefits than in a world without the threat of unionization. See Taschereau-Dumouchel (2012) for an analysis of the threat effect at the macro level.

¹⁰See Alder, Lagakos, and Ohanian (2013).

ing plants. If there are indeed adverse effects of unions on businesses, unions' concentration in large and productive plants may have accelerated the decline of manufacturing in U.S. under increased competition and foreign trade. Conversely, as documented by Holmes (2011), the decline of manufacturing due to competitive pressure, trade, and technological change led to the gradual disappearance of large manufacturing plants. These plants are what unions tend to target, as the findings here indicate. Their decline may have therefore reinforced the decline of unions by depriving them of lucrative organizing opportunities.

The findings also provide guidance on the modeling of the unionization process and on the diffusion of unionization in a population of heterogeneous firms. Models that recognize the selection of unions into larger and more productive businesses are rare.¹¹ Studies that aim to model unions as integral parts of the economy can benefit from the empirical regularities pertaining to the unionization process provided in this paper.¹² The union learning process analyzed in the model can be also generalized to include union's learning from organizing other establishments within an industry or within the same geographic area. For instance, Holmes (2006) finds that unionism exhibits geographic spillover, consistent with a process by which workers in a non-union business learn from the experience of unionized businesses located nearby.

The rest of the paper is organized as follows. The next section presents the model and derives its testable implications. Section 3 describes the empirical methodology, followed by the description of the data in Section 4. Section 5 presents the findings. Section 6 makes concluding remarks. Appendix A contains all proofs, and Appendix B provides additional material for the model. Appendix C describes the data. Robustness analysis and additional results are collected in Appendix D.

¹¹For example, Dinlersoz and Greenwood (2013) posit a framework where unions organize the most productive (and largest) firms because such firms are able to afford both higher employment and higher wages for union members.

¹²Such studies include Acikgoz and Kaymak (2012), Acemoglu, Aghion and Violante (2001), and Krusell and Rudanko (2013).

2 A Model of Union Learning and Organizing

Consider an infinitely-lived industry inhabited by establishments that differ in total factor productivity. Profits and measures of size (e.g. employment, output, or revenue) for an establishment are increasing functions of productivity, as is the case under standard production functions. There is a single industry-wide union whose aim is to organize establishments, each one separately.¹³ The union's per-period benefit from organizing an establishment is taken to be a time-invariant function of the establishment's productivity. Establishments draw their productivity from a stationary distribution, implying that the union targeting decision does not depend on calendar time. Each establishment is born non-unionized, but can become unionized as early as its first period. For simplicity, unionization is an irreversible event, and each establishment can become unionized only once.¹⁴

Neither the union nor an establishment knows the establishment's underlying productivity, but both have prior beliefs on its distribution. Both parties update their priors over time as they obtain more information about the productivity. For simplicity, they observe the same information every period. Their learning processes are therefore identical.¹⁵ Based on the learning process, the union decides whether and when to target an establishment for organizing. There is, however, a cost of organizing, and the outcome of a certification election is random.¹⁶ The organizing cost and the uncertainty about the election result limit the number and type of establishments that the union targets, and hence, the diffusion of unionization in the industry.

The timing of events and decisions for the union within a period is shown in Figure 2. Time is discrete. The union enters a period with prior beliefs about the productivity of each

¹³Alternatively, workers in an establishment may decide to get organized independently of other establishments, with or without the help of an industry-wide union. The predictions of the model are similar under such an interpretation.

¹⁴Once it is certified in an establishment, the decertification of a union is very rare. Only about 1% of unionized establishments experience a decertification election. The most common way for a union to dissolve at an establishment is through establishment exit. Furthermore, multiple certification elections in an establishment are rare (less than 5% of establishments experience multiple elections).

¹⁵It is possible to let the union and establishment observe different signals of profitability and have different learning processes. Such asymmetry is not the focus of the paper.

¹⁶Union organizing costs may include costs associated with monitoring an establishment's performance, penetrating and educating its labor force (e.g. planting union agents in labor force), campaigning to collect signatures for a certification election, and countering an employer's campaign against unionization.

establishment and with knowledge about the cost of organizing the establishment and the likelihood of winning the certification election. At the beginning of the period, the union decides which establishments to target *before* establishments realize their current period productivity. Certification elections are then held. Unionization occurs when the union wins the election. The current period productivity of each establishment is then observed, and production takes place. At this time, the union realizes the benefits. At the end of the period, the union and each establishment update their beliefs about the latter's long-run productivity, and next period's cost of organizing for each non-unionized establishment is revealed.

2.1 The Productivity Process

Let x_a denote (the logarithm of) total factor productivity for an age- a establishment. For $a \geq 1$, the variable x_a follows the process [see Jovanovic (1982)]

$$x_a = \chi + \varepsilon_a, \tag{1}$$

where $\varepsilon_a \sim N(0, \sigma_\varepsilon^2)$ is white noise. Both χ and ε_a are unknown by the union and establishment. The distribution for ε_a is known.

Just after an establishment's entry, χ and ε_1 are drawn. The parameter χ comes from the distribution $N(\bar{\chi}, \sigma_\chi^2)$, and it is fixed for the rest of the establishment's life. While the establishment does not know χ , it knows its distribution. Upon drawing ε_1 and χ , the establishment learns x_1 . It can start first-period production then. Thereafter, the establishment's x_a fluctuates around its average, χ . Both the establishment and the union learn about χ over time based on the information in the realized values of the x_a 's. This information is contaminated by random shocks, the ε_a 's.

2.2 The Learning Process

Suppose the union is monitoring an age- a establishment at the beginning of some period for potential organizing. The union has a prior belief about the establishment's χ . This prior is normally distributed, with mean and variance denoted by ζ_{a-1} and $\sigma_{\zeta_{a-1}}^2$, respectively. The establishment draws a new value, x_a , observed by both the union and the establishment. Using

Bayes' Rule, the union then obtains a posterior distribution for χ with mean

$$\zeta_a = \theta_a \zeta_{a-1} + (1 - \theta_a)x_a, \quad (2)$$

and variance

$$\sigma_{\zeta_a}^2 = \frac{1}{\sigma_\varepsilon^{-2} + \sigma_{\zeta_{a-1}}^{-2}}, \quad (3)$$

where

$$\theta_a \equiv \frac{\sigma_\varepsilon^2}{\sigma_{\zeta_{a-1}}^2 + \sigma_\varepsilon^2},$$

for $a \geq 1$, and

$$\zeta_0 = \bar{\chi} \text{ and } \sigma_{\zeta_0}^2 = \sigma_\chi^2.$$

Now, consider the prior beliefs of the union about x_a . Because χ and ε_a are both normally distributed, (1) implies that the prior distribution of x_a is normal with a mean denoted by μ_{a-1} and variance represented by σ_{a-1}^2 . Taking the expectation of (1) yields

$$\mu_{a-1} = \zeta_{a-1}. \quad (4)$$

The variance σ_{a-1}^2 is given by

$$\sigma_{a-1}^2 = \sigma_{\zeta_{a-1}}^2 + \sigma_\varepsilon^2. \quad (5)$$

Using (2) and (4), one can write the law of motion for μ_a as

$$\begin{aligned} \mu_a &= \theta_a \mu_{a-1} + (1 - \theta_a)x_a \\ &= (1 - \theta_a)\chi + \theta_a \mu_{a-1} + (1 - \theta_a)\varepsilon_a, \end{aligned} \quad (6)$$

where the initial prior, $\mu_0 = \bar{\chi}$, is the same for all new establishments.¹⁷

Let $\Phi(x_a; \mu_{a-1}, \sigma_{a-1}^2)$ be the (normal) cumulative distribution function (c.d.f.) of x_a . Note, from (3) and (5), that σ_{a-1}^2 changes over time only because a changes, since σ_ε^2 is known. Therefore, Φ can be summarized by the pair (μ_{a-1}, a) .

¹⁷Note that μ_a follows an AR(1) process. Equation (6) can also be written as

$$\mu_a = \mu_{a-1} + (1 - \theta_a)(x_a - \mu_{a-1}), \text{ for } a \geq 1,$$

Because $1 - \theta_a > 0$, values of x_a higher (lower) than μ_{a-1} lead to an upward (downward) revision of the prior.

2.3 The Union's Problem

The payoff to the union from organizing an establishment is represented by a union benefit function, $B(x_a) > 0$, which gives the period surplus the union obtains from a unionized establishment with current productivity x_a . The benefit function $B(x_a)$ satisfies the following assumption.

Assumption 1 $B(x_a)$ is bounded, strictly increasing, and strictly convex.

Assumption 1 states that the union obtains an increasingly larger surplus as the productivity of a unionized establishment increases. A higher level of x_a would generally imply a larger, more profitable establishment.¹⁸ Assumption 1 is not arbitrary, and has its foundations in the literature on union-establishment conduct. A union benefit function $B(x_a)$ satisfying Assumption 1 can be obtained in a variety of models governing the relationship between the union and an establishment, including monopoly union, right-to-manage, and efficient bargaining models.¹⁹ The exact mode of the post-unionization behavior of the establishment and the union is therefore not specified. Appendix B gives the derivation of $B(x_a)$ under the above three models.

There is a cost $c > 0$ of organizing an establishment. This cost is known by the union and is incurred regardless of the outcome of the certification election.²⁰ The union wins a certification election with probability ω_a in an age- a establishment. The probability ω_a is an independently

¹⁸For example, imagine an establishment whose production function is given by $\exp(x_a)l^\alpha$, where l is employment and $\exp(x_a)$ is total factor productivity. The production function is a standard one that is frequently used. The form used for total factor productivity is typical when shocks are normal. If the establishment is in a competitive industry and could freely hire labor at the wage rate, w , then its employment would be given by $l(x_a) = [\exp(x_a)/w]^{1/(1-\alpha)}$, which is strictly convex in x_a . Output and profits are also strictly convex in x_a .

¹⁹See, e.g., Manning (1987, 1994) for a discussion of these different models. In the case of a monopoly union, the union picks the wage while the establishment chooses employment given the union wage. In the right-to-manage model, the union and the establishment bargain over the wage, but the latter chooses employment. In the efficient bargaining model, both the wage and employment are chosen simultaneously as a result of bargaining. In the U.S., unionization process is mainly decentralized. In some countries, centralized bargaining prevails, where a union may negotiate a common contract with several employers simultaneously.

²⁰Estimates of union organizing costs are hard to come by. Voos (1984) presents some early estimates and finds that total real organizing expenditures per organizable worker remained relatively constant over the years she studied.

and identically distributed continuous random variable drawn, across establishments and over time, from a cumulative distribution function $\Gamma(\omega_a)$ with support $[0, 1]$. This probability is observed during the previous period *before* the targeting decision is made in the current period.

Let $V^u(s_a)$ represent the value that a union obtains from an age- a unionized establishment, given the state, $s_a \equiv (\mu_a, a, x_a, \omega_{a+1})$. The function V^u is defined by

$$V^u(s_a) = B(x_a) + \beta E[V^u(s_{a+1})], \quad (7)$$

where the μ_{a+1} component of s_{a+1} is governed by the law of motion specified in (6). The expectation on the right hand side of (7) depends on the prior μ_a . This prior is used to forecast both x_{a+1} and μ_{a+1} .

The union's value from a non-unionized establishment, $V^n(s_a)$, arises solely from the option to organize this establishment at some future date. This value can be written as

$$V^n(s_a) = \beta \max\{\omega_{a+1}E[V^u(s_{a+1})] + (1 - \omega_{a+1})E[V^n(s_{a+1})] - c, E[V^n(s_{a+1})]\}. \quad (8)$$

The current benefit to the union from a non-unionized establishment is zero. At the beginning of the next period, the union makes a decision about whether or not to target the establishment. It makes this decision *before* it observes x_{a+1} . Therefore, it compares the expected benefit from targeting, $\omega_{a+1}E[V^u(s_{a+1})] + (1 - \omega_{a+1})E[V^n(s_{a+1})] - c$, with the expected benefit from not targeting, $E[V^n(s_{a+1})]$.

2.4 Union Targeting and Unionization

A certification election occurs in an age- a establishment if and only if the expected *net* gain from targeting, $\omega_a\{E[V^u(s_a) - V^n(s_a)]\}$, exceeds the cost of organizing, c

$$\omega_a\{E[V^u(s_a) - V^n(s_a)]\} > c. \quad (9)$$

Now, let $D(s_a) \equiv V^u(s_a) - V^n(s_a)$. The targeting decision depends on the properties of $D(s_a)$. Using (7) and (8) one can write

$$\begin{aligned} D(s_a) &= B(x_a) + \beta E[V^u(s_{a+1})] - \beta \max\{\omega_{a+1}E[V^u(s_{a+1})] + (1 - \omega_{a+1})E[V^n(s_{a+1})] - c, E[V^n(s_{a+1})]\} \\ &= B(x_a) + \beta \min\{(1 - \omega_{a+1})(E[V^u(s_{a+1})] - E[V^n(s_{a+1})]) - c, E[V^u(s_{a+1})] - E[V^n(s_{a+1})]\}. \end{aligned} \quad (10)$$

By using the definition for $D(s_a)$, the right hand side of (10) reduces to

$$\begin{aligned} D(s_a) &= B(x_a) + \beta \min\{(1 - \omega_{a+1})E[D(s_{a+1})] - c, E[D(s_{a+1})]\}, \\ &= B(x_a) + \beta\{(1 - \omega_{a+1})E[D(s_{a+1})] - c\}, \end{aligned} \quad (11)$$

where μ_{a+1} in s_{a+1} is governed by the law of motion (6). The second equality in (11) holds because $1 - \omega_{a+1} < 1$, $c > 0$, and $D(s_a) \geq 0$.²¹ The function D has the following properties.

Lemma 1 (*Properties of D*) *There exists a unique, continuous and bounded function $D(s_a)$ that satisfies (11). $D(s_a)$ is increasing and strictly convex in μ_a , increasing in x_a , and decreasing in a and c . Furthermore, $E[D(s_a)|\mu_{a-1}, a]$ is increasing in μ_{a-1} and decreasing in a .*

From the targeting rule (9) and Lemma 1, for any given a there exists a unique threshold for the probability of union win in a certification election, $\tilde{\omega}(\mu_{a-1}, a)$, defined by

$$\tilde{\omega}(\mu_{a-1}, a) = \frac{c}{E[D(s_a)|\mu_{a-1}, a]}, \quad (12)$$

such that the union targets an establishment whenever $\omega_a > \tilde{\omega}(\mu_{a-1}, a)$. The probability of the union targeting a non-unionized establishment of age a and with prior μ_{a-1} is then given by

$$T(\mu_{a-1}, a) = 1 - \Gamma(\tilde{\omega}(\mu_{a-1}, a)). \quad (13)$$

The main results can now be presented. What type of establishments do unions target for organizing? Proposition 1 answers this question.

Proposition 1 (*Unions target productive, young firms.*) *The probability of the union targeting an establishment, $T(\mu_{a-1}, a)$, is increasing in μ_{a-1} and decreasing in a .*

By Proposition 1, the probability of the union successfully organizing an establishment

$$O(\mu_{a-1}, a) = \omega_a T(\mu_{a-1}, a), \quad (14)$$

is also increasing in μ_{a-1} and decreasing in a . A higher value for μ_{a-1} implies that the union believes that the establishment will yield a greater stream of benefits. Hence, the probability of union targeting and successful organizing rises. As an establishment ages, the variance around

²¹The non-negativity of $D(s_a)$ follows because $V^n(s_a)$ cannot exceed $V^u(s_a)$.

the prior declines, in line with (3) and (5). This reduces the probability that a high value for x_a will be drawn. The decline in the variance around the prior means a lower expected value for the union, given the strict convexity of D , and hence, a lower likelihood of targeting and successfully organizing an establishment.

Next, consider the expected probability of a union win in a certification election *conditional on* the union targeting an age- a establishment. Using (9), this probability can be written as

$$W(\mu_{a-1}, a) = E [\omega_a | \omega_a > \tilde{\omega}(\mu_{a-1}, a)]. \quad (15)$$

Note that $W(\mu_{a-1}, a)$ depends on μ_{a-1} and a , even though the unconditional probability of win, ω_a , is assumed to be an *i.i.d.* random variable independent of μ_{a-1} and a . $W(\mu_{a-1}, a)$ satisfies the following properties.

Proposition 2 (*Unions win elections in less productive, older firms*) *The expected probability of a union win, conditional on the establishment being targeted, $W(\mu_{a-1}, a)$, is decreasing in μ_{a-1} and increasing in a .*

The expected gain from organizing an establishment is higher for young establishments with a high prior. Therefore, the union is willing to target such establishments even for low levels of the probability of winning a certification election.

The probability that an age- a establishment, with a history of priors $(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0)$, is unionized is given by

$$\begin{aligned} U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) &= \sum_{j=1}^a \{\prod_{k=1}^{j-1} [1 - O(\mu_{k-1}, k)]\} O(\mu_{j-1}, j) \\ &= 1 - \prod_{j=1}^a [1 - O(\mu_{j-1}, j)], \end{aligned} \quad (16)$$

where $O(\mu_{a-1}, a)$ is defined by (14). Observe that U is the probability that unionization occurs by the a -th trial, where the probability of success in trial j is $O(\mu_{j-1}, j)$. The following can be stated about unionized establishments.

Proposition 3 (*Unionization is prevalent in productive, old firms.*) *The probability of an establishment being unionized, $U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a)$, is increasing in μ_{a-1} and a .*

A rise in μ_{a-1} increases the probability that the union is targeted in the current period, if it hasn't been organized in the past. Clearly, the chances that an establishment is unionized

in the current period are then higher. An increase in age, a , raises the likelihood that the establishment is organized, since it increases the time interval over which the union could have potentially engaged in targeting activity.

2.5 Testable Implications

Consider now an outside observer (an econometrician) who sees a plant's current productivity, x_a , and its age, a , but not the union's prior, μ_{a-1} . The observer knows the distribution of μ_{a-1} given x_a and a , the distribution of ω_a , and the union's probability of targeting, $T(\mu_{a-1}, a)$. Given x_a and a , the observer has beliefs on μ_{a-1} , represented by the c.d.f. $\Omega(\mu_{a-1}|x_a, a)$. Based on these beliefs, the observer's assessment of the probability that the union targets an age- a establishment with current productivity x_a is

$$T^o(x_a, a) = \int T(\mu_{a-1}, a) d\Omega(\mu_{a-1}|x_a, a). \quad (17)$$

This probability is positively associated with the observed x_a , as outlined in the following proposition.

Proposition 4 (*The probability of targeting from the observer's perspective*) *The probability of targeting from the observer's perspective, $T^o(x_a, a)$, is increasing in x_a .*

Proposition 4 implies the likelihood of targeting is higher for higher values of x_a . How $T^o(x_a, a)$ changes as a increases, however, depends on the magnitudes of two opposing effects. For any given x ,

$$T^o(x, a+1) - T^o(x, a) = \int T(\mu, a+1) d\Omega(\mu|x, a+1) - \int T(\mu, a) d\Omega(\mu|x, a). \quad (18)$$

By Proposition 1, $T(\mu, a+1) \leq T(\mu, a)$. This effect implies that $T^o(x, a+1)$ is no larger than $T^o(x, a)$, ceteris paribus. However, μ has a lower variance when a is higher. Depending on the curvature of T , the effect on T^o of a lower variance for μ can be positive or negative.²² Thus, the sign of (18) depends on the nature of T . Which effect dominates in practice is an empirical question. For instance, if the first effect dominates, T^o is decreasing in a .

For the observer, the probability of the union successfully organizing an establishment is given by $O^o(x_a, a) = \int \omega_a T^o(x_a, a) d\Gamma(\omega_a)$. The probability O^o shares the properties of T^o in

²²If T is strictly concave (strictly convex) in μ , a reduction in the variance of μ implies a higher (lower) T^o .

Proposition 4. Furthermore, the observer's assessment of the expected probability of a union win conditional on targeting, $W^o(x_a, a) = E[E[\omega_a | \omega_a > \tilde{\omega}(\mu_{a-1}, a)] | x_a, a] = \int W(\mu_{a-1}, a) d\Omega(\mu_{a-1} | x_a, a)$, satisfies the following.

Proposition 5 *(The probability of a union win from the observer's perspective) The expected probability of a union win (conditional on targeting) from the observer's perspective, $W^o(x_a, a)$, is decreasing in x_a .*

As in the case of $T^o(x_a, a)$, how $W^o(x_a, a)$ depends on a is dictated by the shape of $W(\mu_{a-1}, a)$.

Consider next the probability $U^o(x_a, a)$ that an age- a establishment with productivity x_a is a union establishment from the observer's perspective. Let $\Psi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x_a, a)$ denote the joint c.d.f. associated with the history of priors for an age- a establishment, conditional on x_a . Then,

$$U^o(x_a, a) = \int U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) d\Psi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x_a, a) d\Gamma(\omega_{a-1}) \dots d\Gamma(\omega_0). \quad (19)$$

The properties of U^o are presented below.

Proposition 6 *(The probability of unionization from the observer's perspective) The probability of being a union establishment from the observer's perspective, $U^o(x_a, a)$, is increasing in x_a and a .*

Now, consider any proxy for x_a (any increasing function of x_a , such as profit, output or size). Because the c.d.f.'s Ω and Ψ in (17) and (19) remain the same if the conditioning is done on the proxy, the probabilities T^o and U^o don't change if x_a is replaced by the proxy. Therefore, Propositions 4-6 continue to hold. In relating the probabilities T^o , W^o , O^o , and U^o to x_a in empirical analysis, one can thus use variables such as establishment size or measures of productivity as proxies of x_a .

