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### SAMPLE DESIGN

**Introduction.** The estimates for each of the 11 metropolitan areas in this report series (H-170-85) are based on data collected from the 1985 American Housing Survey Metropolitan Sample (AHS-MS) and the 1985 American Housing Survey National Sample, which were conducted by the Bureau of the Census acting as collection agent for the Department of Housing and Urban Development.

The sample areas covered for metropolitan areas that remained in the AHS sample after survey year 1983 are consistent with the 1983 Office of Management and Budget (OMB) definitions of a metropolitan statistical area (MSA), consolidated metropolitan statistical area (CMSA), or primary metropolitan statistical area (PMSA). In some instances, a given metropolitan area is a combination of primary metropolitan statistical areas and will be referred to as PMSA's. In addition to adding new areas to some metropolitan samples in order to comply with the 1983 definitional changes, some new metropolitan areas have been added. Thus, each of the 1985 metropolitan areas will fall into one of three categories—

- a. Areas of the same geographic area as defined for surveys prior to 1984 (i.e., areas in which the 1970 OMB definition of a standard metropolitan statistical area is the same as the 1983 MSA, PMSA, or CMSA definition, 1970-based area)—Dallas, TX, PMSA; Los Angeles-Long Beach, CA, PMSA; Philadelphia, PA-NJ, PMSA; Phoenix, AZ, MSA; and San Francisco-Oakland, CA, area PMSA's.
- b. Areas consisting of new area in addition to the 1970-based area—Boston, MA-NH, CMSA; Detroit, MI, PMSA; Ft. Worth Arlington, TX, PMSA; Minneapolis-St. Paul, MN-WI, MSA; and Washington, DC-MD-VA, MSA.
- c. Areas that are in sample for the first time—Tampa-St. Petersburg, FL, MSA.

The metropolitan areas selected for the 1985 AHS-MS are interviewed on a rotating basis once every four years. Each metropolitan area had an expected sample size of 8,500 or 4,250 housing units uniformly distributed throughout nine panels (panels 4-12). The areas having an expected sample size of 8,500 housing units include the Detroit, MI, PMSA; Los Angeles-Long Beach, CA, PMSA; Philadelphia, PA-NJ, PMSA; San Francisco-Oakland, CA, area PMSA's; and Washington, DC-MD-VA, MSA. The remaining six metropolitan areas have an expected sample size of 4,250

housing units. Due to budget constraints, panel 12 was dropped from sample in all metropolitan areas; and with the exception of the Boston, MA-NH, CMSA; Minneapolis-St. Paul, MN-WI, MSA; Phoenix, AZ, MSA; and Tampa-St. Petersburg, FL, MSA's, panel 11 was dropped as well. In metropolitan areas where only panel 12 was dropped from sample, interviewing was scheduled for April 1985 through November 1985; in the other seven metropolitan areas, interviewing was scheduled for April 1985 through October 1985. Hence, the expected sample sizes were lower than the original goals of 4,250 and 8,500 sample units. In this metropolitan area, 3,949 AHS-MS housing units were eligible for interview. Of these sample housing units, 183 interviews were not obtained, because for occupied sample units, the occupants refused to be interviewed, were not at home after repeated visits, or were unavailable for some other reason; or, for vacant units, no informed respondent could be found after repeated visits. In addition to the AHS-MS housing units eligible for interview, 75 AHS-MS units were visited but were not eligible for interview because they were condemned, unfit, demolished, converted to group quarters use, etc.

The AHS-National sample is interviewed biennially on odd-numbered years. It was conducted from August 1985 through December 1985. The sample covers 878 counties and independent cities with coverage in each of the 50 States and the District of Columbia. In order to increase the reliability of the AHS-MS sample estimates, information from AHS-National sample units was used in the estimation process. For each metropolitan area, interviewed AHS-National units that were located within the 1985 AHS-MS definition of the metropolitan area were used in the estimation procedure. In this metropolitan area, 728 AHS-National units were used.

**Designation of AHS-MS Sample Housing Units for the 1985 Survey.** The sample housing units designated to be interviewed in the 1985 survey consisted of the following categories which are described in detail in the following sections:

Housing units which were in the 1970-based area include the following:

- a. All sample housing units that were interviewed in the previous survey and remained in sample after the 1985 reduction. This sample includes housing units that were selected as part of the 1976-1981 Coverage Improvement Program. These Coverage Improvement cases represented most of the housing units which, until these procedures were implemented, did not have a chance of selection.
- b. All sample housing units that were type A noninterviews (i.e., units eligible to be interviewed) or type B noninterviews (i.e., units not eligible for interview at the time of the survey but which could become eligible in

the future) in the previous survey and remained in sample after the 1985 reduction. (For a list of reasons for type A and type B noninterviews, see the facsimile of the 1985 AHS questionnaire, page App-24.)

- c. All sample housing units selected from a listing of new residential construction building permits issued since the previous survey that remained in sample after the 1985 reduction. This sample represented the housing units built in permit-issuing areas since the previous survey.
- d. All sample housing units that were added since the previous survey in sample segments from the nonpermit universe that remained in sample after the 1985 reduction. This sample represented additions to the housing inventory since the previous survey in nonpermit-issuing areas
- e. In the 1970-based areas of the Boston, MA-NH, CMSA; Detroit, MI, PMSA; Ft. Worth-Arlington, TX, PMSA; Minneapolis-St. Paul, MN-WI, MSA; and Washington, DC-MD-VA, MSA, all sample housing units selected from the 1980 Census of Population and Housing.
- f. All sample housing units reinstated in sample in 1985. This sample represents units which had previously been dropped from sample due to previous sample reductions.

Housing units within new areas added to the metropolitan area in 1980 and for metropolitan areas that are in sample for the first time (1980-based area):

- a. All housing units selected from the 1980 Census of Population and Housing.
- b. All housing units that were selected from a list of new residential construction building permits. This sample represented the housing units built in permit-issuing areas since the 1980 census.
- c. All sample housing units that were selected in sample segments added from the nonpermit universe. This sample represents units enumerated in the 1980 census as well as additions to the housing inventory in nonpermit-issuing areas since the 1980 census.

The following table shows the percent of the AHS-MS old construction sample that is in 1970-based area (i.e., frames a, b, d, e, and f) and 1980-based area (i.e., frames a and c) for each metropolitan area:

Metropolitan Area	Percent 1970-based area	Percent 1980-based area
Boston, MA-NH, CMSA .....	70.1	29.9
Dallas, TX, PMSA .....	100.0	0.0
Detroit, MI, PMSA .....	91.7	8.3
Ft. Worth-Arlington, TX, PMSA .....	96.2	3.8
Los Angeles-Long Beach, CA, PMSA .....	100.0	0.0
Minneapolis-St. Paul, MN-WI, MSA .....	91.6	8.4
Philadelphia, PA-NJ, PMSA .....	100.0	0.0
Phoenix, AZ, MSA .....	100.0	0.0
San Francisco-Oakland, CA, area PMSA's .....	100.0	0.0
Tampa-St. Petersburg, FL, MSA .....	0.0	100.0
Washington, DC-MD-VA, MSA .....	93.3	6.7

**1985 AHS-MS Original Sample Selection for the 1970-Based Area Sample of the Metropolitan Areas.** The 1985 AHS-MS original sample for the 1970-based area of the metropolitan areas which, in 1970, were 100-percent permit-issuing was selected from two frames: housing units enumerated in the 1970 Census of Population and Housing in areas under the jurisdiction of permit-issuing areas (the 1970-based permit-issuing universe); and housing units constructed in permit issuing areas since the 1970 census (the 1970-based new construction universe). In addition, the sample for those metropolitan areas which were not 100-percent permit-issuing in 1970 included a sample selected from a third frame—those housing units located in areas not under the jurisdiction of permit-issuing offices (the 1970-based nonpermit universe). In 1970, the Boston, MA-NH, CMSA; Los Angeles-Long Beach, CA, PMSA; Phoenix, AZ, MSA; San Francisco-Oakland, CA, area PMSA's; and Washington, DC-MD-VA, MSA were the only metropolitan areas that were 100-percent permit-issuing.

Sampling operations, described in the following paragraphs, were performed separately within the central city and balance (outside the central city) using the 1970 OMB definitions of the central city of each metropolitan area for each of the sample frames. The overall sampling rate used to select the sample for each metropolitan area was determined by the size of the sample. Each metropolitan area had an overall sampling rate about the same for both the central city and the balance, since the sample was distributed proportionately between the two according to the corresponding distribution of total housing units.