A few remarks about the model and its implications are in order. The cost of organizing, c , has been assumed to be constant for simplicity. This cost can depend on the pair (μ_{a-1}, a) . For instance, unions may have a lower cost of organizing in younger and less productive establishments, where managerial inexperience may underlie both lower productivity and an inability to counter union activity. In that case, as μ_{a-1} or a increases, the targeting decision depends on how fast the cost c changes relative to the net benefit from targeting, $E[D(s_a) | \mu_{a-1}, a]$. A

similar argument applies to the probability of a win, ω_a , which can be a function of μ_{a-1} and a . Unions may be less likely to win elections in larger and older establishments because such establishments can have better resources to fend off unionization. Consider, for example, the case where the organizing cost c increases, or the union win likelihood ω_a decreases, as μ_{a-1} increases. The probabilities of the union targeting and successfully organizing an establishment can then be decreasing or non-monotonic functions of μ_{a-1} . The empirical analysis, however, does not impose any restrictions on c or ω_a , and tests the implications of the model *without* using the specific structure of the model. If potentially higher organizing costs or the lower likelihood of a win in larger, more productive establishments indeed overwhelm the otherwise greater benefits to the union from such establishments, unions should be less likely to organize larger, more productive establishments. The empirical analysis will therefore examine the monotonicity of the union targeting and win likelihoods.

The model also ignores the likelihood of union decertification, which is a rare event.²³ For simplicity, it also abstracts from establishment exit. The model implies that unions tend not to target less productive, smaller establishments, which are also generally more likely to fail, as prior empirical evidence suggests.²⁴ Furthermore, because larger, more productive establishments offer a larger surplus to the union, a union can target such establishments even if the union's surplus extraction may increase the exit likelihood of such establishments after unionization. Moreover, unions in general value the survival of organized establishments, as longer-lived establishments provide a longer stream of benefits to the union. Unions will therefore tend to internalize the exit likelihood to some extent. For instance, when an establishment experiences a negative shock to its productivity, the union may engage in less surplus extraction to ensure that the establishment survives. It is therefore not obvious that a potentially higher exit likelihood for a unionized establishment can overturn the prediction that unions target large, more productive establishments. Note also that in the data the observed likelihood of unionization, $U^o(x_a, a)$, depends on both the union decertification rate and the exit rate of unionized establishments. It is an empirical question whether these rates alter any of the predictions in Propositions 4-6.

²³The data indicate that decertification elections occur at an annual rate of less than 1% in unionized establishments – see Appendix C.

²⁴See, e.g., Dunne, Roberts and Samuelson (1989).

Establishments may also take costly actions to lower the likelihood of unionization. For instance, establishments that are more likely to be targeted by a union may raise their wages. They may also counter the union threat by investing into reorganization, anti-union campaigns, or management consulting services. All of these actions may reduce the union targeting likelihood in larger, more productive establishments, which are more likely to be targeted. In general, a union and an establishment can engage in a game of unionization where both take costly actions to achieve their goals. Whether the outcome of a such a game changes the monotonic relationship between union targeting and establishment size or productivity predicted by the model would depend on the specific environment of the game. If large, productive establishments are indeed successful in avoiding being targeted by a union, the data should reveal a negative or potentially non-monotonic relationship between union targeting activity and establishment size or productivity. The empirical analysis allows for these possibilities.

3 Empirical Methodology

The model has predictions regarding four main events associated with unionization over the life-cycle of an establishment. These events are: an establishment is targeted by a union for the first time (the first certification election), a union win in the first certification election *conditional* on the establishment being targeted, the first successful organization (the first certification election *and* union win), and the union status of an establishment at any point in time as defined by whether the establishment has ever experienced a successful certification election. The probabilities associated with these events are explored using the model’s predictions.

It is important to note the sequence of events leading to union certification in an establishment for understanding what the theoretical events described above exactly correspond to in the data.²⁵ A collection of workers in a non-unionized establishment decide, for the first time, to organize and form a union. These workers typically contact a union to seek assistance with organizing. With the help of the union, workers carry out a “card drive”, during which they try to obtain support from at least 30 percent of the workers in order to be legally granted an election by the NLRB. The NLRB then makes a determination on what constitutes the bargaining unit for the workers trying to organize. A certification election is then held among

²⁵See Dinardo and Lee (2004, section II) for more details on these events.

workers eligible to vote. This event is referred to as “union targeting” here, corresponding to the first certification election registered for an establishment in the data. A simple majority is required for a union win in the election. Within 7 days after the ballot tally, objections can be made by both parties, and a re-election can be granted by NLRB if there is sufficient evidence of an improperly carried-out election. If the union has a simple majority in the end, the union is certified as the exclusive bargaining agent for the bargaining unit, and the employer is obliged to negotiate in good faith with the union. The event of this exclusive right being granted to the union is labelled here as “successful organizing”. Therefore, in the data this event corresponds to the first certification election that results in a union win. The exclusive right to negotiate is lost if the establishment exits the business, or the union loses a subsequent decertification election, which is either petitioned by the employer (e.g., in the event of a business restructuring) or by a sufficient number of workers whose jobs are represented by the union. An establishment is labelled here as “unionized” until such an event occurs.

A union may try to organize an establishment with no resulting election (e.g., a failed card drive). Such cases are not observed in the data. The first-ever targeting of an establishment by a union is thus defined as the first-ever union organizing drive that leads to a certification election. The observation of a certification election at an establishment may itself be subject to selection. Unions may target establishments where they are more likely to reach the election stage. It is plausible that unions are less likely to secure an election when they target establishments that are able to resist unionization successfully. Such establishments may be larger and more productive, and may offer higher wages. If this is the case, unobserved targetings with no resulting election may be disproportionately concentrated in larger or productive establishments. An implication is that any positive association found between union targeting and establishment productivity or size would be stronger if targetings with no resulting election were also observed.

An establishment may experience more than one certification election along its life-cycle. Only a small fraction of establishments ($< 5\%$) in the data have two or more certification elections over their life-cycles. Subsequent targetings and successful organizations likely depend on the outcome of the first targeting, and hence, cannot be treated as independent events. They are therefore excluded from the analysis of targeting and successful organization. The union status of an establishment, however, takes into account the outcome of all certification

elections that the establishment experiences. That is, if an establishment experiences more than one certification election, a union win in any of these elections results in unionization, and the establishment then stays unionized until either it exits or a union decertification election takes place.

Let E_{it} be the indicator that event E occurs in a establishment i in year t , where $E \in \{T, W, O, U\}$ denotes the event of first-ever targeting by a union (T), union win in the first-ever certification election conditional on the establishment being targeted by a union (W), the first successful organizing by a union (O), or the event that an establishment is a union establishment (U). Because all of these events have binary outcomes and successes are highly rare (except for event W), the observer's assessed probability, $E^o(x_{it}, a_{it})$, that an establishment experiences event E is modelled using the inverse logit transformation²⁶

$$E^o(x_{it}, a_{it}, y_{it}; \mathbf{b}) = \frac{\exp(F(x_{it}, a_{it}, y_{it}; \mathbf{b}))}{1 + \exp(F(x_{it}, a_{it}, y_{it}; \mathbf{b}))}, \quad (20)$$

where

$$F(x_{it}, a_{it}, y_{it}; \mathbf{b}) = b_0 + \sum_{j=1}^J b_x^j I^j(x_{it}) + \sum_{k=1}^K b_a^k I^k(a_{it}) + b'_y y_{it} + \nu_i + s_i + \tau_t. \quad (21)$$

In (21), x_{it} is classified into one of J bins, where $I^j(x_{it})$ is the indicator that x_{it} falls in bin j . Similarly, a_{it} is classified into one of K bins, with the corresponding indicator $I^k(a_{it})$. These specifications offer a flexible way of accounting for the effects of productivity and age, which are not necessarily linear according to the theoretical model. The variable y_{it} is a vector of controls, ν_i is an industry fixed effect, s_i is a state (geography) fixed effect, and τ_t is a year fixed effect, and $\mathbf{b} \equiv \{b_0, \{b_x^j\}_{j=1}^J, \{b_a^k\}_{k=1}^K, b_y\}$ is the set of parameters.

The parameters \mathbf{b} in (21) can be estimated for each event $E \in \{T, W, O, U\}$ separately by maximizing a weighted log-likelihood function that pools all sample years and all relevant observations

$$\mathcal{L}(\mathbf{b}) = \sum_{t=t_0}^{\bar{t}} \sum_{i=1}^{n_t} w_{it} E_{it} \ln E^o(x_{it}, a_{it}, y_{it}; \mathbf{b}) + (1 - w_{it} E_{it}) \ln(1 - E^o(x_{it}, a_{it}, y_{it}; \mathbf{b})), \quad (22)$$

where $t = t_0, t_0 + 1, \dots, \bar{t}$ denotes a year and n_t is the total number of establishments at risk for

²⁶For example, for the observer's probability of union targeting, $T^o(x_{it}, a_{it}) = \int T(\mu_{a-1}, a) d\Omega(\mu_{a-1} | x_a, a)$ is modeled by an inverse logit transformation. Other probabilities are modeled similarly.

event E in year t .²⁷ Weights, w_{it} , are assigned to establishments to account for the uncertainty with which they match to certification elections. The matching and weighting procedure is explained in more detail in the next section. For an establishment that matches with a given certification election, the weight w_{it} takes on a value in $(0, 1]$ equal to the inverse of the frequency of matches for that certification election. The weight $1 - w_{it}$ is then assigned to the establishment's non-matched version. If an establishment is not matched at all to any certification election, the weight is $w_{it} = 0$.²⁸ The calculation of the standard errors for the estimated parameters take into account the clustering of error terms by establishment.

For all events, the focus is on the estimates of the parameters $\{b_x^j\}_{j=1}^J$ and $\{b_a^k\}_{k=1}^K$. These estimates are used to test the model's predictions in Propositions 4-6. Two strategies are followed. First, two measures of establishment size (employment and the value of shipments) are used separately as a proxy for x_a .²⁹ Second, some measures of productivity are used to proxy for x_a . These are the value of shipments per worker, value added per worker, and a measure of total factor productivity. Establishment age, measured by the number of years elapsed from the first observation of an establishment, is included in all estimations in accordance with (21).

In the control variables, y_{it} , a multi-unit firm indicator is included to assess the effect of being part of a multi-unit firm. A multi-unit firm association may signal to the union that the establishment belongs to a successful firm that has expanded. A firm-level unionization indicator is also included. This indicator equals one if the establishment is part of a firm that already has at least one unionized establishment. This variable accounts for potential spillover of unionization within a firm. Firm union presence may signal to the union that the establishment is more prone to unionization. Hence, both the multi-unit and firm union status are expected to contribute to the union's learning about an establishment's eligibility for being

²⁷For the event of first-ever certification election, the establishments at risk are those that have never experienced a certification election before year t ; for the event of a union win in a certification election, the establishments at risk are the ones that experience their first certification election in year t ; for the event of a successful organization, the establishments at risk are those that are targeted for the first time *and* experience a certification election in year t ; finally, for the event of being a union establishment, the relevant establishments are all establishments in year t .

²⁸In estimations using the Economic Census sample, the establishment's weights in the Economic Census are also used to arrive at population estimates.

²⁹This approach is similar to those adopted in some other studies of firm learning (e.g. Dunne, Roberts, and Samuelson (1989)).

organized. Figure 1 suggests that unions indeed seek information on these characteristics for potential organizing.

Another indicator variable in y_{it} takes on a value of one if the establishment is located in a right-to-work state. During the sample period three states adopted a right-to-work law or amended it, so this indicator varies over time for these states, allowing for the identification of the corresponding coefficient in the presence of state fixed effects.³⁰ Right-to-work states generally have laws and regulations less favorable for union activity. Additionally, establishments may favor location in these states because such states tend also to foster other policies friendly to businesses.³¹ Union organizing activity may therefore be less intense in such states. In the estimation of the probability W^o , the ratio of the workers eligible to vote in the certification election to the establishment's total employment is also included in y_{it} . The higher this ratio, the higher the stakes for the union and the establishment in the election. Therefore, both the union and the management in the establishment may devote more resources to influence the outcome of the certification election. As a result, the effect of this ratio can go in either direction.

For each event of interest, the estimation is carried out for all private sector establishments, and also separately for the manufacturing sector, where unionization has traditionally been more concentrated. For specifications using an establishment's employment as a measure of size, the sample period for estimation is 1977-2007, as employment data are available annually. For specifications where value of shipments or measures of productivity are used, the sample period includes only the Economic Census years (every five years between 1977 and 2007, inclusive), as these variables are available only in the Economic Census, as discussed in the next section (except in the case of total factor productivity, which can be calculated for 1977-2007 using the Census of Manufactures). In all estimations, the sample excludes those establishments that

³⁰Idaho and Oklahoma adopted a right-to-work law in 1986 and 2002, respectively. In 1993, Texas amended its original right-to-work law passed in 1947. Whether Texas is included in this group of three states does not alter the estimates materially. For a chronology of the adoption of right-to-work laws by states, see <http://www.dol.gov/whd/state/righttowork.htm>.

³¹Holmes (1998) documents the sharp increase in manufacturing activity when one crosses the border from a non-right-to-work law state to one with a right-to-work law. As Holmes notes, the law itself may not be the cause of this increase. The presence of the law in a state may serve as a proxy for other business-friendly policies and regulations of the state.

entered before 1977, because for such establishments there is no information on prior union activity before 1977, as discussed in the next section.

4 Data

National Labor Relations Board (NLRB) certification election data for the years 1977-2007 is linked with the corresponding years from the U.S. Census Bureau’s Longitudinal Business Database (LBD), quinquennial revenue data from the Economic Census (EC), and data on total factor productivity from the Census of Manufactures.³² The NLRB data contain information on union certification and decertification elections that took place between 1977-2007. NLRB elections in the year 1977 are only partially observed, see Appendix C. For each election, the data contains the employer’s name, address, and industry. It also contains the number of workers eligible to participate in the election, how many ballots were cast, and how many were cast in favor of the union.

Over the sample period, the NLRB data contain information on a total of 103,064 certification elections.³³ In most years there are roughly 3,000 certification elections. The frequency of these elections in general declines over the sample period. In particular, the number of certification elections drops sharply from about 9,000 in 1977 to about 3,500 in 1983, and continues to drift lower for the rest of the period, with about 1,600 elections in 2007. For an analysis of the trends in NLRB elections, see Appendix C.

The LBD, with which the NLRB elections data is matched, contains the universe of private sector employers in the U.S. at the establishment level. Key variables are the number of employees, industry affiliation, location, the year the establishment enters and exits, and the identifier of the firm that has operational control over the establishment. The LBD is also matched at the establishment level with the EC, which is conducted every five years. This match is done for each quinquennial census between 1977 and 2007, inclusive. The time series

³²The NLRB certification election data come from two sources: the 1977-1999 data was kindly provided by Thomas J. Holmes; data from 1999-2007 is available from data.gov.

³³The 497 of elections that are labelled “Union Shop Clause” are omitted from all empirical specifications, as these only appear in the data obtained from data.gov. The 3,418 “Employer Requested” elections are also omitted, as these tend to be elections that are sought when some fraction of the workforce was previously unionized.

coverage by the EC varies by industry.³⁴ The data collected permit the construction of a revenue-based labor productivity measure, that is, the value of shipments per employee, for all industries. Another measure of labor productivity, value added per worker, is available for manufacturing, and is used as an alternative to the value of shipments-based one. A total factor productivity measure is also used for manufacturing.

NLRB data is consistently available by the employer’s name and the employer’s city and state. The LBD is linked to the NLRB data via a multi-stage matching process. Each stage involves considering establishments in both the LBD and the NLRB data at some level of geography – city and state, fuzzed city name and state, or county and state. Having so “blocked” the data at a particular level of geography, the similarity of business names and industry are considered between the two data sets. Inspection of individual records is used to validate the name and industry agreement rules, along with inspection of the address and zip code that are available consistently for the LBD data, and also for a subset of the years of the NLRB data. The NLRB data contains the number of employees eligible to vote in the certification election. This information is used to reject potential matches, while allowing for the somewhat uncommon event of multiple establishments being included in the same certification election (for example, an election might cover all cashiers in a particular geographic region of a retail chain). If the size of the firm that has operational control over the establishment is less than 80% of the number of employees eligible to vote in the election matched to the establishment, then the match is rejected, and another match is sought. For establishments that have less employees than the number eligible, a progressive search is performed within the firm at increasingly higher levels of aggregation (the address, city, county, state, and national level) until the total number of LBD employees is at least 80% of the number eligible to vote in the election.³⁵ About 73% of certification elections match reliably with the LBD for the sample period.

³⁴The Census of Construction Industries, the Census of Manufactures, the Census of Retail Trade, the Census of Services, and the Census of Wholesale Trade are available every five years from 1977-2007. Other parts of the economy are only available for more recent years. The Census of Finance, Insurance, and Real Estate is available for 1992-2007, the Census of Mining and the Census of Transportation, Communications, and Utilities are available for 1987-2007, and the Census of Finance, Insurance, and Real Estate are available for 1992-2007.

³⁵This rule was established using the records downloaded from www.data.gov as training data. This data contains a free-form text description field that often includes the phrase “at all” when describing elections that cover multiple establishments.

Weights are calculated to account for the uncertainty of matching an election to an establishment. If there are multiple matches to a given certification or decertification election in a year, each establishment receives a weight, w_{it} , equal to the inverse of the total number of such matches. For simplicity, and for longitudinal consistency in the case of the relatively few establishments that link to more than one election (certification or decertification), the largest weight that an establishment receives among all such elections is given as its weight over time. Each establishment involved in a multiple match was also given an additional weight, equal to $1 - w_{it}$, to represent the non-unionized version of this establishment. The weights in the EC were also retained to be used in the analysis to make inference about the population of establishments in the estimations using EC data for size and productivity measures.

Both the NLRB data and the LBD are left-censored. The NLRB data available for this study begins in 1977, and an establishment's union status is unknown if it entered prior to 1977. Therefore, there is no way of identifying whether a certification election that occurs during the 1977-2007 period at such an establishment is, in fact, that establishment's first certification election. Furthermore, the LBD coverage starts in 1976, so there is no exact entry year, hence, no age information for establishments that first appear in 1976. To identify age and union status accurately, the analysis is therefore restricted to all establishments that first appear in LBD in or after 1977.

The constructed dataset contains a weighted sum of nearly 30 million establishments. About 89,400 establishments match to certification elections. Of those that match to an election, about 95% match to exactly one election, and about 4% match to exactly two elections. Most of the remainder match to exactly three elections. Thus, the cases where multiple certification elections occur at a given establishment are rare. Less than half of the elections occur in establishments that are left-censored, and the remainder are among those establishments that entered during or after the year 1977. This skewness reflects the fact that between 1981 and 1982 the number of certification elections dropped from 6,000 – 7,000 to around 3,000 per year and never recovered to its previous level.³⁶

³⁶This sharp drop in early 1980's is also documented by other researchers, see, e.g. Farber and Western (2001).

5 Results

5.1 Size and Age Effects

How are the probabilities of interest related to establishment age and size, as measured by employment or the value of shipments?³⁷ Tables 1, 2, and 3 present the estimated odds ratios based on the estimation of the logit model in (20) for all private sector and manufacturing, respectively. For easy interpretation, Figures 3 and 4 contain the predicted effects of size and age, and also, of year, to explore the trends.³⁸ In what follows, the predicted effects for all remaining control variables are discussed based on the specification that uses employment as the measure of size, rather than the value of shipments. The specification with employment is estimated using all years of data, whereas the value of shipments is available only every five years, effectively reducing the sample size by 80%. Therefore, the effects of all other variables are more precisely estimated in the specification with employment as the size measure, and the year effects are obtained for all years, rather than every five years.

Consider first the predicted probabilities of union targeting in Figures 3 and 4. The probability of targeting increases with both measures of size, although in the manufacturing sector the size effect tapers off and declines slightly, only at very large employment classes.³⁹ When the entire private sector is considered, an establishment in the largest employment (value of shipments) class has almost 10 (4) times the probability of being targeted compared with an establishment in the smallest size class. In manufacturing, this relative likelihood is much higher, 23 (49).⁴⁰ Age effects for the manufacturing sector suggest that union targeting activity is at its peak within the first couple of years after an establishment's entry and flattens out after 10

³⁷The variables in nominal values used in constructing the value of shipments/receipts categories and productivity categories in the next section are not deflated. Because the empirical methodology relies on the ranking of these measures, instead of the absolute values, and because year fixed effects are included, there is no need for deflating. Similarly, time dummies also capture any change in the likelihood of an event over time due to secular changes in the average sales or productivity of establishments. Results are generally qualitatively similar when size and productivity measures in nominal terms are alternatively measured as deviations from their annual means.

³⁸The predicted average marginal effect for a given size or age category is calculated by averaging the predicted probability for that category over all values of the remaining variables across all observations.

³⁹This decline may result from large firms' ability to better counter the threat of union organizing.