The major portion of the sample in each of the metropolitan areas was selected from a file which represented the 20-percent sample of housing units enumerated in permit-issuing areas of the metropolitan areas during the 1970 Census of Population and Housing. This file contained records for occupied housing units, vacant housing units, and housing units in certain special places or group quarters. Sampling operations were done separately for the special place and group quarters records, and for the occupied and vacant housing unit records. Before the sample was selected from the occupied and vacant housing unit records, the records were stratified by race of the

head of household (non-Black/Black), and the vacant records were stratified into four categories pertaining to the value or rent associated with the vacant housing units. The occupied housing unit records were further stratified so that each unit was assigned to one of 50 strata according to its tenure (owner/renter), family size, and family income category as illustrated by the following table:

Family income	Tenure													
	Owner family size					Renter family size								
	1	2	3	4	5+	1	2	3	4	5+				
Under \$3,000 .....														
\$3,000 to \$5,999 .....														
6,000 to \$9,999 .....														
\$10,000 to \$14,999 .....														
\$15,000 and over .....														

Thus, for the metropolitan areas, the occupied housing unit records from the permit-issuing universe were assigned to one of 100 strata for either the central city or for the balance, and the vacant housing unit records were assigned to one of the 4 vacant strata for either the central city or for the balance of the metropolitan areas. A sample selection procedure was then instituted that would produce one-half of the desired sample. However, whenever a record was selected to be in sample, the housing unit record adjacent to it on the file was also selected to be in sample, thereby insuring the necessary designated sample size.

Before the sample was selected from the group quarters and special place records, the records were stratified by census tract and census enumeration district (ED) within the central city and within the balance of the metropolitan areas. A sample of special place records was then selected by a procedure that produced one-quarter of the desired sample size. However, at the time of the survey, the housing units at each of the special places were listed and subsampled at a rate which produced an expected four sample units, thereby insuring the necessary designated sample size.

The second frame from which the metropolitan area sample was selected was a list of new construction building permits issued since 1970 (i.e., the new construction universe). The sample selection from the list of new construction building permits was an independent operation within the metropolitan area. Under clerically selected procedures, the list of permits was chronologically stratified by the date the permits were issued, and clusters of an expected four (usually adjacent) housing units were formed. These clusters were then sampled for inclusion at the overall sampling rate. In February 1984, the new construction sampling operation for the 1970-based and 1980 based areas were combined into one computerized system. The universe sampled in the computerized system will be referred to in the estimation section as the 1980-based

permit universe. Under these procedures, prior to sample selection, the list of permits was chronologically stratified by the date of issue, State, 1980 central city and balance, county or minor civil division, and permit office. Clusters of an expected four (usually adjacent) housing units were formed. These clusters were then sampled for inclusion at twice the overall sampling rate. The housing units within each of the clusters were then subsampled so that two of the four housing units originally selected were kept in sample.

For those metropolitan areas which were not 100-percent permit issuing, the remainder of the AHS-MS sample was selected from a frame consisting of areas not under the jurisdiction of permit-issuing offices (i.e., the nonpermit universe). The first step in the sampling operation for the nonpermit universe was the selection of a sample of census enumeration districts. Prior to this sample selection, the ED's were stratified by census tract within the central city and within the balance of the metropolitan area. The probability of selection of an ED was proportionate to the following:

$$\frac{\text{Number of housing units in 1970 census ED} + \text{Group quarters population in 1970 census ED}}{4} = 3$$

The sample ED's were then divided into segments (i.e., small land areas with well-defined boundaries having an expected size of four, or a multiple of four, housing units). At the time of the survey, those segments that did not have an expected size of four were further subdivided to produce an expected four sample housing units. The next step was the selection of one of these segments within each sample ED. All housing units in existence at the time of interview in these selected segments were eligible for sample. Thus, housing units enumerated in the 1970 census as well as housing units built since the 1970 census were included.

**Sample Selection for the AHS-MS Coverage Improvement Program.** The AHS-MS Coverage Improvement Program was undertaken to correct certain deficiencies in the AHS-Metropolitan Area sample from the 1970 permit-issuing universe and the 1970 new construction universe within the 1970-based area. The coverage deficiencies included the following units:

- New construction from building permits issued prior to January 1970, but completed after April 1, 1970.
- Mobile homes placed in parks either missed during the 1970 census or established since the 1970 census.
- Housing units missed in the 1970 census.
- Housing units converted to residential use that were nonresidential at the time of the 1970 census.