⁴⁰These ratios, and others that follow, come directly from the estimated odds ratios in Tables 1 and 2.

to 12 years. The age effects indicate a similar decline in the likelihood of targeting over time in the case of the entire private sector. For the entire private sector, the youngest group of establishments are about 1.76 times as likely to be targeted compared with the oldest group. In manufacturing, this relative likelihood is around 1.72. These patterns are consistent with a learning process and support Propositions 1 and 4. The predicted effects for year in Figures 3 and 4 also point to a decline in the probability of targeting over the sample period. This secular decline in union organizing activity is one reason behind the decline in private sector union membership rate in the U.S.⁴¹

Next, turn to the predicted probability of a union win in a certification election, conditional on an establishment being targeted by a union. In general, the predicted probability of a win declines as establishment size increases, consistent with Propositions 2 and 5. For the case of manufacturing, in the largest employment (value of shipments) category the probability of win is about 22% (30%), and about 60% (50%) in the smallest category. The likelihood of win does not appear to change substantially with establishment age. When the entire private sector is considered, the union win likelihood also declines as establishment size increases, regardless of the size measure. Overall, there is evidence that unions are less successful in winning certification elections in larger establishments.⁴² The year effects indicate that, for much of the sample period, the probability of a union win has had little or no trend, with one exception: there is some rise in the union win likelihood starting in the early 2000's when the entire private sector is considered.⁴³ The relatively stable average union win likelihood in certification elections, combined with the declining union targeting activity, implies a declining rate of unionized business formation over the years.

Consider now the predicted probability that a union organizes an establishment successfully for the first time. These are also presented in Figures 3 and 4. In manufacturing, the probability

⁴¹While the model considers a stationary environment, such a secular decline in union targeting activity can be obtained from the model, for instance, when the cost of union organizing is rising over time, reflecting an increasingly unfavorable environment for the unions.

⁴²While the model does not posit a specific connection between union win probability and establishment characteristics, larger establishments may also be better in campaigning against unions in certification elections, perhaps due to better management, organizational capabilities, and better labor conditions and compensation.

⁴³See Farber (2013) for an exploration of the trends in voter turnout in union elections and the implications of declining union organizing on union win likelihood in elections.

of a union successfully organizing an establishment increases as both measures of establishment size increase. Therefore, the decline in the likelihood of a union win as size increases is not enough to overcome the steep increase in the likelihood of targeting. For the largest employment (value of shipments) class, the probability of the union successfully organizing an establishment is nearly 15 (30) times the probability in the smallest class. The likelihood of successfully organizing an establishment in the manufacturing sector declines with age, as a consequence of the fact that the likelihood of targeting declines with age, whereas the likelihood of a union win in an election is not significantly different across age categories. For the entire private sector, the likelihood of a union successfully organizing an establishment in the largest employment (value of shipments) category is about 12 (5) times that in the smallest category. Note also that likelihood of a union successfully organizing an establishment has declined persistently over the sample period, as indicated by the year effects.

Next, observe that the effects of size and age on the probability of being a union establishment are highly pronounced. Larger and older establishments are more likely to be unionized. Considering the entire private sector, establishments in the largest employment (value of shipments) size class are about 11 (4) times more likely to be unionized, compared with the ones in the smallest class. In manufacturing, this relative likelihood is about 7 (12). The age effects indicate that the oldest group of establishments are about 10 times more likely to be unionized compared with the youngest group when all sectors are considered, and 21 times more likely in manufacturing. The relationship between unionization likelihood, on the one hand, and the establishment size and age, on the other, provide support for Propositions 3 and 6. The year effects indicate that the probability of being unionized has declined substantially since late 1970's. As mentioned before, this decline is in part driven by the decline in the rate of certification elections shown in Figures 3 and 4, as the probability of a union win in a certification remained relatively stable over. The exit of union establishments and union decertification also contribute to deunionization. The former effect is a much bigger source of union dissolution compared with the latter, which is a rare event.⁴⁴ While these two modes of union dissolution are not the focus of this study, their effects across size and age categories do not appear to be large enough to overturn the monotonic size and age effects predicted by Propositions 3 and 6.

⁴⁴Less than 1% of establishments that previously had a certification election experience a subsequent decertification election over the sample period – see Appendix C.

5.2 The Effects of Other Controls

Other controls included in the model (20) have effects that are generally consistent with what was expected a priori, as can be seen in Tables 1 and 2. Establishments that are part of a multi-unit firm and have at least one unionized sister establishment have higher odds of experiencing union certification elections. In manufacturing, multi-unit status and having a sister establishment that is unionized each double the odds of being targeted by a union. When all sectors are considered, the relative odds are higher: about 5.5 and 3.5, respectively, for multi-unit status and firm-union status. Unions win certification elections with higher probability in cases where there is already at least one unionized establishment in a firm (37% versus 56% win rate in manufacturing). Multi-unit status has the opposite effect on the probability of union win (44% versus 38% win rate in manufacturing), consistent with unions winning elections with lower likelihood in larger establishments. Establishments that are part of a multi-unit firm, and establishments that have unionized sister establishments, have higher odds of being unionized. All sectors taken together, the predicted probability of being a union establishment is about 3 times larger if an establishment is part of a multi-unit firm. This predicted probability is also nearly 5 times larger, if there is at least one sister establishment that is already unionized. In manufacturing, these relative likelihoods are about 2.8 each.

When state fixed effects are present (as in Tables 1 and 2), the right-to-work status of a state usually does not have a highly significant effect on targeting, union win in certification elections, or union status. With state effects included, the coefficient for the right-to-work law status is identified only through a small number of states that changed their status during the sample period. The effect of the law does not seem to be estimated with precision. When state fixed effects are *not* included in the estimations, establishments located in states with a right-to-work law have significantly lower odds of being targeted.⁴⁵ For example, in manufacturing the predicted marginal effect of being located in a non-right-to-work law state on union targeting is about 1.5 times that of being located in a right-to-work law state. Unionization odds are also lower in right-to-work law states. In manufacturing, the predicted likelihood of being a union establishment in a right-to-work law state is about half that in a non-right-to-work law state.

Finally, the share of employees in an establishment eligible to vote in a certification elec-

⁴⁵These estimates are available upon request.

tion consistently tends to be negatively associated with the odds of a union win. One possible explanation for this effect is that when a large fraction of the employees are at risk of becoming organized, the management in the establishment may choose to devote more resources to thwarting the union campaign, resulting in a lower likelihood of a win for the union. In general, the estimated effects of the control variables support the view that unions use favorable signals such as multi-unit affiliation, union presence in the firm, and location in a non-right-to-work-law state to determine an establishment's eligibility for organizing, in addition to size and age signals.

5.3 The Effects of Productivity

Three measures of productivity are considered in turn. The first one is an establishment's total value of shipments per employee. The appeal of this measure is that it can be calculated for all sectors in the private economy. The second measure is the value added per employee, and it is available only for certain sectors. This measure is used for the manufacturing sector for comparison with the other measures. A third measure, total factor productivity, is also available for manufacturing.⁴⁶ Total factor productivity is the ideal measure from the model's perspective. From an empirical point of view, however, it is the most difficult to measure and is in general subject to more measurement error than other productivity measures. All productivity measures are computed for Economic Census years – every five years between 1977 and 2007, inclusive.⁴⁷

Table 4 presents the odds ratios obtained from the estimation of the logit model in (20) using measures of productivity instead of establishment size, and including all other controls as in Tables 1-3. The estimates for the control variables are generally similar to the ones in Tables 1 and 2, and are omitted. Tables 1 and 2 give much more precise estimates for age and other controls, and complete estimates of year effects, as these estimates are based on annual data, rather than the quinquennial observations used in Table 4. The predicted probabilities

⁴⁶Zoltan Wolf has kindly provided help with the data on the revenue-based total factor productivity measure, which is calculated using the methodology in Foster, Haltiwanger and Krizan (2001) – see their work for the details of total factor productivity calculation.

⁴⁷In the analysis, the highest and lowest percentiles of the distributions of all productivity measures are trimmed to prevent any influence of likely outliers.

for productivity measures based on Table 4 are in Figures 5 and 6.

Consider first the effects of productivity measures on the union targeting (election) likelihood. The probability of targeting increases as total value of shipments per employee increases, both in the case of all sectors and manufacturing. In manufacturing, establishments in the top decile are about 3.8 times more likely to be targeted compared with the bottom decile. In manufacturing, there is also a rise in the likelihood of targeting as value added per employee increases, but the effects are less pronounced and the differences across productivity percentiles are not always highly significant. The top decile of value added per employee in manufacturing is about 2.7 times more likely to be targeted compared with the bottom decile. The probability of union targeting also increases as total factor productivity increases. The top decile in manufacturing has about 2.5 times the likelihood of being targeted compared with the bottom decile. When the entire private sector is considered in Figure 6, the differences across percentiles of value of shipments per labor are less pronounced and not highly significant. The top decile is about 2.3 times more likely to be targeted compared with the bottom decile. Overall, these effects give support to Propositions 1 and 4.

Turn, next, to the relationship between the productivity measures and the probability of union win in a certification election. In manufacturing, there is some decline in this probability as productivity increases, except in the case of total factor productivity. This pattern gives some support to Propositions 2 and 5. In manufacturing establishments experiencing a certification election, the union win probability is about 1.6 (1.5) times higher in establishments in the bottom decile compared with the ones in the top decile, based on the value of shipments per employee (value added per employee). In the case of the entire private sector, the likelihood of a union win does not appear to change significantly across productivity categories.

The predicted probability for successfully organizing an establishment follows a similar pattern to that for targeting an establishment for the first certification election. The increase in the likelihood of targeting generally overwhelms the slight decline in the likelihood of a union win in a certification election, leading to a positive association between productivity measures and the likelihood of first successful union organizing in an establishment. In the manufacturing sector, based on the value of shipments per employee (total factor productivity) unions are 2.6 (2.7) times more likely to successfully organize an establishment in the top decile of productivity compared with the bottom decile. The association between successful organizing

and productivity is somewhat weaker and is not highly significantly different across categories in the case of value added per employee. The top decile is about 1.8 times more likely to be successfully organized compared with the bottom decile. For the entire private sector, the estimated likelihood of successfully organizing approximately doubles going from the bottom decile to the top decile of value of shipments per employee.

Finally, observe that more productive establishments are also more likely to be unionized regardless of the productivity measure. This conclusion holds for all sectors as well as manufacturing. In manufacturing, establishments in the highest productivity decile are about 3, 2.2, and 2.3 times more likely to be unionized compared with the ones in the lowest decile, for value of shipments per employee, value added, and total factor productivity, respectively. In the entire private sector, the likelihood of unionization is about twice as high in the top decile of value added per employee, compared with the bottom decile. The positive association between unionization and productivity generally support Propositions 3 and 6.

5.4 Robustness Analysis and Additional Results

Some robustness checks reinforce the patterns observed so far. First, the estimated age effects analyzed pertain to all establishments that were born in or after 1977. The estimated year effects, however, indicate that there was a much higher rate of union organizing activity in the late 1970s and early to mid 1980s. This higher rate of organizing activity can lead to a disproportionate targeting of young firms born in late 1970s and early 1980s, possibly resulting in a spurious negative correlation between age and union targeting likelihood (though the year effects would absorb some of this effect). As a robustness check for the age effects, all estimations were repeated after restricting the sample to the 1990-2007 period. The results are shown in Tables D.1 and D.2 in Appendix D. All estimated odds ratios, including those for age categories, are largely similar in magnitude and significance to those obtained in Tables 1 and 2. Therefore, the results do not seem to be driven by the much higher rate of union organizing activity in the early part of the sample period.

The empirical analysis controls for any general effects of calendar time by using year fixed effects. However, union certification elections have been on a secular decline, especially after 1982. In this unfavorable environment, unions may be increasingly focusing on more lucrative

targets. The estimated odds ratios for size categories in Table D.2 indicate that between 1990-2007 there was indeed a steeper size profile for the odds of targeting in the manufacturing sector. Unions had even higher odds of targeting larger establishments compared to smaller ones during this period, as opposed to the entire 1977-2007 period. To further investigate whether the estimated effects are robust across different time periods, Table D.3 repeats the analysis in Table 2 for two different time periods: 1977-1982 and 2000-2007. These periods are chosen to highlight any stark differences. During the 1977-1982 period, union organizing activity was much more intense, with many certification elections taking place. By 2000, the union organizing activity had already experienced a long decline, and the unions had likely adjusted to this unfavorable environment with potentially new strategies for targeting. For both periods, the estimated odds ratios in Table D.3 are generally consistent with the estimates for the entire period in Table 2, and the general conclusions about size and age effects remain.⁴⁸ However, there are some important differences. Table D.3 points to a much steeper size profile for the odds of targeting for the 2000-2007 period. As their environment continued to become less favorable for organizing, unions tended to target larger establishments in manufacturing with even higher odds. This finding is also consistent with the evidence documented by Farber (2013) regarding the union election and voter turnout patterns in a deteriorating environment for unions.

Another concern is that very small establishments (those with less than 5 employees) may be matched with some certification elections as a result of the matching algorithm, even though these establishments may not have significant union activity associated with them. Employment, age, and other data associated with these establishments may also in general be subject to more measurement error compared to the larger ones. In addition, many of these establishments tend to be young. The age effects found may therefore be in part driven by these establishments. As a robustness check, the estimations in Tables 1 and 2 were repeated after excluding small firms with 1-4 employees. The results are shown in Tables D.4 and D.5 in Appendix D. The size and age effects remain significant and qualitatively similar to those reported in Tables 1 and 2.

In results not reported, the differences in the age effects for establishments of multi-unit

⁴⁸Note that the estimated age effects for the 1977-1982 period are limited to the age groups of 0-3 and 4-6 years, as establishments born during the 1977-1982 period cannot be more than 5 years old.

versus single-unit firms were also explored. For this purpose, interaction variables between the multi-unit indicator and age categories were added to the specifications in Tables 1 and 2. For manufacturing, the age effects observed in union targeting remain significant and become even more negative for establishments of single-unit firms, but there is little evidence that age matters much for establishments of multi-unit firms. When all private sector is considered, the effects of age are somewhat less pronounced for both types of establishments.

Another factor that may be relevant for a union's targeting decision is an establishment's ability to offer wages above and beyond what it currently pays to its employees. For instance, a productive establishment may face higher demand for unionization if its wages are below what its productivity would imply. A full analysis of the union targeting-establishment wage relationship is not the focus of this paper. This relationship requires a more detailed exploration. Here, an initial look at whether unions target establishments that are able to provide higher wages is provided based on the logistic regression framework used so far. One potential measure of this ability is the gap between an establishment's labor productivity and the average wage paid to its employees. Establishments that have higher labor productivity, but pay a relatively low average wage, offer potentially a greater surplus to a union, which the latter can partially extract by demanding higher wages. More generally, establishments that offer lower wages may also be more likely to be targeted by unions, as demand for unionization would be higher by workers in such establishments. However, some unions may also choose to target high-wage establishments because in such establishments a union can readily secure high wages and benefits for its members even when the union does not have a high bargaining power. It is well known that larger and older establishments tend to pay higher wages.⁴⁹ Given that the results so far indicate a positive association between union targeting and establishment size measures, it may be that unions target high-wage establishments. Another consideration is that establishments that are more likely to be targeted by a union may raise their wages to avoid unionization.⁵⁰ Such an effect implies that wages cannot be treated as exogenous to the union targeting decision.⁵¹ In addition, wages in already unionized establishments are likely not

⁴⁹See, e.g., Brown and Medoff (1989, 2003).

⁵⁰See Taschereau-Dumouchel (2012) for an analysis of the union threat effect.

⁵¹Under this union threat effect on wages, unobserved characteristics of an establishment that lead to a higher likelihood of unionization can also imply higher wages. Therefore, the unobserved factors determining the union's expected net benefit from targeting could be positively correlated with wages.

independent of the unobservable characteristics of an establishment that lead to union presence. An analysis of these issues of endogeneity are left for future work.

The analysis in Table 4 was repeated using the difference between labor productivity and the average wage in an establishment in place of productivity. The results, shown in Table D.6, are generally similar to the productivity effects. Establishments that have a larger gap between labor productivity and average wage tend to have higher odds of being targeted by a union. Furthermore, unions appear to have lower odds of winning a certification election in establishments with a larger gap. Tables D.6-D.9 provide estimates of the parameters of interest by including two measures of wages in the specifications for Tables 1 and 2. The first specification (Tables D.7 and D.8) includes the logarithm of the average wage paid by an establishment.⁵² The second specification (Tables D.9 and D.10) uses the difference between the logarithm of the average wage and its predicted value based on observable characteristics of an establishment.⁵³ This measure aims to identify establishments that pay a relatively high average wage compared to what they would be expected to offer based on their characteristics. The results in Tables D.7-D.10 indicate that conditional on size and age, establishments with high average wage and high difference between actual and predicted average wage have higher odds of being targeted, but, at the same time, unions have lower odds of winning elections in establishments that offer higher average wage and have larger difference between actual and predicted average wage.

6 Discussion

The size effects in Figures 3 and 4, and the productivity effects in Figures 5 and 6, are generally consistent with the main predictions of the theory. Productive and larger establishments

⁵²It should also be noted that the average wage pertains to all employees in an establishment, not just the ones in the potential bargaining unit. The data does not allow the separate measurement of wages for the workers that try to organize. This measurement error can be important, especially when there is a large wage inequality among workers in an establishment. Note also that establishments with zero reported payroll are excluded from the analysis.

⁵³The predicted (log) average wage is obtained from an establishment-level regression of the logarithm of the average wage on employment, age, multi-unit status, right-to-work status, state, 2-digit industry, and year fixed effects.

are more likely to be targeted by unions, and more likely to be successfully organized despite a lower likelihood of a union win in certification elections in such establishments. Unionization also tends to be concentrated in productive and large establishments at any point in time. The age effects in Figures 3 and 4 also indicate that unions generally organize in establishments when they are younger, consistent with the learning process in the proposed model.

The productivity effects are generally much less pronounced and statistically less significant than the size effects. This result is not an artifact of the way size and productivity categories are chosen to generate the estimates. The relatively weaker effects of productivity are apparent even when the estimates of the size effects in Figures 3 and 4 are obtained for the same percentile categories used for the productivity measures.⁵⁴ There are several reasons why the productivity effects may be weaker than size effects. The productivity effects are estimated using much smaller samples (about 20% of the samples in the case of the size effects), decreasing the precision of the estimates to some degree. Productivity measures also contain more measurement error than size measures, as they confound the measurement errors in employment, revenue, and the costs of inputs. In addition, measures of current productivity and current size need not always be strongly related. Suppose there are adjustment costs for capital or labor. Then, a temporary fall in productivity may not be associated with a decline in size.⁵⁵ It is also possible that unions care more about size (employment) than the underlying productivity of an establishment. In particular, unions that value broader employment for their members would target businesses that are large, but not necessarily highly productive. Moreover, unions may be using establishment size and sales as a stronger indicator of eligibility for organizing, as they are much easier to observe than productivity. Theoretically speaking, there should exist a weaker relationship between unionization and labor productivity, as measured by the value of shipments per worker. For instance, imagine an environment where non-unionized

⁵⁴For instance, when employment is measured in categories of percentiles, the top decile has 30 times the odds of being targeted compared to the bottom decile in manufacturing. From an expository point of view, size measures expressed in terms of the absolute-value categories of employment and the value of shipments/receipts are more readily interpreted than percentiles.

⁵⁵These adjustment costs may also apply to the “customer capital” of a firm. In models of stochastic customer acquisition under informational frictions, such as Fishman and Rob (2002) or Dinlersoz and Yorukoglu (2012), larger firms need not necessarily have high current productivity. In general, size and productivity can also have different levels of persistence.

establishments choose employment given an exogenous competitive wage common across all establishments. In such a case, non-unionized establishments with different levels of total factor productivity would hire labor up to the point where the marginal product of labor is equal (to the common wage) across establishments. Labor productivity would then not be related to underlying total factor productivity.⁵⁶ Hence, it may not be related to the patterns of union targeting. One way to gauge the relative importance of size and productivity in union activity is to include both measures simultaneously in the estimations. Tables D.11 and D.12 repeat the analysis in Tables 1 and 2 by including both size (employment) and labor productivity (the value of shipments per worker) together in the estimations. These head-to-head comparisons indicate that size retains its sign and significance as in Tables 1 and 2, but labor productivity effects tend to be weaker, especially when all sectors are considered together. In manufacturing, however, the size and productivity effects still retain their signs and significance.

Note also that size and productivity effects are generally much stronger and statistically more significant in manufacturing than in the entire private sector. These differences may be driven by a number of reasons. One reason is the differences in the size distribution of establishments across sectors. In manufacturing, there is a higher fraction of very large establishments and a smaller fraction of small ones, compared with sectors such as retail or services. The higher dispersion of establishment size in manufacturing implies that the size effects may be stronger and more pronounced across size categories. Another reason may be the stronger tradition of unionization in manufacturing than the rest of the economy. During the sample period 1977-2007, manufacturing employment exhibited an average union membership rate of 20.8%, whereas the corresponding figure in the entire private sector was 12.4%.⁵⁷ The higher level of union activity in manufacturing suggests that unions in manufacturing may be more experienced in identifying and organizing establishments that are more productive.

The model emphasizes union learning as a potential mechanism behind the unionization process. The age effects documented in the data are consistent with this mechanism. There may be alternative explanations for why age matters, though. For instance, unions may target

⁵⁶As an example, take a production function of the form $e^x l^\alpha$, where l is labor and e^x is the total factor productivity. If the wage is w , then the first-order condition for labor is $\alpha e^x l^{\alpha-1} = w$. Labor productivity is then w/α , which will be common across firms even if total productivity differs.

⁵⁷See the *Union Membership and Coverage Database* at www.unionstats.com.

younger businesses in part because management in such businesses may be inexperienced and less aware of union strategies. It may be easier for unions to organize such young businesses, as management opposition would be less intense, effectively lowering the cost of organizing and also potentially increasing the likelihood of a union win in a certification election.

The theory and empirics abstracted from establishment exit. Note, however, that the estimated probabilities of being a union establishment embed the effects of unions on exit likelihood. If, for instance, unionization increases the likelihood of exit, the steady-state rate of unionization in the establishment population would be lower than in the case where unions had no effect on exit. These considerations, however, do not lead to empirical patterns that are inconsistent with the model's predictions. The analysis also assumed away union decertification. The rate of decertification in the population of unionized establishments is reflected in the estimated rate of unionization. Because decertifications are rare, they do not seem to have a large effect on the overall rate of unionization, and the relationship between unionization and size, age, or productivity.

7 Conclusion

Despite the long presence of union activity in the U.S., there has been a lack of systematic evidence on which types of businesses unions select for organizing and when. A dynamic model of union learning and organizing is proposed here to analyze the nature of union activity. A union monitors establishments in an industry to learn about their underlying unknown productivity over time. Based on the learning process, it decides which firms are the most lucrative for organizing. The model predicts that the union organizes productive and larger establishments at any given age, and younger ones at any given size or productivity level. In the model, the union's benefit from organizing an establishment is increasing in the establishment's productivity, because a more productive establishment can provide higher wages and larger employment to the union. Establishment age also matters because the union's information about productivity becomes more precise over time. Thus, the probability of obtaining a high level of benefits that occur in the right tail of the productivity distribution diminish as an establishment ages. Conditional on size, an older establishment then generates a lower expected benefit to the union, at least when benefits are convex in productivity. A convex benefit function emerges

in a wide variety of models of firm-union conduct. The likelihood of being targeted by the union therefore declines with age, conditional on size. Furthermore, conditional on targeting an establishment, the union has a lower expected probability of winning a certification election when the establishment targeted is more productive or larger.

The data on NLRB union elections matched with establishment-level data from U.S. Census Bureau's LBD and EC data generally support the model's predictions. The main message is that there are clear selection effects in union organizing. First, unions tend to target large and productive establishments early in their life-cycles. Second, unions are less likely to win certification elections in larger and more productive establishments, conditional on targeting. This second effect is generally dominated by the first one, resulting in a positive relationship between successful union organizing and productivity, as well as size. Size, age, and productivity effects appear to be more pronounced in the manufacturing sector, where unions have been traditionally stronger.

The selection effects documented here are relevant for studies on the effect of unions on firm- and establishment-level outcomes. Unionized establishments are not a random sample of the establishment population. If post-unionization outcomes depend on size, age, and productivity around the time of unionization, this selection matters. The effect of unions on the economy overall, and on its individual sectors, can also depend on the nature of this selection. To the extent that unions target larger, more productive firms in an economy, any effects of unions on business outcomes may be amplified. For instance, the decline of manufacturing could have been influenced by the presence of unions in larger establishments. Increasing global competition has been putting downward pressure on the profitability of all manufacturing establishments, but especially the larger ones that produce basic goods highly substitutable with imports, leading to their downsizing or exit.⁵⁸ Whether unions have been able to extract rents from such establishments and reduce their profitability, and hence have accelerated the decline in manufacturing, is still an open question. For example, Dinardo and Lee (2004) find little effect of unions on the survival of relatively large manufacturing establishments, a result that may in part be driven by the selection of unions into large, productive establishments. These large establishments may be able to withstand any adverse effects of unions on survival. Future work can quantify the potentially disparate effects of unions on small and medium versus large

⁵⁸See Holmes (2011).

establishments by using the comprehensive dataset constructed in this study. The decline in manufacturing itself may have led to fewer targets for unions, as larger establishments have gradually disappeared. The role of such effects in the decline of unionization can also be assessed. Finally, studies of unionization at the micro level can use the results and estimates in this paper as a guide in constructing models of the unionization process consistent with the patterns observed in the data.

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TABLE 1. Odds ratios based on logit model estimates – All Sectors
(Size (employment) and age effects)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	2.66 *** [0.038]	0.75 *** [0.023]	3.48 *** [0.074]	2.46 *** [0.043]
20-49 employees	4.44 *** [0.063]	0.59 *** [0.019]	5.56 *** [0.118]	4.02 *** [0.080]
50-99 employees	6.46 *** [0.109]	0.49 *** [0.019]	7.73 *** [0.198]	6.11 *** [0.143]
100-249 employees	8.29 *** [0.152]	0.43 *** [0.018]	9.46 *** [0.226]	8.04 *** [0.215]
250-499 employees	9.25 *** [0.263]	0.39 *** [0.027]	10.74 *** [0.455]	8.75 *** [0.341]
500+ employees	10.24 *** [0.366]	0.27 *** [0.024]	11.84 *** [0.628]	11.49 *** [0.624]
4-6 years	0.82 *** [0.010]	1.04 * [0.046]	0.88 *** [0.016]	2.58 *** [0.023]
7-9 years	0.75 *** [0.011]	1.00 [0.033]	0.81 *** [0.017]	3.85 *** [0.046]
10-12 years	0.69 *** [0.012]	1.09 ** [0.043]	0.78 *** [0.019]	4.94 *** [0.069]
13-15 years	0.68 *** [0.014]	0.98 [0.046]	0.75 *** [0.022]	6.26 *** [0.100]
16-18 years	0.64 *** [0.016]	1.03 [0.060]	0.74 *** [0.026]	7.65 *** [0.139]
19-21 years	0.63 *** [0.020]	1.01 [0.071]	0.71 *** [0.031]	9.22 *** [0.191]
22-24 years	0.59 *** [0.024]	0.93 [0.088]	0.67 *** [0.038]	10.84 *** [0.263]
25+ years	0.57 *** [0.029]	1.06 [0.129]	0.62 *** [0.043]	13.17 *** [0.400]
Multi-unit status	4.92 *** [0.063]	0.44 *** [0.012]	3.29 *** [0.061]	3.05 *** [0.063]
Firm union status	3.40 *** [0.029]	5.45 *** [0.142]	5.51 *** [0.062]	5.04 *** [0.063]
Right-to-work status	0.88 * [0.054]	0.89 [0.106]	0.83 * [0.051]	0.93 [0.051]
Eligible employees %	—	0.75 *** [0.005]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates

significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects.

The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE 2. Odds ratios based on logit model estimates – Manufacturing
(Size (employment) and age effects)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	5.79 ^{***} [0.290]	0.63 ^{***} [0.070]	6.46 ^{***} [0.475]	2.78 ^{***} [0.132]
20-49 employees	14.45 ^{***} [0.663]	0.49 ^{***} [0.050]	14.75 ^{***} [1.014]	5.65 ^{***} [0.280]
50-99 employees	25.05 ^{***} [1.265]	0.37 ^{***} [0.040]	21.73 ^{***} [1.691]	8.58 ^{***} [0.489]
100-249 employees	31.69 ^{***} [1.692]	0.28 ^{***} [0.031]	23.35 ^{***} [1.954]	9.72 ^{***} [0.619]
250-499 employees	32.82 ^{***} [2.163]	0.25 ^{***} [0.036]	23.72 ^{***} [2.548]	8.91 ^{***} [0.743]
500+ employees	22.98 ^{***} [1.964]	0.17 ^{***} [0.035]	14.96 ^{***} [2.194]	7.11 ^{***} [0.838]
4-6 years	0.85 ^{***} [0.028]	0.87 ^{**} [0.060]	0.78 ^{***} [0.040]	2.68 ^{***} [0.681]
7-9 years	0.77 ^{***} [0.030]	0.89 [0.073]	0.73 ^{***} [0.046]	4.18 ^{***} [0.152]
10-12 years	0.68 ^{***} [0.032]	1.04 [*] [0.103]	0.71 ^{***} [0.052]	5.83 ^{***} [0.256]
13-15 years	0.66 ^{***} [0.035]	0.78 ^{**} [0.097]	0.57 ^{***} [0.055]	7.75 ^{***} [0.393]
16-18 years	0.64 ^{***} [0.043]	0.98 [0.140]	0.63 ^{***} [0.069]	10.02 ^{***} [0.583]
19-21 years	0.62 ^{***} [0.052]	1.08 [0.188]	0.64 ^{***} [0.084]	12.84 ^{***} [0.861]
22-24 years	0.53 ^{***} [0.062]	0.52 ^{**} [0.136]	0.38 ^{***} [0.080]	15.94 ^{***} [1.248]
25+ years	0.57 ^{***} [0.086]	1.19 [0.379]	0.59 ^{***} [0.132]	20.62 ^{***} [1.969]
Multi-unit status	2.09 ^{***} [0.067]	0.78 ^{***} [0.043]	1.92 ^{***} [0.095]	2.78 ^{***} [0.143]
Firm-union status	2.01 ^{***} [0.066]	2.16 ^{***} [0.156]	2.82 ^{***} [0.136]	2.84 ^{**} [0.109]
Right-to-work status	0.85 [0.112]	0.98 [0.274]	0.83 [0.177]	0.99 [0.110]
Eligible employees %	—	0.78 ^{***} [0.017]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates

significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects.

The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE 3. Odds ratios based on logit model estimates
(Size (value of shipments/receipts) effects)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
All Sectors (Value of Shipments or Receipts)				
\$250-500K	1.13 *** [0.042]	1.06 [0.085]	1.68 *** [0.097]	1.24 *** [0.029]
\$500K-1M	1.21 *** [0.047]	0.87 * [0.072]	1.81 *** [0.111]	1.35 *** [0.034]
\$1M-2.5M	1.72 ** [0.066]	0.73 *** [0.060]	2.44 *** [0.151]	1.86 *** [0.047]
\$2.5-5M	2.49 *** [0.107]	0.54 *** [0.053]	3.37 *** [0.234]	2.99 *** [0.084]
\$5-10M	3.00 *** [0.138]	0.49 *** [0.052]	3.99 *** [0.298]	3.87 *** [0.116]
\$10M+	3.74 *** [0.160]	0.34 *** [0.036]	4.95 *** [0.350]	4.43 *** [0.135]
Manufacturing (Value of shipments)				
\$250-500K	2.03 *** [0.428]	1.23 [0.603]	2.14 *** [0.631]	0.87 [0.088]
\$500K-1M	6.09 *** [1.034]	1.43 [0.555]	7.38 *** [1.744]	1.94 *** [0.166]
\$1M-2.5M	12.24 *** [1.915]	0.87 [0.308]	11.24 *** [2.526]	3.96 *** [0.300]
\$2.5-5M	23.89 *** [3.856]	0.68 [0.237]	19.08 *** [4.605]	6.61 *** [0.522]
\$5-10M	31.42 *** [5.315]	0.56 * [0.196]	22.42 *** [5.762]	8.78 *** [0.731]
\$10M+	48.78 *** [8.374]	0.42 *** [0.141]	29.97 *** [7.900]	11.18 *** [0.959]

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include all other explanatory variables in Tables 1 and 2. The following categories are omitted: \$0-250K value of shipments and 0-3 years of age..

TABLE 4. Odds ratios based on logit model estimates

(Productivity effects)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
All Sectors (Value of shipments or Receipts per employee)				
11-25 percentile	0.89 [0.255]	0.46* [0.203]	0.59 [0.203]	1.21 [0.213]
26-50 percentile	1.76** [0.461]	0.68 [0.237]	1.39 [0.530]	1.66*** [0.215]
51-75 percentile	1.96*** [0.511]	0.74 [0.313]	1.86 [0.752]	2.16*** [0.239]
76-90 percentile	1.53* [0.359]	0.97 [0.421]	1.38 [0.537]	2.01*** [0.246]
91-100 percentile	2.27** [0.802]	1.32 [0.781]	2.16* [0.984]	2.15*** [0.384]
Manufacturing (Value of shipments per employee)				
11-25 percentile	1.80*** [0.274]	0.66 [0.221]	1.62** [0.356]	1.22** [0.107]
26-50 percentile	2.98*** [0.393]	0.67 [0.199]	2.57*** [0.479]	1.78*** [0.132]
51-75 percentile	2.91*** [0.388]	0.59* [0.176]	2.31*** [0.438]	2.09*** [0.150]
76-90 percentile	3.54*** [0.485]	0.38*** [0.118]	2.15*** [0.435]	2.46*** [0.177]
91-100 percentile	3.88*** [0.534]	0.43*** [0.135]	2.57*** [0.506]	2.93*** [0.218]
Manufacturing (Value added per employee)				
11-25 percentile	1.96*** [0.277]	0.70 [0.211]	1.76*** [0.357]	1.44*** [0.113]
26-50 percentile	2.45*** [0.315]	0.72 [0.200]	2.15*** [0.394]	1.85*** [0.126]
51-75 percentile	2.26*** [0.297]	0.51** [0.146]	1.61*** [0.310]	1.97*** [0.131]
76-90 percentile	2.40*** [0.342]	0.56** [0.172]	1.82*** [0.383]	2.04*** [0.139]
91-100 percentile	2.65*** [0.376]	0.48*** [0.146]	1.77*** [0.380]	2.17*** [0.153]
Manufacturing (Total factor productivity)				
11-25 percentile	1.33*** [0.136]	0.94 [0.212]	1.33* [0.220]	1.06 [0.077]
26-50 percentile	1.66*** [0.153]	0.86 [0.181]	1.58*** [0.238]	1.28*** [0.097]
51-75 percentile	2.21*** [0.198]	0.80 [0.163]	1.98*** [0.290]	1.61*** [0.124]
76-90 percentile	2.26*** [0.211]	0.87 [0.183]	2.11*** [0.322]	1.72*** [0.138]
91-100 percentile	2.54*** [0.242]	1.08 [0.231]	2.70*** [0.417]	2.33*** [0.204]

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates

significance at the 10%, 5%, 1% level, respectively. Models include all other explanatory variables in

Tables 1 and 2. The 1-10 percentile category is omitted.

Figure 1. Information requested by UAW in online applications for union organizing
(as of October 21, 2013)

Contact UAW Organizing

Do you want to know more about joining the UAW?

If so, we need some basic information so we can provide the best possible support and assistance to you and your co-workers.

ALL INFORMATION IS CONFIDENTIAL!

In order to respond all fields are required except those marked optional.

An organizer will contact you within the next ten business days. Thanks for your interest!

If you prefer, you can call the UAW Organizing Department at 1-800 2GET-UAW (1-800-243-8829). You'll be connected to (or get a call back from) a UAW organizer who can answer questions and tell you what it takes to organize a union at your workplace.

Your Name:

Your Address:

Your City:

Your State/Province: *

Email Address: *

Phone: include area code: *

Cell Phone (optional):

Best time to return your call: *

About the company

Name of your employer: *

Parent company, if any (optional):

City: *

State/Prov: *

Number of employees (best guess): *

Figure 1. Continued.

Product or Service: *

The employer's main customers: *

Does your employer have more than one workplace?: *

- Yes
 No

Are other workplaces owned by the employer organized by the UAW?: *

- Yes
 No

Are other workplaces owned by the employer organized by another union?: *

- Yes
 No

Is there another union involved at your workplace?: *

- Yes
 No

Have there been previous attempts to organize at your workplace?: *

- Yes
 No

Please describe what needs to be improved at your workplace.: *

Submit

Figure 2. The union's timing of events and decisions for an age-*a* establishment

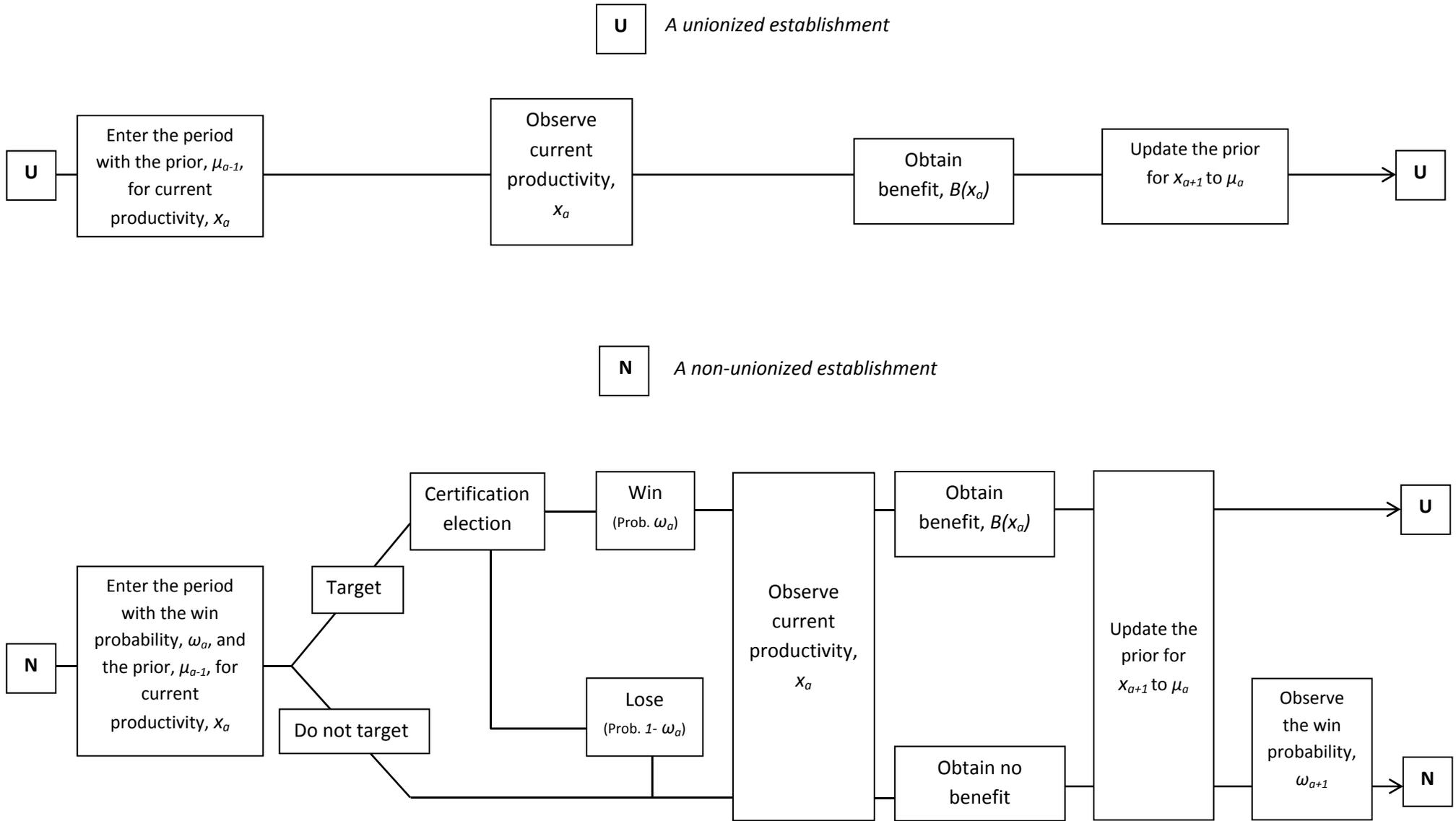
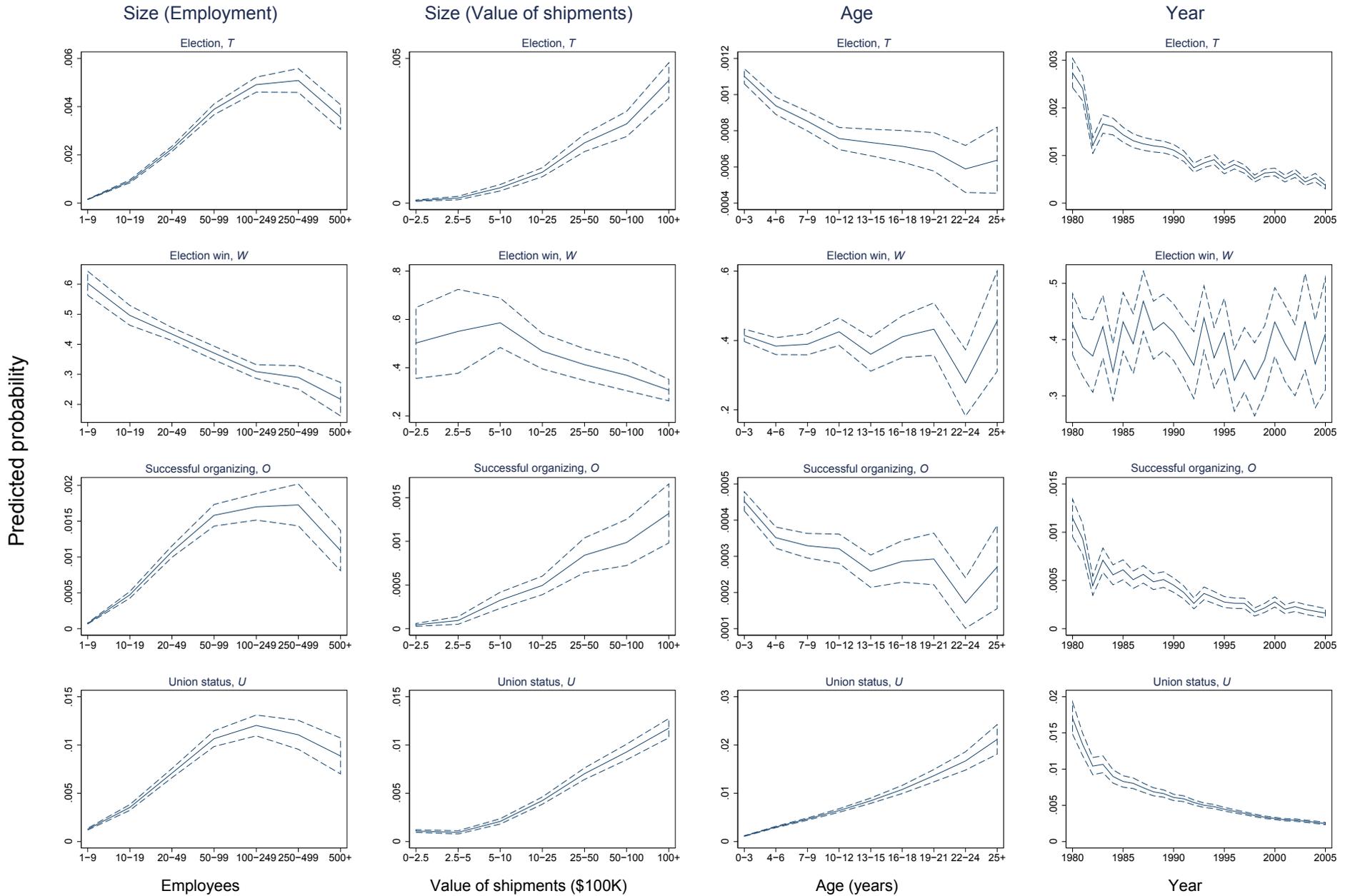
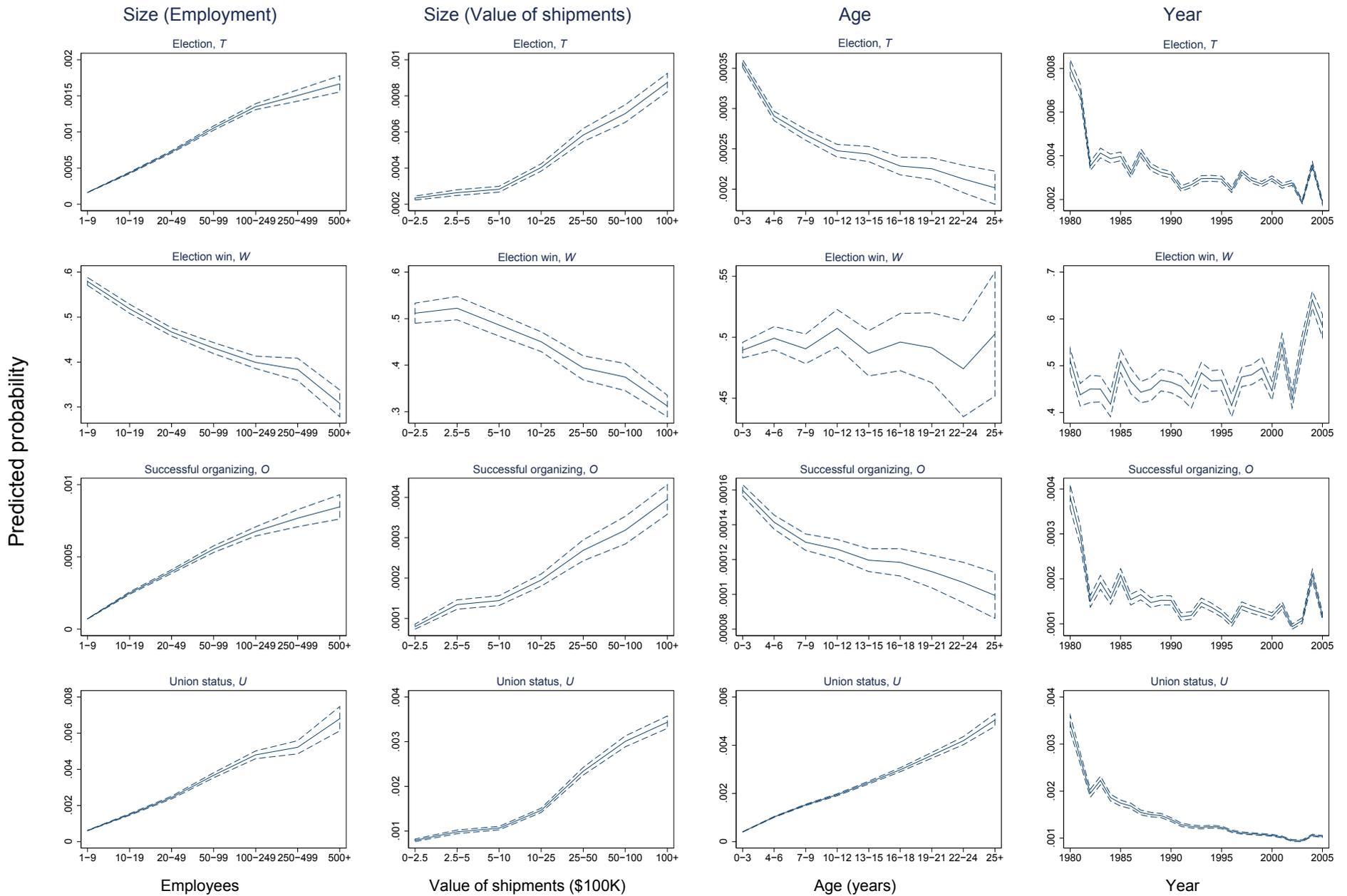