- Houses that have been moved onto their present site since the 1970 census.
- Mobile homes placed outside parks since the 1970 census or vacant at the time of the 1970 census.

For a detailed description of the coverage improvement sample selection process, see earlier reports in the H-170 series for the years 1976 through 1981.

**1985 AHS-MS Sample Reduction and Sample Reinstatement.** The 1985 AHS-MS sample reduction dropped units from sample, whereas the 1985 AHS-MS sample reinstatement added enumerated units which were previously dropped from sample. The universes involved were: the 1970-based permit-issuing universe, the 1970-based new construction universe, and the 1970-based nonpermit universe. These procedures involved dropping or adding individual housing units from the permit issuing universe, whole clusters from the new construction universe, and whole segments from the nonpermit universe.

The reduction/reinstatement was implemented to achieve two criteria:

- A sample size of 8,500 in the Detroit, MI, PMSA; Los Angeles-Long Beach, CA, PMSA; Philadelphia, PA-NJ, PMSA; San Francisco-Oakland, CA, area PMSA's; and Washington, DC-MD-VA, MSA; and a sample size of 4,250 in the other six metropolitan areas;
- A sample having an equal number of owners and renters.

In order to achieve these results, each unit was classified according to the original panel number (the original sample was divided into 12 panels, with one-twelfth of the sample being in each panel) and 1985 tenure (each housing unit was given a 1985 tenure based on the previous year's tenure status). In order to simplify field procedures, panels 1-3 (i.e., a random one-fourth of the original sample) were dropped from sample whenever possible. More sample reductions were implemented separately for each 1985 tenure group (using different selection rates) across the remaining panels.

**AHS-MS Sample Selection for the 1980-Based Area Sample of the Metropolitan Areas.** The sample for new areas added to the 1970-metropolitan areas, and metropolitan areas in sample for the first time which, in 1980, were 100-percent permit-issuing was selected from two frames—housing units enumerated in the 1980 Census of Population and Housing in areas under the jurisdiction of permit-issuing areas (the 1980-based permit-issuing universe) and housing units constructed in permit-issuing areas since the 1980 census (the 1980-based new construction universe). In addition, the sample for those metropolitan areas which were not 100-percent permit-issuing in 1980 included a sample from a third frame—

those housing units not under the jurisdiction of permit-issuing offices (the 1980-based nonpermit universe). In 1980, the Boston, MA-NH, CMSA; Ft. Worth-Arlington, TX, PMSA; Minneapolis-St. Paul, MN-WI, MSA; and Washington, DC-MD-VA, MSA were the only metropolitan areas that added new areas which were not 100-percent permit-issuing.

In order to satisfy confidentiality requirements in the Boston, MA-NH, CMSA; Ft. Worth-Arlington, TX, PMSA; and Washington, DC-MD-VA, MSA, it was necessary to supplement the existing sample within the 1970-based area for each metropolitan area. The additional sample housing units were selected separately for each metropolitan area from the 1980-based permit-issuing universe.

The major portion of the sample in each metropolitan area was selected from a file which represented all the housing units enumerated in permit-issuing areas of the metropolitan area during the 1980 Census of Population and Housing. This file contained records for occupied housing units, vacant housing units, and housing units in group quarters. Sampling operations were done separately for noninstitutionalized group quarters and for all other housing units in permit-issuing areas. In addition, in order that an equal number of owner and renter housing units were selected in each metropolitan area, a selection rate that differed by tenure group was used. Before the sample was selected, the housing units that were not classified as group quarters were stratified into 60 categories by tenure, contract rent, value, and number of rooms as illustrated by the following table:

Contract rent and value	Number of rooms		
	1-3	4-5	6+
<b>RENTER</b>			
Contract rent			
Less than \$100 .....			
\$100 to \$149 .....			
\$150 to \$199 .....			
\$200 to \$249 .....			
\$250 to \$299 .....			
\$300 to \$349 .....			
\$350 to \$399 .....			
\$400 or more .....			
Not available .....			
<b>OWNER</b>			
Value			
Less than \$20,000 .....			
\$20,000 to \$29,999 .....			
\$30,000 to \$34,999 .....			
\$35,000 to \$39,999 .....			
\$40,000 to \$49,999 .....			
\$50,000 to \$64,999 .....			
\$65,000 to \$79,999 .....			
\$80,000 to \$99,999 .....			
\$100,000 to \$149,999 .....			
\$150,000 or more .....			
Not available .....			