Figure 3. Predicted probabilities of union activity in manufacturing establishments -- Size, Age, and Year Effects



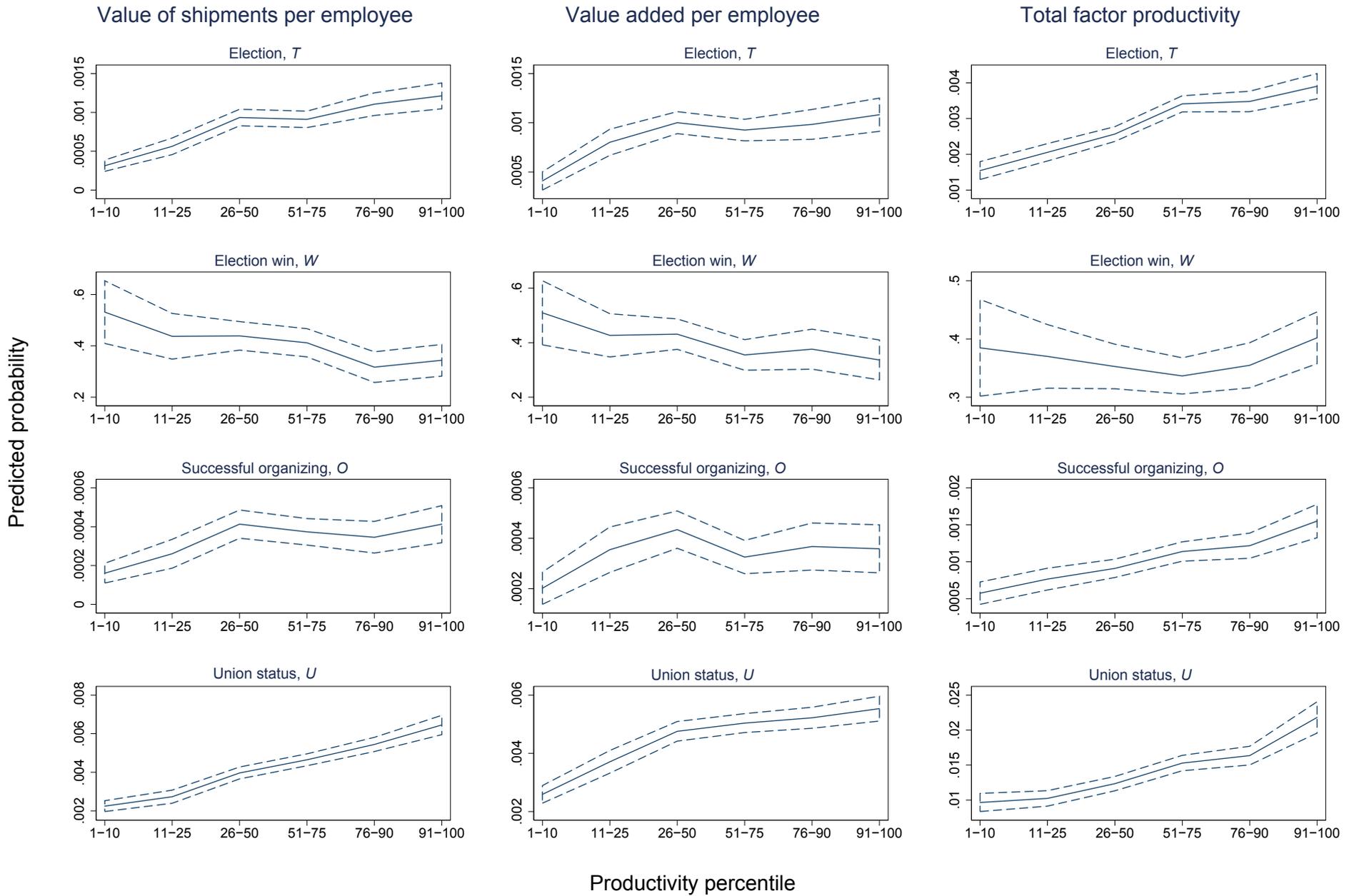
Note: Dashed curves indicate 95% pointwise confidence intervals.

Figure 4. Predicted probabilities of union activity in private sector establishments -- Size, Age, and Year Effects



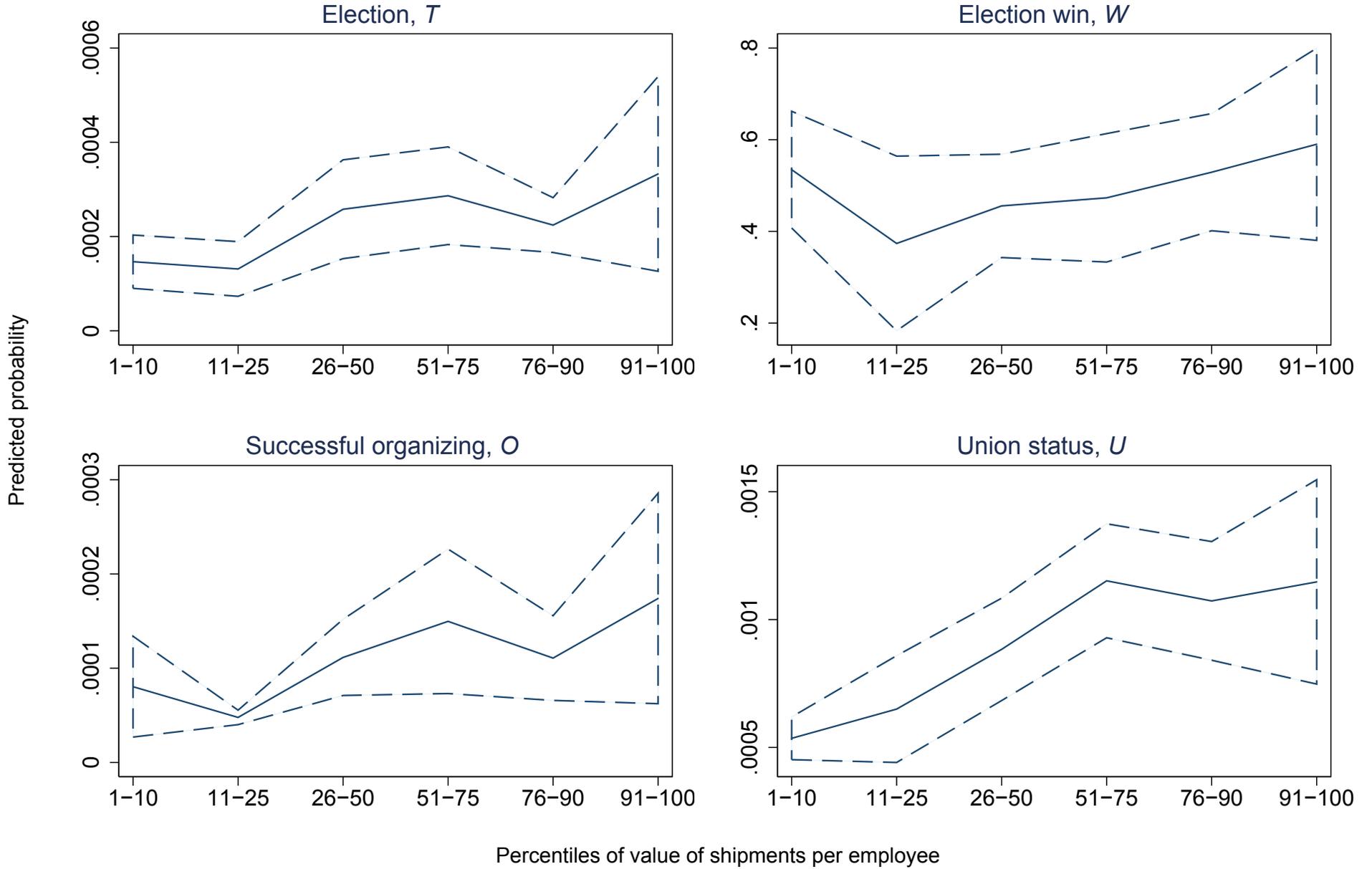
Note: Dashed curves indicate 95% pointwise confidence intervals.

Figure 5. Predicted probabilities of union activity in manufacturing establishments -- Productivity effects



Note: Dashed curves indicate 95% pointwise confidence intervals.

Figure 6. Predicted probabilities of union activity in private sector establishments -- Productivity effects



Note: Dashed curves indicate 95% pointwise confidence intervals.

Appendix

A Proofs

The following lemma will be used in the proof for Lemma 1.

Lemma 2 *Let $G(x)$ be a bounded, (non-decreasing) increasing, (strictly) convex function in x , where x is a normally distributed random variable with mean μ and variance σ^2 . $E[G(x)]$ is (non-decreasing) increasing and (strictly) convex in μ .*

Proof. Observe that

$$E[G(x)] = \int G(x) \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] dx.$$

Now,

$$\begin{aligned} E[G(x)] &= \int G(x - \mu + \mu) \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] dx \\ &= \int G(\tilde{x} + \mu) \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{\tilde{x}^2}{2\sigma^2}\right] dx, \end{aligned}$$

where $\tilde{x} = x - \mu$. It immediately follows that $E[G(x)]$ is (strictly) increasing and convex in μ by differentiating with respect to μ . ■

Proof of Lemma 1. *Existence and uniqueness of D .* Let \mathbf{T} be the operator defined by

$$D^{n+1} \equiv \mathbf{T}D^n = B(x_a) + \beta\{E[(1 - \omega_{a+1})D^n | \mu_a, a] - c\}. \quad (23)$$

\mathbf{T} has a unique fixed point, D , by the Banach fixed point theorem. To establish this result, note that \mathbf{T} satisfies Blackwell's sufficiency condition for a contraction mapping – Stokey and Lucas (1989, Theorem 3.3). The operator \mathbf{T} maps a bounded function D^n into another bounded function, D^{n+1} , because B , ω_{a+1} , and c are bounded. Similarly, if D^n is continuous in s_{a+1} , then so is $\mathbf{T}D^n$ in s_a . On this, $\mathbf{T}D^n$ can be trivially seen from (23) to be continuous in x_a and c . Note that D^n is a function of the random variables x_{a+1} and μ_{a+1} . The distribution for x_{a+1} is normal with mean μ_a and variance σ_a^2 . Recall that σ_a^2 evolves as a deterministic function of a . The distribution function for μ_{a+1} is also normal with mean μ_a and variance $(1 - \theta_a)^2\sigma_a^2$, from (6). Therefore, $E[(1 - \omega_{a+1})D^n | \mu_a, a]$ is a continuous function of μ_a , and hence so is $\mathbf{T}D^n$. To

see that Blackwell's sufficiency condition holds, note that, first, \mathbf{T} is monotone: for any two functions $D_1^n \geq D_2^n$, it follows that $\mathbf{T}D_1^n \geq \mathbf{T}D_2^n$. Second, \mathbf{T} satisfies the discounting hypothesis for any constant $b > 0$:

$$\begin{aligned}
\mathbf{T}(D^n + b) &= B(x_a) + \beta\{E[(1 - \omega_{a+1})(D^n + b)|\mu_a, a] - c\} \\
&= B(x_a) + \beta\{E[(1 - \omega_{a+1})D^n|\mu_a, a] - c\} + \beta(1 - \omega_{a+1})b \\
&= \mathbf{T}D^n + \beta(1 - \omega_{a+1})b \\
&< \mathbf{T}D^n + \beta b,
\end{aligned}$$

because $1 - \omega_{a+1} < 1$. Hence, \mathbf{T} satisfies Blackwell's sufficiency condition.

D is increasing in x_a . Observe that x_a only enters into $B(x_a)$ in (23), given μ_a . The result then immediately follows from the fact that $B(x_a)$ is increasing in x_a .

D is decreasing in c . Trivially, the function $\beta(E[(1 - \omega_{a+1})D(s_{a+1})] - c)$ is decreasing in c . Therefore, so is $D(s_a)$.

D is increasing in μ_a . Assume that D^n is non-decreasing in μ_{a+1} . It will now be shown that this implies that $\mathbf{T}D^n$ is increasing in μ_a . Again, D^n is a function of the random variables x_{a+1} and μ_{a+1} . The distribution function for x_{a+1} is normal with mean μ_a . Now, $B(x_{a+1})$ is increasing in x_{a+1} . Hence, on this account, a higher value for μ_a implies a higher value for $E[(1 - \omega_{a+1})D^n|\mu_a, a]$, because D^n is an increasing function of x_{a+1} – Lemma 2. Likewise, μ_{a+1} is normally distributed with mean μ_a . Therefore, on this account, $\mathbf{T}D^n$ is non-decreasing in μ_a , and D^n is non-decreasing in μ_{a+1} – again, Lemma 2. Putting both pieces together implies that $\mathbf{T}D^n$ is increasing in μ_a . Consequently, \mathbf{T} maps non-decreasing functions of μ_a into increasing ones. By Stokey and Lucas (1989, Theorem 3.2, Corollary 1) the fixed point D must then be increasing in μ_a .

Convexity of D in μ_a and x_a . It is easy to see that D^{n+1} is strictly convex in x_a because $B(x_a)$ is strictly convex in x_a and $\beta\{E[(1 - \omega_{a+1})D^n|\mu_a, a] - c\}$ does not depend on x_a , given μ_a . Suppose now that D^n is a convex function of μ_a . Consider two priors, μ_1 and μ_2 . Let $\mu_\lambda = \lambda\mu_1 + (1 - \lambda)\mu_2$, for $\lambda \in (0, 1)$. Convexity of $\mathbf{T}D^n$ requires $\lambda(\mathbf{T}D^n)(\mu_1, a, x_a, \omega_{a+1}) + (1 - \lambda)(\mathbf{T}D^n)(\mu_2, a, x_a, \omega_{a+1}) \geq (\mathbf{T}D^n)(\mu_\lambda, a, x_a, \omega_{a+1})$. Now, again D^n is a function of the random variables x_{a+1} and μ_{a+1} . The distribution of x_{a+1} is normal with mean μ_a and variance σ_a^2 ,

while the distribution of μ_{a+1} is also normal with mean μ_a and variance $(1 - \theta_a)^2 \sigma_a^2$. Note that

$$\begin{aligned}
& \lambda(\mathbf{T}D^n)(\mu_1, a, x_a, \omega_{a+1}) + (1 - \lambda)(\mathbf{T}D^n)(\mu_2, a, x_a, \omega_{a+1}) \\
&= \lambda\{B(x_a) + \beta E[(1 - \omega_{a+2})D^n(\mu_{a+1}, a + 1, x_{a+1}, \omega_{a+2})|\mu_1, a] - \beta c\} \\
&+ (1 - \lambda)\{B(x_a) + \beta E[(1 - \omega_{a+2})D^n(\mu_{a+1}, a + 1, x_{a+1}, \omega_{a+2})|\mu_2, a] - \beta c\} \\
&> B(x_a) + \beta E[(1 - \omega_{a+2})D^n(\mu_{a+1}, a + 1, x_{a+1}, \omega_{a+2})|\mu_\lambda, a] - \beta c \\
&= (\mathbf{T}D^n)(\mu_\lambda, a, x_a, \omega_{a+1}).
\end{aligned}$$

The inequality follows from the facts that D^n is convex in μ_{a+1} and strictly convex in x_{a+1} and an application of Lemma 2. Thus, \mathbf{T} maps convex functions into strictly convex functions. Therefore, D is strictly convex in μ_a —Stokey and Lucas (1989, Theorem 3.2, Corollary 1).

D is decreasing in a. From above, $D(\mu_{a+1}, a + 1, x_{a+1}, \omega_{a+2})$ is a bounded, increasing, strictly convex function of the random variables x_{a+1} and μ_{a+1} . The random variable x_{a+1} is normally distributed with mean μ_a and variance σ_a^2 , while μ_{a+1} is normally distributed with mean μ_a and variance $(1 - \theta_a)^2 \sigma_a^2$. Now, as can be seen from (3) and (5), σ_a^2 decreases with age, a . Therefore, an increase in a , ceteris paribus, amounts to a mean-preserving shrinkage in x_{a+1} and μ_{a+1} . As a result, $E[D(\mu_{a+1}, a + 1, x_{a+1}, \omega_{a+2})|\mu_{a-1}, a]$ is decreasing in a by Hadar and Russell (1971, Theorem 3).

$E[D|\mu_{a-1}, a]$ is increasing in μ_{a-1} and decreasing in a . Prior to observing x_a , the union will take μ_a and x_a to be normally distributed random variables with mean μ_{a-1} . From the parts above, $D(\mu_a, a, x_a, \omega_{a+1})$ is increasing in both μ_a and x_a , decreasing in a and strictly convex in μ_a and x_a . Consequently, $E[D(\mu_a, a, x_a, \omega_{a+1})|\mu_{a-1}, a]$ is increasing in μ_{a-1} and decreasing in a .

■

Proof of Proposition 1. By definition, $T(\mu_{a-1}, a) = 1 - \Gamma(c/E[D(\mu_a, a, x_a, \omega_{a+1})|\mu_{a-1}, a])$. First, note that Γ is a decreasing function of $E[D(\mu_a, a, x_a, \omega_{a+1})|\mu_{a-1}, a]$, since Γ is a c.d.f. Therefore, $1 - \Gamma$ is an increasing function of $E[D(\mu_a, a, x_a, \omega_{a+1})|\mu_{a-1}, a]$. By Lemma 1, this last expectation is increasing in μ_{a-1} and decreasing in a . Therefore, so is T . ■

Proof of Proposition 2. By definition, $W(\mu_{a-1}, a) = E[\omega_a | \omega_a > \tilde{\omega}(\mu_{a-1}, a)]$. But the definition of $\tilde{\omega}(\mu_{a-1}, a)$ in (12) and Lemma 1 imply that $\tilde{\omega}(\mu_{a-1}, a)$ is decreasing in μ_{a-1} and increasing in a . Therefore, $W(\mu_{a-1}, a)$ is also decreasing in μ_{a-1} and increasing in a . ■

Proof of Proposition 3. By the definition of U in (16), U is increasing in μ_{a-1} if T is. But, T is increasing in μ_{a-1} by Proposition 1. Therefore, so is U . Moreover, U is increasing in a

because $U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) - U(\mu_{a-2}, \dots, \mu_0, a-1) = \{\prod_{k=1}^a [1 - O(\mu_{k-1}, k)]\} O(\mu_{a-1}, a) > 0$. ■

Proof of Proposition 4. The proof is in two parts. First, suppose that the c.d.f. governing the observer's beliefs $\Omega(\mu_{a-1}|x_a, a)$ is increasing in x_a in the sense of first-order stochastic dominance. Now, $T(\mu_{a-1}, a)$ is increasing in μ_{a-1} by Proposition 1. The integral in (17) is then increasing in x_a – Hadar and Russell (1971, Theorem 1). As a consequence, $T^o(x_a, a)$ is increasing in x_a .

Second, it will now be established that $\Omega(\mu_{a-1}|x_a, a)$ is increasing in x_a in the sense of first-order stochastic dominance. Let $\psi(\mu_{a-1}|x_a, a)$ be the density function associated with Ω . Bayes' Rule implies

$$\psi(\mu_{a-1}|x_a, a) = \frac{\phi(x_a|\mu_{a-1}, a)\psi(\mu_{a-1}|a)}{\phi(x_a|a)}, \quad (24)$$

where $\psi(\mu_{a-1}|a) = \int \psi(\mu_{a-1}|x_a, a)dx_a$, $\phi(x_a|\mu_{a-1}, a)$ is the density associated with $\Phi(x_a|\mu_{a-1}, a)$, and $\phi(x_a|a) = \int \phi(x_a|\mu_{a-1}, a)d\mu_{a-1}$. First, it will be shown that $\psi(\mu_{a-1}|x_a, a)$ satisfies the monotone likelihood ratio property (MLRP). The MLRP is satisfied strictly if, given $x_2 > x_1$, the following inequality holds

$$\psi(\mu_{a-1}|x_2, a)\psi(\mu'_{a-1}|x_1, a) - \psi(\mu_{a-1}|x_1, a)\psi(\mu'_{a-1}|x_2, a) > 0, \text{ for } \mu_{a-1} > \mu'_{a-1}. \quad (25)$$

[See, e.g., Karlin and Rubin (1956), equation (2)]. For differentiable density functions, (25) implies

$$\frac{d}{d\mu_{a-1}} \frac{\psi(\mu_{a-1}|x_2, a)}{\psi(\mu_{a-1}|x_1, a)} > 0, \quad (26)$$

assuming that $\psi(\mu_{a-1}|x_1, a) \neq 0$ (which will be satisfied for a normal density). Using the definition of ψ in (24), rewrite the sufficient condition for the MLRP (26) as

$$\frac{d}{d\mu_{a-1}} \left[\frac{\phi(x_1|a)\phi(x_2|\mu_{a-1}, a)}{\phi(x_2|a)\phi(x_1|\mu_{a-1}, a)} \right] = \frac{\phi(x_1|a)}{\phi(x_2|a)} \frac{d}{d\mu_{a-1}} \left[\frac{\phi(x_2|\mu_{a-1}, a)}{\phi(x_1|\mu_{a-1}, a)} \right] > 0. \quad (27)$$

Now, ϕ is the density of a normal random variable with mean μ_{a-1} and variance σ_{a-1}^2 . Therefore,

$$\begin{aligned} \frac{d}{d\mu_{a-1}} \left[\frac{\phi(x_2|\mu_{a-1}, a)}{\phi(x_1|\mu_{a-1}, a)} \right] &= \frac{1}{\sqrt{2\pi}\sigma_{a-1}} \frac{d}{d\mu_{a-1}} \exp\left[\frac{(x_1 - \mu_{a-1})^2 - (x_2 - \mu_{a-1})^2}{2\sigma_{a-1}^2} \right] \\ &= \frac{1}{\sqrt{2\pi}\sigma_{a-1}} \frac{(x_2 - x_1)}{\sigma_{a-1}^2} \exp\left[\frac{(x_1 - \mu_{a-1})^2 - (x_2 - \mu_{a-1})^2}{2\sigma_{a-1}^2} \right] > 0, \end{aligned}$$

where the inequality follows because $x_2 > x_1$. Thus, $\psi(\mu_{a-1}|x_a, a)$ satisfies the MLRP strictly. This implies that Ω is increasing in x_a in the sense of first-order stochastic dominance – Milgrom (1981). ■

Proof of Proposition 5. Note that

$$\begin{aligned} W^o(x_a, a) &= \int W(\mu_{a-1}, a) d\Omega(\mu_{a-1}|x_a, a) \\ &= - \left[\int (-W(\mu_{a-1}, a)) d\Omega(\mu_{a-1}|x_a, a) \right] \end{aligned} \quad (28)$$

By Proposition 2, $W(\mu_{a-1}, a)$ is decreasing in μ_{a-1} . Therefore, $-W(\mu_{a-1}, a)$ is increasing in μ_{a-1} . Furthermore, $\Omega(\mu_{a-1}|x_a, a)$ is increasing in x_a in the sense of first-order stochastic dominance, as shown in the proof of Proposition 4. Consequently, the integral inside the brackets (28) is increasing in x_a – Hadar and Russell (1971, Theorem 1). It follows that $W^o(x_a, a)$ is decreasing in x_a . ■

Proof of Proposition 6. U^o is increasing in x_a . Observe that $U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a)$ is increasing in μ_{m-1} , for $m = 1, \dots, a$, because from (16)

$$\frac{dU(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a)}{d\mu_{m-1}} = \prod_{j=1, j \neq m}^a [1 - \omega_j T(\mu_{j-1}, j)] \omega_m \frac{dT(\mu_{m-1}, m)}{d\mu_{m-1}} > 0.$$

The sign of the expression follows from Proposition 1. Next, let $\xi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0|x_a, a)$ be the density function for the sequence of priors $(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0)$ conditional on x_a and a . This density can be expressed in terms of a product of one-step conditional densities

$$\xi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0|x_a, a) = \psi(\mu_{a-1}|x_a, a) \zeta(\mu_{a-2}|\mu_{a-1}, a) \cdots \zeta(\mu_1|\mu_2, a),$$

where $\zeta(\mu_{m-2}|\mu_{m-1}, m)$ is the density of μ_{m-2} conditioned on μ_{m-1} and m . The form of the above expression is justified from (6). Note that $\mu_0 = \bar{\chi}$ is fixed (non-random). Therefore,

$$\begin{aligned} U^o(x_a, a) &= \int U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \xi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0|x_a, a) d\mu_{a-1} \cdots d\mu_1 d\Gamma(\omega_{a-1}) \cdots d\Gamma(\omega_0) \\ &= \int I_{a-2}(\mu_{a-1}, a) \psi(\mu_{a-1}|x_a, a) d\mu_{a-1} d\Gamma(\omega_{a-1}) \cdots d\Gamma(\omega_0), \end{aligned}$$

where

$$I_m(\mu_{m+1}, a) = \int I_{m-1}(\mu_m, a) \zeta(\mu_m|\mu_{m+1}, a) d\mu_m, \text{ for } m = 2, \dots, a-2,$$

and

$$I_1(\mu_2, a) = \int U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \zeta(\mu_1|\mu_2, a) d\mu_1.$$

Suppose $I_{m-1}(\mu_m, a)$ is increasing in μ_m . Then, $I_m(\mu_{m+1}, a)$ is increasing in μ_{m+1} . This occurs because the c.d.f. for μ_m is increasing in μ_{m+1} , in the sense of first-order stochastic dominance,

by an argument similar to that employed in the proof of Proposition 4.⁵⁹ To start the induction hypothesis off, note that $I_1(\mu_2, a)$ will be increasing in μ_2 , because U is strictly increasing in μ_2 and the c.d.f. associated with $\zeta(\mu_1|\mu_2, a)$ is stochastically increasing in μ_2 (in the sense of first-order stochastic dominance).

U^o is increasing in a . Note that for all x_a

$$\begin{aligned} U^o(x_a, a+1) - U^o(x_a, a) &= \\ &= \int U(\mu_a, \mu_{a-2}, \dots, \mu_0, a+1) \varphi(\mu_a, \mu_{a-2}, \dots, \mu_0 | x_a, a+1) d\mu_a d\mu_{a-1} \dots d\mu_1 d\Gamma(\omega_a) \dots d\Gamma(\omega_0) \\ &\quad - \int U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \varphi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x_a, a) d\mu_{a-1} \dots d\mu_1 d\Gamma(\omega_{a-1}) \dots d\Gamma(\omega_0). \end{aligned}$$

Using the definition of $U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a)$

$$\begin{aligned} &U(\mu_a, \mu_{a-2}, \dots, \mu_0, a+1) - U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \\ &= \prod_{k=1}^a [1 - O(\mu_{k-1}, k)] O(\mu_a, a+1) > 0. \end{aligned} \tag{30}$$

Also,

$$\varphi(\mu_a, \mu_{a-2}, \dots, \mu_0 | x_a, a+1) = \rho(\mu_a | \mu_{a-1}) \varphi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x_a, a),$$

where ρ is the density of μ_a conditional on μ_{a-1} . Therefore,

$$\begin{aligned} U^o(x, a+1) - U^o(x, a) &= \\ &= \int U(\mu_a, \mu_{a-2}, \dots, \mu_0, a+1) \rho(\mu_a | \mu_{a-1}) \varphi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x, a) d\mu_a d\mu_{a-1} \dots d\mu_1 d\Gamma(\omega_a) \dots d\Gamma(\omega_0) \\ &\quad - \int U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \varphi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x, a) d\mu_{a-1} \dots d\mu_1 d\Gamma(\omega_{a-1}) \dots d\Gamma(\omega_0) \\ &= \int \left[\int U(\mu_a, \mu_{a-2}, \dots, \mu_0, a+1) \rho(\mu_a | \mu_{a-1}) d\mu_a d\Gamma(\omega_a) - U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \right] \\ &\quad \times \varphi(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0 | x, a) d\mu_{a-1} \dots d\mu_1 d\Gamma(\omega_{a-1}) \dots d\Gamma(\omega_0). \end{aligned}$$

⁵⁹To see this, let $\Upsilon(\mu_{a-1}|\mu_a, a)$ represent the cdf that is associated with the density function $\zeta(\mu_{a-1}|\mu_a, a)$. Establishing MLRP for $\zeta(\mu_{a-1}|\mu_a, a)$ is equivalent to showing

$$\text{sign} \frac{d}{d\mu_{a-1}} \left[\frac{\xi(\mu_{a-1}|\mu'_a, a)}{\xi(\mu_{a-1}|\mu_a, a)} \right] = \text{sign} \frac{d}{d\mu_{a-1}} \left[\frac{\rho(\mu'_a|\mu_{a-1}, a)}{\rho(\mu_a|\mu_{a-1}, a)} \right] > 0, \tag{29}$$

for any $\mu'_a > \mu_a$, where $\rho(\mu_a|\mu_{a-1}, a)$ is the density of μ_a conditional on μ_{a-1} and a -follow steps similar to those used in the proof of Proposition 4. The derivation of equation (29) parrots that of (27). Note from (6) that $\rho(\mu_a|\mu_{a-1}, a)$ is the density of a normal random variable with mean μ_{a-1} and variance $(1 - \theta_a)^2 (\sigma_{\zeta_{a-1}}^2 + \sigma_\varepsilon^2)$. By mimicing the argument outlined in the proof of Proposition 4, it can be shown that (29) holds. It follows that $\Upsilon(\mu_{a-1}|\mu_a, a)$ is increasing in μ_a (in the sense of first-order stochastic dominance).

The last expression is positive if the term in brackets is positive, or if

$$\int U(\mu_a, \mu_{a-2}, \dots, \mu_0, a+1) \rho(\mu_a | \mu_{a-1}) d\mu_a d\Gamma(\omega_a) > U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a). \quad (31)$$

But note that

$$\begin{aligned} \int U(\mu_a, \mu_{a-2}, \dots, \mu_0, a+1) \rho(\mu_a | \mu_{a-1}) d\mu_a d\Gamma(\omega_a) &> \int U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a) \rho(\mu_a | \mu_{a-1}) d\mu_a \\ &= U(\mu_{a-1}, \mu_{a-2}, \dots, \mu_0, a), \end{aligned}$$

where the inequality follows from (30). Therefore, U^o is increasing in a . ■

B Derivation of $B(x_a)$ Under Alternative Models

This Appendix shows that the union benefit function, $B(x_a)$, in Assumption 1 is an increasing and strictly convex function under many specifications of the monopoly union, right-to-manage and efficient bargaining paradigms of union and establishment conduct.

B.1 Setup

Consider a setting where an establishment produces output, o , according to the standard production function,

$$o(l; x_a) = e^{x_a} l^\alpha, \quad 0 < \alpha < 1,$$

where x_a drives total factor productivity, e^{x_a} , and l is the amount of labor hired. This formulation for total productivity is standard when productivity shocks are assumed to be normally distributed. Endow the union with the utility function

$$u(w, l; w_c) = (w - w_c)^\delta l^\gamma, \quad \delta, \gamma > 0, \quad (32)$$

where w is the union wage rate and w_c is the fixed competitive wage rate non-union establishments pay. The union values a high wage premium, $w - w_c$, and a high employment (which equates with union membership). The objective function (32) is a Stone-Geary type utility function.⁶⁰ Variants of (32) are frequently used to model union preferences. For instance, Dunlop (1944) proposes the wage bill as the union's objective function, $u(w, l) = wl$, which is a

⁶⁰The general form of the Stone-Geary utility function is $(w - w_c)^\delta (l - \underline{l})^\gamma$, $\underline{l} \geq 0$. The function in (32) sets $\underline{l} = 0$, i.e. the union desires any positive employment. Note that setting $\underline{l} > 0$ would trivially imply that the union does not organize small firms that cannot provide the union an employment of at least \underline{l} .

special case of (32) with $\delta = \gamma = 1$ and $w_c = 0$. Rosen (1969), Calvo (1978), Oswald (1982), and Manning (1987, 1994) use rent maximization as the objective, $u(w, l; w_c) = (w - w_c)l$, which is another special case of (32) with $\delta = \gamma = 1$. Another frequently used objective function is the utilitarian one, $u(w, l) = lh(w)$, where $h(w)$ is a strictly concave function. This formulation implies that the union cares about the total utility of its l members, and corresponds to (32) with $\gamma = 1$ and $w_c = 0$ when $h(w) = w^\delta$, $\delta < 1$, a standard concave function. Finally, note that the version of (32) with $\gamma < 1$, $\delta = 1 - \gamma$, and $w_c = 0$ is the familiar Cobb-Douglas utility function.

B.2 Monopoly Union Model

In the monopoly union model, the union picks w , and then the establishment chooses l .⁶¹ The establishment's problem is

$$\max_l \{e^{x_a} l^\alpha - wl\},$$

which yields a demand for labor given by

$$l^* = L(w, x_a) = \left[\frac{\alpha e^{x_a}}{w}\right]^{1/(1-\alpha)}. \quad (33)$$

The union's problem is

$$\max_w \{(w - w_c)^\delta \left[\frac{\alpha e^{x_a}}{w}\right]^{\gamma/(1-\alpha)}\}. \quad (34)$$

The first order condition for this problem is

$$\delta(w - w_c)^{\delta-1} \left[\frac{\alpha e^{x_a}}{w}\right]^{\gamma/(1-\alpha)} + (w - w_c)^\delta \frac{\gamma}{1-\alpha} \left[\frac{\alpha e^{x_a}}{w}\right]^{\gamma/(1-\alpha)} \left(-\frac{1}{w}\right) = 0,$$

which has the solution

$$w^* = \frac{\gamma}{\gamma - \delta(1-\alpha)} w_c,$$

provided that $\gamma - \delta(1-\alpha) > 0$, which is the condition for an interior solution w^* to exist.⁶²

Note that w^* is not a function of x_a . Plugging the expression for w^* back into the union's objective function yields

$$B(x_a) = (w^* - w_c)^\delta \left[\frac{\alpha}{w^*}\right]^{\gamma/(1-\alpha)} e^{\frac{\gamma}{1-\alpha} x_a},$$

⁶¹See Oswald (1982) for the monopoly union model. For a general exposition and discussion of all three models discussed, see also Manning (1987, 1994).

⁶²The second-order condition associated with the maximization problem in (34) is also satisfied for the given parameter restriction.

which is increasing and strictly convex in x_a .

B.3 Right-to-Manage Model

Consider now the case where the establishment is free to choose l (hence, right-to-manage), given the union wage, w , but where w is determined via Nash Bargaining.⁶³ Once again, l will be determined by (33). The bargaining problem reads

$$\max_w \{(e^{x_a} l^\alpha - wl)^{(1-\phi)} [(w - w_c)^\delta l^\gamma]^\phi\}, \text{ for } 0 < \phi < 1,$$

subject to (33). The objective function weights the establishment's profits and the union's objective function, where the weight ϕ reflects the union's bargaining power. Differentiate the objective function with respect to w , while making use of the fact that the establishment has chosen l to maximize its profits, to obtain

$$(1 - \phi)l(w - w_c)^\delta - 1 = \phi(1 - \alpha)(w - w_c)^{\delta-1} e^{x_a} l^\alpha [(1 - \gamma) + \gamma(w - w_c)l^{-1} (\alpha e^{x_a})^{1/(1-\alpha)} w^{-1/(1-\alpha)-1}]. \quad (35)$$

The solution is

$$w^* = \frac{\phi\gamma + (1 - \phi)\alpha}{\phi[\gamma - \delta(1 - \alpha)] + (1 - \phi)\alpha} w_c, \quad (36)$$

provided that $\phi[\gamma - \delta(1 - \alpha)] + (1 - \phi)\alpha > 0$, which guarantees an interior solution, w^* .⁶⁴ Note, again, that w^* does not depend on x_a . Therefore, $B(x_a)$ now reads

$$B(x_a) = (w^* - w_c)^\delta \left[\frac{\alpha}{w^*} \right]^{\gamma/(1-\alpha)} e^{\frac{\gamma}{1-\alpha} x_a},$$

which is strictly convex in x_a .

B.4 Efficient Bargaining Model

Finally, consider efficient bargaining.⁶⁵ Both l and w are chosen simultaneously via Nash Bargaining to solve the maximization problem

$$\max_{w,l} \{(e^{x_a} l^\alpha - wl)^{(1-\phi)} [(w - w_c)^\delta l^\gamma]^\phi\}, \text{ for } 0 < \phi < 1.$$

⁶³See Nickell (1982) for the right-to-manage model.

⁶⁴The second-order condition must also be satisfied for w^* to be a maximizer. Note, however, that the derived properties of $B(x_a)$ hold at *any* interior solution w^* .

⁶⁵See MacDonald and Solow (1981) for the efficient bargaining model.