The group quarters housing units were grouped into two strata: (1) institutionalized group quarters; and (2) noninstitutionalized group quarters.

The following sample selection procedures were then implemented separately within the 1980 central city and balance of the metropolitan area. For the Boston, MA-NH, CMSA; Ft. Worth-Arlington, TX, PMSA; and Washington, DC-MD-VA, MSA, the sample selections were implemented separately by the 1970-based and 1980-based areas. First, all units were sorted by the 1980 central city and balance, stratum, State, district office, ED, and census serial number. The sample selection procedure was then implemented separately for noninstitutionalized group quarters and for institutionalized group quarters and nongroup quarters housing units. For the institutionalized group quarters and nongroup quarters housing units, the sample selection was done across the 61 strata. Individual housing units were selected for the nongroup quarters while each institutionalized group quarters had one chance of selection. Before the sample selection for the noninstitutionalized group quarters was implemented, the following measure of size was calculated for each record:

$$(1/4) \times \frac{\text{Total group quarters population}}{2.75}$$

The noninstitutionalized group quarters were then selected proportionate to the measure of size.

The second frame from which the metropolitan area sample was selected was a list of new construction building permits issued since 1980 (i.e., the new construction universe). The sample selection from the list of new construction building permits was an independent operation within each metropolitan area. This operation was described previously in the discussion on computerized building permit sampling in 1970-based areas.

The following table shows the percent of the new construction sample that was clerically selected since the previous survey (i.e., cluster size = 4) and computer selected (i.e., cluster size = 2) for each metropolitan area:

Metropolitan area	Percent clerically selected	Percent computer selected
Boston, MA-NH, CMSA .....	17.0	83.0
Dallas, TX, PMSA .....	66.7	33.3
Detroit, MI, PMSA .....	41.4	58.6
Ft. Worth-Arlington, TX, PMSA .....	62.1	37.9
Los Angeles-Long Beach, CA, PMSA .....	48.0	52.0
Minneapolis-St. Paul, MN-WI, MSA .....	53.4	46.6
Philadelphia, PA-NJ, PMSA .....	48.8	51.2
Phoenix, AZ, MSA .....	52.4	47.6
San Francisco-Oakland, CA, area PMSA's .....	44.3	55.7
Tampa-St. Petersburg, FL, MSA .....	0.0	100.0
Washington, DC-MD-VA, MSA .....	49.0	51.0

For those metropolitan areas which were not 100-percent permit issuing, the remainder of the AHS-MS sample was selected from a frame consisting of areas not

under the jurisdiction of permit-issuing offices (i.e., the 1980-based nonpermit universe). The first step in the sampling operation for the nonpermit universe was the selection of a sample of census ED's within these areas (using the overall sampling rate). Prior to this sample selection, the ED's were sorted by State, district office and enumeration district number. The probability of selection of an ED was proportionate to the following:

$$\frac{\text{Number of housing units in 1980 census ED} + \text{Noninstitutionalized group quarters population in 1980 census ED}}{4} = 2.75$$

The sample ED's were then divided into segments (i.e., small land areas with well-defined boundaries having an expected size of four, or a multiple of four, housing units). At the time of the survey, those segments that did not have an expected size of four housing units were further subdivided to produce an expected four sample housing units. Following the division, a segment from each sample ED was selected. All housing units in existence at the time of interview in these selected segments were eligible for sample. Thus, housing units enumerated in the 1980 census as well as housing units built since the 1980 census are included.

**AHS-National Sample Selection.** This sample was set up as a multistage design in which the United States was divided into areas made up of counties and independent cities called primary sampling units (PSU's). These PSU's were grouped into strata consisting of one or more PSU's and then one PSU was selected from each stratum to represent all PSU's in that stratum. Sample units were selected from 1980 census units in these PSU's at an overall sampling rate of about 1 in 2,148. In areas where addresses were, for the most part, complete, and where new construction is monitored by permits, a sample of housing units which received long-form questionnaires in the 1980 census was selected directly from a list of all such housing units based on certain housing and geographic information of housing unit.

In areas where at least 4 percent of the addresses were incomplete or inadequate, or where new construction was not monitored by building permits (most rural areas), a sample of 1980 census units which received long-form questionnaires was selected in several steps.