The first-order conditions for w and l , respectively, read

$$\phi\delta(e^{x_a}l^\alpha - wl) - (w - w_c)(1 - \phi)l = 0, \quad (37)$$

$$\phi\gamma(e^{x_a}l^\alpha - wl) + (1 - \phi)(\alpha e^{x_a}l^\alpha - wl) = 0. \quad (38)$$

Adding the two equations together and rearranging yields the following relationship between the optimal choices for w and l

$$l^* = \left\{ \frac{[\phi + (1 - \phi)\alpha]e^{x_a}}{[2(1 - \phi) - \phi]w^* - (1 - \phi)w_c} \right\}^{1/(1-\alpha)}. \quad (39)$$

Furthermore, (38) implies

$$w^* = \frac{\phi\gamma + (1 - \phi)\alpha}{\phi\gamma + 1 - \phi} e^{x_a} l^{*\alpha-1}. \quad (40)$$

Substituting (39) into (40) and rearranging yields

$$w^* = \frac{(\phi\gamma + (1 - \phi)\alpha)(1 - \phi)}{[(\phi\gamma + (1 - \phi)\alpha)(2(1 - \phi) - \phi) - (\phi\gamma + 1 - \phi)(\phi + (1 - \phi)\alpha)]} w_c,$$

with the condition that $(\phi\gamma + (1 - \phi)\alpha)(2(1 - \phi) - \phi) - (\phi\gamma + 1 - \phi)(\phi + (1 - \phi)\alpha) > 0$, which, again, ensures an interior solution, w^* .⁶⁶ Once again, w^* does not depend on x_a . One can then write

$$B(x_a) = (w^* - w_c)^\delta \left\{ \frac{(\phi + (1 - \phi)\alpha)}{[2(1 - \phi) - \phi]w - (1 - \phi)w_c} \right\}^{\gamma/(1-\alpha)} e^{\frac{\gamma}{1-\alpha}x_a},$$

which is a strictly convex function of x_a .

C Data

C.1 Trends in Certification and Decertification Elections

The raw data from the NLRB was benchmarked against the published NLRB Annual Reports, and the agreement is quite high, as is shown in Figure C.1 (left axis: certification elections, right axis: decertification elections). The NLRB Annual Reports aggregate the certification elections (RC) and the employer-requested elections (RM). The total number of certification elections is about 7,000 per year for the period 1978-1980. Then, in 1981, it drops to 6,000, and to about 3,500 in 1982. It remains relatively flat until 1992, and drops

⁶⁶Again, the second order condition must be satisfied for w^* to be a maximizer. The derived properties of $B(x_a)$ hold at *any* interior solution w^* .

to 3,000 per year. A further fall occurs during the 2000's. The raw data show basically the same trend as in the published annual reports, with two exceptions. One is a clear instance of a coverage gap in the year 1977: the raw data on NLRB certification elections contain only 4,500 elections rather than the nearly 9,000 in the NLRB Annual Report. The other occurs when the data series switches from the data for 1977-1999 to the one that was downloaded directly from `www.data.gov` for the period 1999-2007, with the greatest dip for the year 1997.

Figure C.1 also shows the total number of decertification elections (RD) in the raw data and the NLRB Annual Reports. These elections occur at a rate of around 800 – 900 per year for the 1977-1986 period. This rate then shows a clear, gradual downward trend until 1997, when it levels at about 400 per year. Similar to the certification elections, there is a clear coverage gap in the year 1977.

Figure C.2 plots the union win rate in the NLRB Annual Reports and the raw data used for empirical analysis, again combining certification elections with employer-requested elections for comparability. The rate at which unions win both certification and decertification elections are basically flat until the year 1987. Unions consistently win about 47% of certification elections with a slight dip during 1981-1982, and lose about 75% of decertification elections. After 1987, the rate at which unions lose decertification elections falls to 71% and trends downward to around 65%. Starting in 1987, the rate at which unions win certification elections increases to about 50%, where it remains until 1995, with a slight dip during the 1990-1991 recession. Thereafter, there is a slight drop around 1996, but then the rate at which unions win certification elections rises to nearly 60% by 2007. The NLRB raw data appear to slightly understate the increase in the election win rate in recent years relative to the published report.

C.2 Statistics on Matching

The match rate for certification and decertification elections over time are shown in Figure C.3. Considering both types of elections, the minimum match rate declines from about 80% in 1977 to 71% in 1986. It is then stable until about 1994, when it starts trending downward, reaching a low of 65% in 2000. Then, in 2001 it sharply returns to about 70%. This discontinuity corresponds roughly with a change in the structuring of the source data for the LBD that resulted in more complete and comprehensive source data. The trends in the match rate

are similar for certification and decertification elections, and do not vary substantially by the election outcome, although match rates are lower in the case of decertification.

In Table C.1, the rates of match for certification elections are shown by NLRB sector, which differ from the SIC sector definitions in general. For those sectors with a substantial number of elections (Manufacturing, Retail Trade, Services, Trade, Transportation, and Utilities, and Wholesale Trade), the match rates range from 70% to 77%. Construction and Manufacturing elections have higher match rates, Wholesale Trade elections match to the LBD somewhat less frequently, and Retail Trade, Services, and Trade, Transportation, and Utilities have lower match rates.

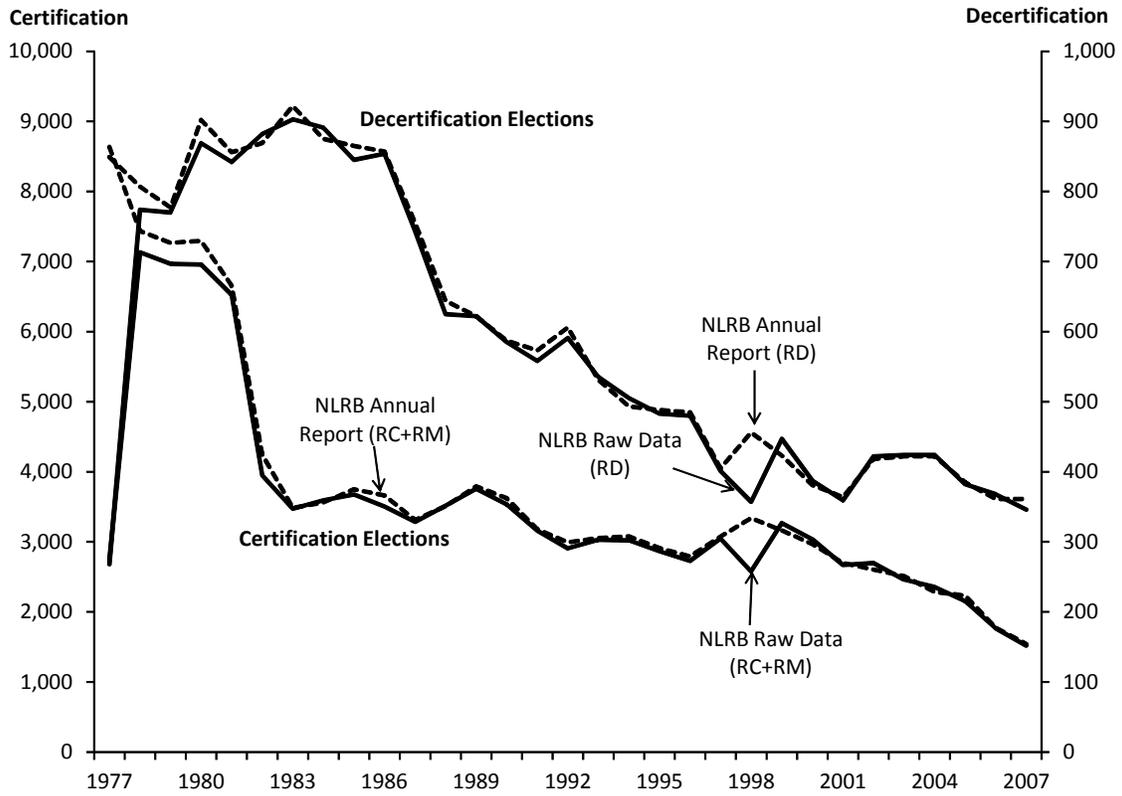


Figure C.1: The number of certification (RC+RM) and decertification (RD) elections – NLRB Annual Report versus NLRB Raw Data

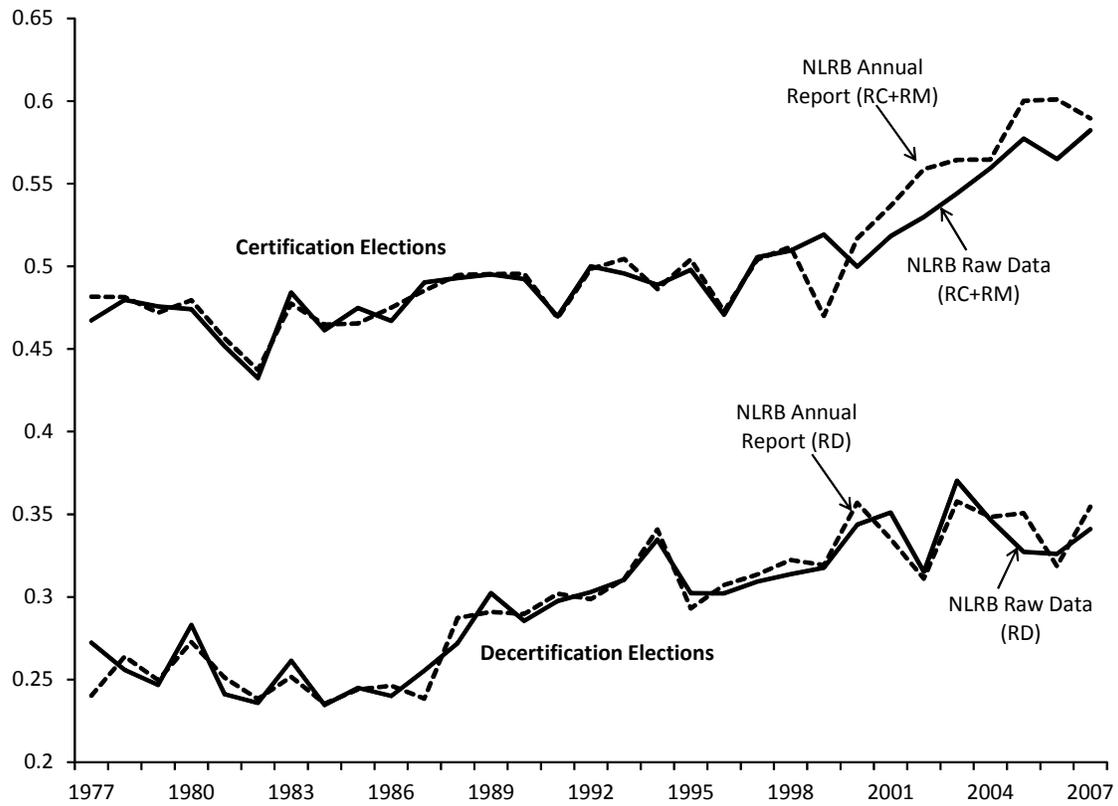


Figure C.2: Union win rate in certification (RC+RM) and decertification (RD) elections – NLRB Annual Report versus NLRB Raw Data

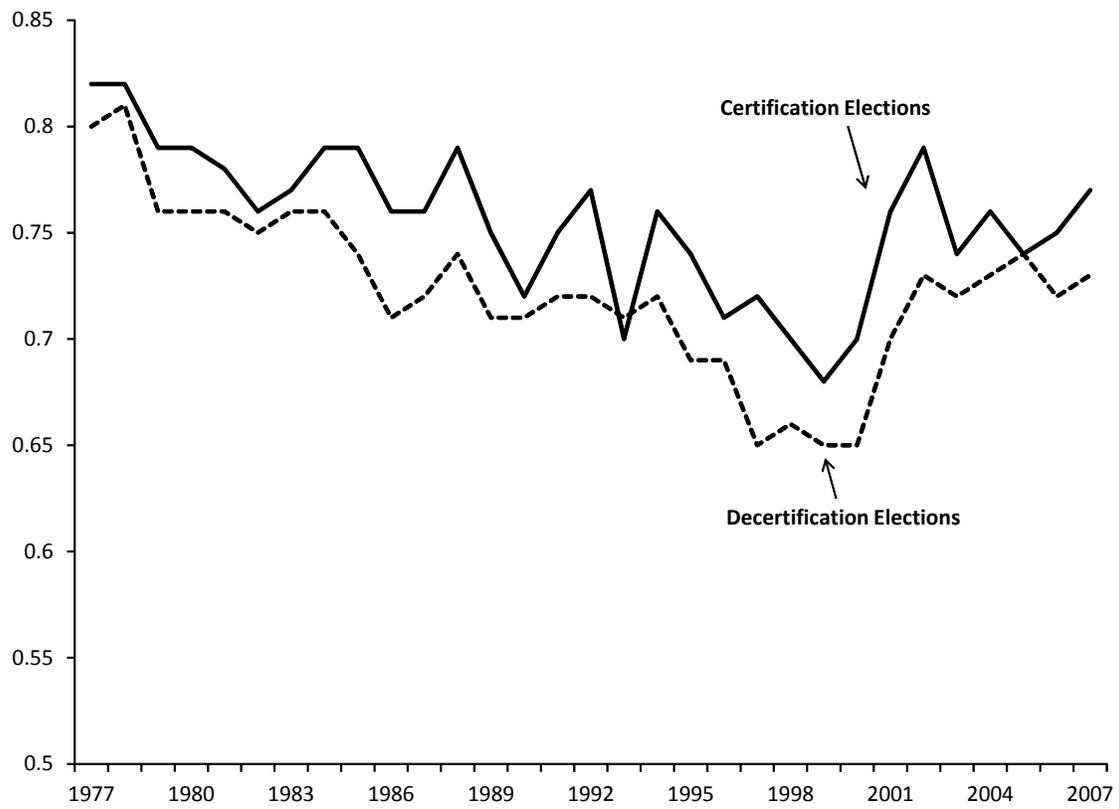


Figure C.3: Match rates by year – All sectors

TABLE C.1. Match rates for selected sectors – 1977-2007

NLRB Sector Name	Number of Certification Elections	Match Rate
Construction	7,380	76.9%
Manufacturing	34,496	77.6%
Retail Trade	8,809	70.0%
Services	26,123	68.8%
Trade, Transportation, and Utilities	15,616	70.3%
Wholesale Trade	7,246	74.7%
All other sectors	3,394	70.1%
Total (All sectors)	103,064	73.1%

D Robustness Analysis and Additional Results

TABLE D.1. Odds ratios based on logit model estimates – All Sectors
(Sample restricted to the 1990-2007 period)

Event:	Certification:		Successful Organizing	Union Status
	Election	Win		
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	2.34*** [0.041]	0.76*** [0.029]	2.98*** [0.076]	2.28*** [0.043]
20-49 employees	3.67*** [0.063]	0.62*** [0.025]	4.53*** [0.115]	3.71*** [0.078]
50-99 employees	5.46*** [0.113]	0.49*** [0.025]	6.29*** [0.194]	5.66*** [0.141]
100-249 employees	7.18*** [0.157]	0.43*** [0.024]	7.90*** [0.261]	7.60*** [0.212]
250-499 employees	8.16*** [0.276]	0.41*** [0.036]	9.16*** [0.456]	8.31*** [0.333]
500+ employees	9.41*** [0.396]	0.28*** [0.031]	10.53*** [0.650]	10.91*** [0.590]
4-6 years	0.84*** [0.012]	0.99 [0.035]	0.91*** [0.020]	2.49*** [0.027]
7-9 years	0.79*** [0.013]	0.97 [0.038]	0.86*** [0.021]	3.61*** [0.049]
10-12 years	0.74*** [0.014]	1.04 [0.047]	0.84*** [0.023]	4.70*** [0.069]
13-15 years	0.72*** [0.015]	0.97 [0.048]	0.80*** [0.025]	6.07*** [0.095]
16-18 years	0.68*** [0.018]	1.02 [0.062]	0.80*** [0.029]	7.45*** [0.131]
19-21 years	0.67*** [0.021]	1.02 [0.074]	0.77*** [0.034]	9.01*** [0.181]
22-24 years	0.64*** [0.027]	0.92 [0.090]	0.73*** [0.041]	10.59*** [0.245]
25+ years	0.61*** [0.032]	1.09 [0.138]	0.68*** [0.047]	12.82*** [1.982]
Multi-unit status	5.06*** [0.082]	0.37*** [0.014]	3.33*** [0.076]	2.94*** [0.160]
Firm union status	3.59*** [0.037]	6.51*** [0.216]	6.10*** [0.081]	5.18*** [0.063]
Right-to-work status	1.13* [0.074]	0.78 [0.387]	0.93 [0.058]	1.12** [0.057]
Eligible employees %	—	0.77*** [0.006]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.2. Odds ratios based on logit model estimates – Manufacturing
(Sample restricted to 1990-2007 period)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	5.48*** [0.420]	0.66*** [0.107]	6.34*** [0.719]	2.42*** [0.135]
20-49 employees	13.76*** [0.964]	0.48*** [0.073]	15.16*** [1.593]	4.97*** [0.279]
50-99 employees	26.33*** [2.006]	0.31*** [0.049]	22.31*** [2.650]	7.44*** [0.478]
100-249 employees	33.44*** [2.675]	0.23*** [0.038]	23.39*** [2.956]	8.29*** [0.589]
250-499 employees	36.45*** [3.442]	0.22*** [0.045]	26.68*** [4.117]	7.81*** [0.711]
500+ employees	24.56*** [2.926]	0.17*** [0.048]	17.74*** [3.504]	6.34*** [0.790]
4-6 years	0.84*** [0.041]	0.97 [0.102]	0.82*** [0.065]	2.66*** [0.105]
7-9 years	0.79*** [0.040]	0.88 [0.033]	0.76*** [0.064]	4.08*** [0.197]
10-12 years	0.69*** [0.038]	1.04 [0.125]	0.73*** [0.063]	5.71*** [0.298]
13-15 years	0.67*** [0.040]	0.76** [0.104]	0.59*** [0.060]	7.72*** [0.434]
16-18 years	0.65*** [0.046]	0.99 [0.153]	0.65*** [0.074]	10.01*** [0.633]
19-21 years	0.62*** [0.053]	1.09 [0.203]	0.67*** [0.091]	12.91*** [0.921]
22-24 years	0.54*** [0.063]	0.55** [0.147]	0.39*** [0.084]	16.09*** [1.295]
25+ years	0.58*** [0.088]	1.26 [0.427]	0.62*** [0.141]	20.82*** [1.982]
Multi-unit status	2.12*** [0.098]	0.72*** [0.059]	1.85*** [0.137]	2.81*** [0.160]
Firm union status	2.02*** [0.086]	2.20*** [0.281]	2.99*** [0.194]	2.75*** [0.118]
Right-to-work status	0.68** [0.120]	0.92 [0.373]	0.68 [0.192]	0.97 [0.088]
Eligible employees %	—	0.72*** [0.027]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and the 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.3. Odds ratios based on logit model estimates – Manufacturing

(Sample restricted to the 1982-1977 period) (Sample restricted to the 2000-2007 period)

Event:	Certification:		Successful	Union	Certification:		Successful	Union
	Election	Win	Organizing	Status	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 emp.	5.53*** [0.544]	0.66* [0.172]	5.96*** [0.852]	5.10*** [0.553]	4.69*** [0.638]	0.64 [0.200]	5.11*** [0.973]	2.28*** [0.158]
20-49 emp.	11.67*** [1.121]	0.56** [0.137]	11.41*** [1.638]	10.14*** [1.113]	10.11*** [1.279]	0.38*** [0.109]	9.60*** [1.736]	4.11*** [0.275]
50-99 emp.	17.05*** [1.843]	0.54** [0.143]	16.50*** [2.655]	14.65*** [1.846]	18.03*** [2.471]	0.27*** [0.081]	13.41*** [2.731]	6.17*** [0.464]
100-249 emp.	21.09*** [2.426]	0.38*** [0.107]	16.85*** [2.979]	17.64*** [2.377]	22.87*** [3.284]	0.18*** [0.058]	12.99*** [2.821]	6.62*** [0.539]
250-499 emp.	18.58*** [2.94]	0.21*** [0.083]	10.26*** [2.907]	15.30*** [3.129]	23.32*** [3.929]	0.23*** [0.086]	17.68*** [4.525]	6.37*** [0.657]
500+ emp.	14.58*** [3.13]	0.24*** [0.128]	9.18*** [3.331]	10.35*** [2.858]	20.85*** [4.110]	0.12*** [0.057]	12.55*** [4.030]	5.19*** [0.692]
4-6 years	0.80** [0.041]	1.07 [0.102]	0.76* [0.126]	2.62*** [0.186]	0.83* [0.086]	0.87 [0.204]	0.74* [0.119]	2.08*** [0.165]
7-9 years	—	—	—	—	0.74*** [0.082]	0.72 [0.179]	0.65*** [0.112]	2.94*** [0.299]
10-12 years	—	—	—	—	0.71*** [0.080]	1.04 [0.261]	0.74* [0.125]	4.23*** [0.461]
13-15 years	—	—	—	—	0.63*** [0.074]	0.66 [0.175]	0.59*** [0.096]	5.75*** [0.618]
16-18 years	—	—	—	—	0.64*** [0.076]	1.03 [0.276]	0.61*** [0.074]	7.80*** [0.833]
19-21 years	—	—	—	—	0.62*** [0.075]	1.17 [0.331]	0.55** [0.011]	9.65*** [1.035]
22-24 years	—	—	—	—	0.55*** [0.076]	0.54** [0.172]	0.38*** [0.091]	12.26*** [1.325]
25+ years	—	—	—	—	0.58*** [0.097]	1.13 [0.444]	0.39*** [0.081]	15.80*** [1.871]
Multi-unit	2.68*** [0.199]	0.72*** [0.059]	2.55*** [0.286]	2.66*** [0.288]	2.72*** [0.245]	0.66*** [0.110]	2.44*** [0.341]	3.09*** [0.205]
Firm union	1.51*** [0.117]	2.20*** [0.281]	2.19*** [0.237]	2.69*** [0.236]	2.28*** [0.154]	1.73*** [0.294]	3.12*** [0.330]	2.75*** [0.134]
Right-to-work	1.07 [0.268]	0.92 [0.373]	0.97 [0.422]	0.96 [0.446]	0.26** [0.182]	1.41 [2.026]	0.32 [0.396]	0.92 [0.108]
Eligible emp.	—	0.83*** [0.054]	—	—	—	0.71*** [0.042]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and the 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.4. Odds ratios based on logit model estimates – All Sectors
(Sample restricted to establishments with 5+ employees)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	1.75*** [0.028]	0.84*** [0.029]	1.81*** [0.042]	1.61*** [0.029]
20-49 employees	3.20*** [0.052]	0.67*** [0.024]	3.18*** [0.074]	2.84*** [0.062]
50-99 employees	5.04*** [0.097]	0.56*** [0.024]	4.79*** [0.133]	4.64*** [0.119]
100-249 employees	6.83*** [0.142]	0.49*** [0.023]	6.21*** [0.190]	6.45*** [0.189]
250-499 employees	7.87*** [0.239]	0.46*** [0.032]	7.29*** [0.324]	7.33*** [0.301]
500+ employees	8.71*** [0.325]	0.31*** [0.027]	7.98*** [0.435]	9.67*** [0.535]
4-6 years	0.82*** [0.012]	1.06** [0.032]	0.87*** [0.017]	2.22*** [0.022]
7-9 years	0.76*** [0.013]	1.00 [0.036]	0.79*** [0.018]	3.24*** [0.044]
10-12 years	0.69*** [0.014]	1.07 [0.046]	0.75*** [0.021]	4.12*** [0.066]
13-15 years	0.66*** [0.016]	0.96 [0.049]	0.68*** [0.023]	5.16*** [0.094]
16-18 years	0.64*** [0.018]	1.03 [0.062]	0.69*** [0.027]	6.33*** [0.131]
19-21 years	0.63*** [0.022]	0.98 [0.071]	0.66*** [0.031]	7.68*** [0.182]
22-24 years	0.59*** [0.026]	0.91 [0.088]	0.62*** [0.038]	9.04*** [0.249]
25+ years	0.58*** [0.032]	1.08 [0.134]	0.63*** [0.046]	11.15*** [0.380]
Multi-unit status	1.66*** [0.023]	0.51*** [0.014]	0.98 [0.022]	1.40*** [0.033]
Firm union status	4.17*** [0.053]	4.36*** [0.124]	8.22*** [0.168]	5.98*** [0.122]
Right-to-work status	1.03 [0.053]	0.99 [0.387]	1.06 [0.080]	0.94 [0.054]
Eligible employees %	—	0.76*** [0.006]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 5-9 employees and 0-3 years of age.