The sample of permit new construction was selected from issued building permits such that the units are expected to be completed after April 1, 1980. The sampling procedure was similar to that of AHS-MS; however, the subsampling rate used was 1 in 4.

Housing units added to the inventory since the 1980 census were represented using two methods. One method identified within structure additions. These are units in structures which had a chance of being in sample because they contained at least one unit enumerated in the 1980

census. The other method identified whole structure additions. These are units in structures for which none of the units in the structure were enumerated in the 1980 census.

Additional information concerning the 1985 AHS-National survey is available in the current housing report series H-150-85.

**ESTIMATION**

The 1985 AHS-Metropolitan Area sample produced estimates pertaining to characteristics of the housing inventory at the time of the interview (i.e., the 1985 housing inventory). The combined estimates used information from both the AHS-MS and AHS-National samples (i.e., the combined sample estimates).

**AHS-MS.** Prior to performing estimation procedures using the combined sample, the AHS-MS sample housing units were weighted according to a one-stage ratio estimation procedure. Before the implementation of the ratio estimation procedure, the basic weight (i.e., the inverse of the probability of selection) for each interviewed sample housing unit was adjusted to account for Type M and Type A noninterviews.

**Type M Noninterview Adjustment.** The Type M noninterviews are sample units which were dropped due to selection by another survey. These noninterviews occur only in the 1980-based permit-issuing area universe, the 1980-based nonpermit-issuing area universe, and the 1980-based new construction universe. This adjustment was done separately for the 1980 central city and balance of each metropolitan area for housing units in the 1980-based permit-issuing universe, in group quarters, in the 1980-based nonpermit-issuing area universe, and in the 1980-based new construction universe. The Type M noninterview adjustment was computed separately for each cell and was equal to the following:

$$\frac{\text{AHS-MS sample estimate of 1980 housing units in the cell} + \text{Weighted count of Type M noninterviewed housing units}}{\text{AHS-MS sample estimate of 1980 housing units in the cell}}$$

**Type A Noninterview Adjustment.** The next adjustment was the Type A noninterview adjustment. This adjustment was done on occupied housing units and was computed separately for units in the 1980-based permit-issuing area universe, for new construction, and for all other housing units (this includes the 1970-based permit-issuing universe, the 1970-based and 1980-based nonpermit-issuing universes and the 1970-based new construction housing units built prior to the last survey). For units in the 1980-based permit-issuing universe, a Type A noninterview adjustment factor was computed separately for each of the 62 strata used in the sample selection process by 1980 central city and balance. For new construction units, a Type A noninterview adjustment factor was computed

separately by tenure for each of the central city and balance. For all other units, a Type A noninterview adjustment factor was calculated separately by tenure and 1970 central city and balance for each of the following: (1) 24 noninterview cells for sample housing units from the permit-issuing universe (where the cell consisted of one or more of the 50 different strata used in the 1970-based permit-issuing universe as previously described); (2) one noninterview cell for new construction housing units; (3) one noninterview cell for mobile homes or trailers from the nonpermit-issuing universe; (4) one noninterview cell for units that were not mobile homes or trailers from the nonpermit-issuing universe; (5) three noninterview cells for units from the coverage improvement universe; (6) one noninterview cell for units classified as vacants at the time of the 1970 Census; and (7) one noninterview cell for units classified as group quarters at the time of the 1970 Census. Within a given cell, the Type A noninterview adjustment factor was equal to the following ratio using the basic weight times the Type M noninterview adjustment factor for the sample weight:

$$\frac{\text{Weighted count of interviewed housing units} + \text{Weighted count of Type A noninterviewed housing units}}{\text{Weighted count of interviewed housing units}}$$

**AHS-MS Ratio Estimation Procedure.** The following ratio estimation procedure was employed for all sample housing units from the permit issuing universe. This factor was computed separately for all sample housing units within each 1970-based permit-issuing universe noninterview cell mentioned previously. The ratio estimation factor for each cell was equal to the following:

$$\frac{\text{1970 census count of housing units from the 1970-based permit-issuing universe in the corresponding cell}}{\text{AHS-MS sample estimate of 1970 housing units from the permit-issuing universe in the corresponding cell}}$$

For each metropolitan area, the numerators of the ratios were obtained from the 1970 Census of Population and Housing 20-percent file of housing units enumerated in areas under the jurisdiction of permit-issuing offices.