TABLE D.5. Odds ratios based on logit model estimates – Manufacturing
(Sample restricted to establishments with 5+ employees)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	3.51 ^{***} [0.233]	0.65 ^{***} [0.107]	3.05 ^{***} [0.278]	2.22 ^{***} [0.136]
20-49 employees	9.16 ^{***} [0.576]	0.49 ^{***} [0.073]	7.27 ^{***} [0.627]	4.71 ^{***} [0.326]
50-99 employees	16.56 ^{***} [1.097]	0.37 ^{***} [0.049]	11.17 ^{***} [1.045]	7.55 ^{***} [0.575]
100-249 employees	21.65 ^{***} [1.484]	0.28 ^{***} [0.038]	12.36 ^{***} [1.225]	8.82 ^{***} [0.726]
250-499 employees	23.16 ^{***} [1.837]	0.25 ^{***} [0.045]	12.87 ^{***} [1.558]	8.38 ^{***} [0.837]
500+ employees	16.55 ^{***} [1.594]	0.17 ^{***} [0.048]	8.15 ^{***} [1.284]	6.83 ^{***} [0.891]
4-6 years	0.83 ^{***} [0.028]	0.87 [*] [0.102]	0.76 ^{***} [0.041]	2.36 ^{***} [0.062]
7-9 years	0.75 ^{***} [0.030]	0.87 [0.033]	0.71 ^{***} [0.046]	3.57 ^{***} [0.135]
10-12 years	0.66 ^{***} [0.032]	1.03 [0.125]	0.67 ^{***} [0.051]	4.89 ^{***} [0.224]
13-15 years	0.64 ^{***} [0.037]	0.73 ^{**} [0.104]	0.54 ^{***} [0.053]	6.40 ^{***} [0.340]
16-18 years	0.62 ^{***} [0.043]	0.97 [0.153]	0.60 ^{***} [0.067]	8.21 ^{***} [0.502]
19-21 years	0.59 ^{***} [0.051]	0.99 [0.203]	0.58 ^{***} [0.080]	10.54 ^{***} [0.742]
22-24 years	0.51 ^{***} [0.061]	0.57 ^{**} [0.147]	0.37 ^{***} [0.081]	12.92 ^{***} [1.060]
25+ years	0.56 ^{***} [0.086]	1.30 [0.427]	0.61 ^{**} [0.136]	19.96 ^{***} [1.701]
Multi-unit status	1.61 ^{***} [0.049]	0.77 ^{***} [0.059]	1.40 ^{***} [0.070]	2.17 ^{***} [0.112]
Firm union status	2.11 ^{***} [0.080]	2.21 ^{***} [0.281]	3.36 ^{***} [0.189]	2.65 ^{***} [0.132]
Right-to-work status	0.91 [0.125]	0.75 [0.373]	0.77 [0.172]	0.97 [0.114]
Eligible employees %	—	0.77 ^{***} [0.027]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 5-9 employees and 0-3 years of age.

TABLE D.6. Odds ratios based on logit model estimates
(Labor productivity minus average wage as explanatory variable)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
All Sectors				
11-25 percentile	0.58 ^{***} [0.104]	1.13 [0.213]	0.49 ^{***} [0.147]	0.77* [0.106]
26-50 percentile	0.78 ^{**} [0.202]	0.31 ^{***} [0.093]	0.45 ^{**} [0.160]	0.68 ^{***} [0.083]
51-75 percentile	1.15 ^{**} [0.261]	0.84 [0.246]	0.99 [0.349]	1.24 ^{***} [0.126]
76-90 percentile	1.29* [0.366]	0.71 [0.262]	0.82 [0.290]	1.38 ^{***} [0.171]
91-100 percentile	0.86 [0.263]	1.05 [0.345]	0.83 [0.424]	0.99 [0.223]
Manufacturing				
11-25 percentile	1.25 [0.186]	0.77 [0.221]	1.21 [0.251]	1.31 ^{***} [0.109]
26-50 percentile	1.86 ^{***} [0.239]	0.59* [0.199]	1.54 ^{***} [0.281]	1.66 ^{***} [0.123]
51-75 percentile	1.94 ^{***} [0.249]	0.51 ^{**} [0.176]	1.48 ^{**} [0.271]	1.95 ^{***} [0.142]
76-90 percentile	2.33 ^{***} [0.308]	0.37 ^{***} [0.118]	1.41* [0.278]	2.37 ^{***} [0.176]
91-100 percentile	2.63 ^{***} [0.350]	0.43 ^{***} [0.135]	1.79 ^{***} [0.349]	2.88 ^{***} [0.225]

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicate significance at the 10%, 5%, and 1% level, respectively. Models include all other explanatory variables in Tables 1 and 2. The 1-10 percentile category is omitted.

TABLE D.7. Odds ratios based on logit model estimates – All Sectors
(The logarithm of average wage as explanatory variable)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
log(average wage)	1.22*** [0.006]	0.99 [0.029]	1.19*** [0.047]	1.16*** [0.009]
10-19 employees	2.21*** [0.034]	0.73*** [0.023]	2.53*** [0.055]	2.12*** [0.040]
20-49 employees	3.61*** [0.057]	0.59*** [0.019]	3.94*** [0.091]	3.44*** [0.078]
50-99 employees	5.03*** [0.094]	0.48*** [0.019]	5.14*** [0.143]	4.95*** [0.133]
100-249 employees	6.11*** [0.124]	0.39*** [0.018]	5.86*** [0.179]	6.09*** [0.187]
250-499 employees	6.50*** [0.195]	0.37*** [0.026]	6.29*** [0.279]	6.46*** [0.274]
500+ employees	6.81*** [0.254]	0.22*** [0.020]	6.35*** [0.349]	7.83*** [0.446]
4-6 years	0.84*** [0.011]	1.04 [0.029]	0.92*** [0.017]	2.45*** [0.023]
7-9 years	0.77*** [0.012]	0.99 [0.033]	0.85*** [0.019]	3.64*** [0.047]
10-12 years	0.71*** [0.013]	1.10** [0.045]	0.82*** [0.022]	4.67*** [0.070]
13-15 years	0.68*** [0.015]	0.98 [0.047]	0.77*** [0.024]	5.86*** [0.100]
16-18 years	0.64*** [0.017]	1.03 [0.061]	0.75*** [0.028]	7.13*** [0.139]
19-21 years	0.62*** [0.020]	1.00 [0.071]	0.72*** [0.033]	8.55*** [0.190]
22-24 years	0.58*** [0.024]	0.93 [0.088]	0.68*** [0.039]	10.02*** [0.260]
25+ years	0.58*** [0.031]	1.11 [0.138]	0.67** [0.047]	12.49*** [0.403]
Multi-unit status	3.66*** [0.058]	0.44*** [0.013]	1.94*** [0.047]	2.33*** [0.061]
Firm union status	5.11*** [0.061]	6.01*** [0.161]	11.87*** [0.238]	7.99*** [0.164]
Right-to-work status	1.02 [0.046]	1.15 [0.115]	1.26*** [0.084]	1.01 [0.054]
Eligible employees %	—	0.75*** [0.005]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.8. Odds ratios based on logit model estimates – Manufacturing
(The logarithm of average wage as explanatory variable)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
log(average wage)	1.23*** [0.054]	0.73*** [0.036]	1.13*** [0.032]	1.01 [0.025]
10-19 employees	6.29*** [0.348]	0.58*** [0.070]	6.12*** [0.481]	3.32*** [0.043]
20-49 employees	15.93*** [0.825]	0.44*** [0.049]	14.21*** [1.061]	6.87*** [0.078]
50-99 employees	28.18*** [1.592]	0.33*** [0.039]	21.39*** [1.804]	10.76*** [0.141]
100-249 employees	36.11*** [2.154]	0.25*** [0.030]	23.31*** [2.126]	12.37*** [0.212]
250-499 employees	37.89*** [2.731]	0.22*** [0.034]	23.73*** [2.736]	11.60*** [0.333]
500+ employees	26.19*** [2.385]	0.15*** [0.033]	14.65*** [2.254]	9.34*** [0.590]
4-6 years	0.84*** [0.028]	0.85 [0.061]	0.78*** [0.041]	2.53*** [0.027]
7-9 years	0.76*** [0.030]	0.88 [0.075]	0.72*** [0.046]	3.87*** [0.049]
10-12 years	0.67*** [0.032]	1.04 [0.106]	0.70*** [0.053]	5.36*** [0.069]
13-15 years	0.64*** [0.037]	0.74 [0.096]	0.55*** [0.053]	7.07*** [0.095]
16-18 years	0.62*** [0.045]	0.98 [0.144]	0.60*** [0.068]	9.11*** [0.131]
19-21 years	0.59*** [0.050]	1.05 [0.189]	0.61*** [0.082]	11.72*** [0.181]
22-24 years	0.52*** [0.060]	0.56 [0.146]	0.37*** [0.080]	14.47*** [0.245]
25+ years	0.53*** [0.082]	1.31 [0.431]	0.57** [0.129]	18.69*** [1.982]
Multi-unit status	1.68*** [0.054]	0.78*** [0.045]	1.49*** [0.078]	2.36*** [0.127]
Firm union status	2.13*** [0.079]	2.39*** [0.185]	3.48*** [0.191]	2.77*** [0.136]
Right-to-work status	0.88 [0.119]	0.87 [0.252]	0.82 [0.179]	0.97 [0.111]
Eligible employees %	—	0.77*** [0.019]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.9. Odds ratios based on logit model estimates – All Sectors
 (Actual minus predicted logarithm of average wage as explanatory variable)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
Actual – predicted log(average wage)	1.30*** [0.007]	0.98 [0.014]	1.25*** [0.011]	1.21*** [0.010]
10-19 employees	2.19*** [0.033]	0.74*** [0.023]	2.50*** [0.054]	2.11*** [0.039]
20-49 employees	3.59*** [0.056]	0.59*** [0.020]	3.90*** [0.088]	3.41*** [0.077]
50-99 employees	5.04*** [0.093]	0.48*** [0.019]	5.12*** [0.141]	4.95*** [0.133]
100-249 employees	6.16*** [0.123]	0.40*** [0.018]	5.88*** [0.178]	6.13*** [0.187]
250-499 employees	6.58*** [0.196]	0.37*** [0.026]	6.33*** [0.280]	6.50*** [0.275]
500+ employees	7.05*** [0.262]	0.22*** [0.020]	6.52*** [0.357]	8.02*** [0.457]
4-6 years	0.85*** [0.011]	1.03 [0.029]	0.93*** [0.017]	2.45*** [0.023]
7-9 years	0.78*** [0.012]	0.99 [0.034]	0.87*** [0.019]	3.67*** [0.046]
10-12 years	0.73*** [0.013]	1.09** [0.045]	0.85*** [0.022]	4.74*** [0.070]
13-15 years	0.71*** [0.015]	0.98 [0.047]	0.80*** [0.025]	5.98*** [0.100]
16-18 years	0.67*** [0.018]	1.03 [0.061]	0.79*** [0.029]	7.31*** [0.139]
19-21 years	0.66*** [0.021]	1.00 [0.071]	0.76*** [0.034]	8.81*** [0.193]
22-24 years	0.62*** [0.026]	0.93 [0.088]	0.72*** [0.042]	10.39*** [0.266]
25+ years	0.63*** [0.033]	1.11 [0.136]	0.73** [0.051]	13.06*** [0.416]
Multi-unit status	3.89*** [0.059]	0.44*** [0.013]	2.03*** [0.049]	2.45*** [0.063]
Firm union status	5.13*** [0.060]	6.00*** [0.161]	11.92*** [0.238]	8.01*** [0.164]
Right-to-work status	1.02 [0.046]	1.15 [0.116]	1.26*** [0.083]	1.01 [0.054]
Eligible employees %	—	0.75*** [0.005]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.10. Odds ratios based on logit model estimates – Manufacturing
(Actual minus predicted logarithm of average wage as explanatory variable)

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
Actual – predicted log(average wage)	1.18 *** [0.023]	0.76 *** [0.037]	1.09 *** [0.033]	0.99 [0.026]
10-19 employees	6.33 *** [0.351]	0.59 *** [0.071]	6.17 *** [0.485]	3.33 *** [0.197]
20-49 employees	16.03 *** [0.833]	0.45 *** [0.050]	14.32 *** [1.074]	6.91 *** [0.457]
50-99 employees	28.38 *** [1.607]	0.34 *** [0.040]	21.56 *** [1.825]	10.81 *** [0.807]
100-249 employees	36.46 *** [2.178]	0.25 *** [0.031]	23.53 *** [2.151]	12.43 *** [1.016]
250-499 employees	38.56 *** [2.778]	0.23 *** [0.035]	24.06 *** [2.772]	11.66 *** [1.167]
500+ employees	27.53 *** [2.491]	0.15 *** [0.032]	15.14 *** [2.315]	9.42 *** [1.234]
4-6 years	0.85 *** [0.028]	0.85 ** [0.061]	0.78 *** [0.041]	2.53 *** [0.066]
7-9 years	0.77 *** [0.031]	0.86 * [0.073]	0.73 *** [0.047]	3.88 *** [0.145]
10-12 years	0.69 *** [0.033]	1.01 [0.102]	0.71 *** [0.053]	5.37 *** [0.242]
13-15 years	0.66 *** [0.038]	0.71 *** [0.092]	0.56 *** [0.054]	7.09 *** [0.372]
16-18 years	0.65 *** [0.044]	0.91 [0.135]	0.62 *** [0.069]	9.13 *** [0.551]
19-21 years	0.62 *** [0.053]	0.99 [0.177]	0.63 *** [0.084]	11.75 *** [0.818]
22-24 years	0.55 *** [0.064]	0.51 *** [0.134]	0.39 *** [0.082]	14.50 *** [1.177]
25+ years	0.57 *** [0.088]	1.19 [0.389]	0.59 ** [0.133]	18.73 *** [1.856]
Multi-unit status	1.84 *** [0.060]	0.68 *** [0.041]	1.57 *** [0.082]	2.37 *** [0.128]
Firm union status	2.14 *** [0.080]	2.36 *** [0.183]	3.50 *** [0.192]	2.77 *** [0.137]
Right-to-work status	0.88 [0.119]	0.88 [0.254]	0.82 [0.178]	0.98 [0.112]
Eligible employees %	—	0.77 *** [0.019]	—	—

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include 2-digit SIC industry, state, and year fixed effects. The following categories are omitted: 1-9 employees and 0-3 years of age.

TABLE D.11. Odds ratios based on logit model estimates – All Sectors

[Size (employment) and productivity (value of shipments per employee) effects together]

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	6.25 ^{***} [2.210]	1.51 [0.886]	7.15 ^{***} [3.341]	4.79 ^{***} [1.060]
20-49 employees	10.89 ^{***} [3.471]	0.75 [0.424]	10.32 ^{***} [3.436]	9.54 ^{***} [1.857]
50-99 employees	16.46 ^{***} [4.725]	0.73 [0.551]	18.15 ^{***} [7.011]	13.71 ^{***} [2.605]
100-249 employees	18.45 ^{***} [5.793]	0.67 [0.523]	18.73 ^{***} [7.845]	14.36 ^{***} [2.723]
250-499 employees	22.19 ^{***} [8.946]	0.45 [0.647]	25.73 ^{***} [14.463]	23.60 ^{***} [5.437]
500+ employees	14.71 ^{***} [4.296]	0.42 [0.300]	20.53 ^{***} [4.677]	27.12 ^{***} [7.450]
11-25 percentile	0.81 [0.242]	0.58 [*] [0.191]	0.52 ^{**} [0.186]	0.83 [0.218]
26-50 percentile	1.49 [0.404]	0.54 [0.211]	1.08 [0.444]	1.05 [0.259]
51-75 percentile	1.46 [0.459]	0.43 [0.317]	1.36 [0.564]	1.30 [0.319]
76-90 percentile	1.09 [0.297]	0.35 [0.511]	1.75 [*] [0.421]	1.24 [0.318]
91-100 percentile	1.43 [0.578]	0.51 [0.352]	1.77 [*] [0.417]	1.04 [0.255]

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include all other explanatory variables in Table 1. The following categories are omitted: 1-9 employees, 0-3 years of age, 1-10 percentile.

TABLE D.12. Odds ratios based on logit model estimates – Manufacturing
 [Size (employment) and productivity (value of shipments per employee) effects together]

Event:	Certification:		Successful	Union
	Election	Win	Organizing	Status
Probability:	$T^o(x_a, a)$	$W^o(x_a, a)$	$O^o(x_a, a)$	$U^o(x_a, a)$
10-19 employees	7.61 ^{***} [1.086]	0.74 [0.107]	6.94 ^{***} [0.278]	3.75 ^{***} [1.176]
20-49 employees	19.49 ^{***} [2.605]	0.77 [0.073]	16.83 ^{***} [0.627]	5.46 ^{***} [1.741]
50-99 employees	34.38 ^{***} [4.921]	0.80 [0.049]	26.03 ^{***} [1.045]	9.31 ^{***} [3.165]
100-249 employees	41.66 ^{***} [6.343]	0.84 [0.038]	28.25 ^{***} [1.225]	9.50 ^{***} [3.284]
250-499 employees	44.22 ^{***} [7.845]	0.76 [0.045]	25.42 ^{***} [1.558]	7.74 ^{***} [2.810]
500+ employees	37.85 ^{***} [8.268]	0.80 [0.048]	21.51 ^{***} [1.284]	6.58 ^{***} [2.496]
11-25 percentile	1.38 ^{**} [0.192]	0.56 ^{**} [0.212]	1.07 [0.220]	1.38 [*] [0.244]
26-50 percentile	1.98 ^{***} [0.237]	0.63 [*] [0.181]	1.58 ^{***} [0.238]	1.86 ^{***} [0.268]
51-75 percentile	1.90 ^{***} [0.229]	0.57 ^{**} [0.163]	1.42 ^{**} [0.290]	2.81 ^{***} [0.497]
76-90 percentile	2.30 ^{***} [0.290]	0.36 ^{***} [0.183]	1.32 [0.322]	2.78 ^{***} [0.399]
91-100 percentile	2.70 ^{***} [0.347]	0.41 ^{***} [0.231]	1.67 ^{***} [0.417]	2.95 ^{***} [0.331]

Notes: Robust standard errors, clustered by establishment, are in brackets. (*), (**), (***) indicates significance at the 10%, 5%, and 1% level, respectively. Models include all other explanatory variables in Table 1. The following categories are omitted: 1-9 employees, 0-3 years of age, 1-10 percentile.