The denominators of the ratio estimation factors were then computed and were obtained from weighted estimates of all the AHS-MS sample housing units from the 1970-based permit-issuing universe within the corresponding ratio estimation categories using the existing weight (i.e., the basic weight times the Type A noninterview adjustment). The computed ratio estimation factor was then applied to the existing weight for each sample housing unit within the corresponding ratio estimation cells. This ratio estimation procedure was introduced to correct the probabilities of selection for samples in each of the strata used in the sample selection of the 1970-based permit-issuing universe. Prior to the AHS-MS sample selection

within each metropolitan area, housing units already selected for other Census Bureau surveys were deleted from the permit-issuing universe. The same probability of selection was then applied to the remaining units to select the AHS-MS sample. Since the number of housing units deleted from the AHS-MS universe frame was not necessarily proportional among all strata, some variation in the actual probability of selection between strata was introduced during the sample selection process.

The following ratio estimation procedure was employed for all sample units from the 1980-based permit-issuing universe. This factor was computed separately for all metropolitan areas excluding the Dallas, TX, PMSA; Los Angeles-Long Beach, CA, PMSA; Philadelphia, PA-NJ, PMSA; Phoenix, AZ, MSA; and San Francisco-Oakland, CA, area PMSA's within each 1980-based permit-issuing universe noninterview cell mentioned previously. The ratio estimation factor was equal to the following:

$$\frac{\text{1980 census count of housing units from the 1980 permit-issuing universe in the corresponding cell}}{\text{AHS-MS sample estimate of 1980 housing units from the 1980 permit-issuing universe in the corresponding cell}}$$

For each metropolitan area, the numerator of the ratio was obtained from the 1980 Census of Population and Housing 100-percent file of housing units enumerated in areas under the jurisdiction of permit issuing offices. The denominator of the ratio was obtained from weighted estimates of all the AHS-MS sample housing units within the corresponding ratio estimation categories using the existing weight (i.e., the basic weight times the Type M noninterview adjustment factor times the Type A noninterview adjustment factor).

The computed ratio estimation factor was then applied to the existing weight for each sample housing unit within the corresponding ratio estimation categories.

This ratio estimation procedure was introduced to adjust the sample estimate in each of the strata used in the sample selection of the 1980-based permit issuing universe to an independent estimate (1980 census count) for the strata. This adjustment was necessary since some sample units were dropped during the processing procedures.

**AHS-National.** Before implementing estimation procedures using the AHS-National units for the combined sample, the AHS-National sample units were assigned a weight which reflected the probability of selection for the unit. The AHS-National weighting procedure then made adjustments for units which could not be interviewed for a variety of reasons. For each of these adjustments, a factor was computed and applied to the appropriate units.

The first of these adjustments was done in permit segments only, to account for permits which could not be sampled and units which could not be located. These were

represented by all other units in permit segments including both interviews and noninterviews (excluding unable-to-locate noninterviews).

The second of the adjustments was done for units in structures built before April 1, 1980. It was done to account for units which could not be located. The unlocatable units were represented by both interviews and noninterviews (excluding unable-to-locate noninterviews).

The last of these adjustments was done to account for units which could not be interviewed because either no one was home after repeated visits or the respondent refused to be interviewed. When 1980 census data was available, this information was used to determine the noninterview adjustment cell. The cells included characteristics such as tenure, geography, units in structure, and number of rooms. When 1980 census data was not available, adjustment factors were computed separately using more general characteristics such as type of area and type of housing unit (i.e., mobile home, nonmobile home). Additional information on the AHS-National weighting procedure can be found in the current housing reports H-150-85 series.

## COMBINED SAMPLE WEIGHTING

**Introduction.** The estimates for the combined sample were obtained by summing the sample weights of interviewed AHS-MS and AHS-National units. For AHS-MS sample units, the starting weight was obtained after the AHS-MS ratio estimation procedure. For AHS-National units, the starting weight was obtained after the Type A noninterview adjustment. In order to account for the use of two different samples representing one metropolitan area, weighting factors were assigned to each unit prior to the combined sample ratio estimation procedures.

**Weighting Factor Adjustment.** The weighting factor adjustment was computed separately for each metropolitan area by sample design (AHS-MS or AHS-National) according to "new construction" or "old construction" classification. New construction was defined as units built in permit-issuing areas since the 1980 census; old construction units were then categorized by tenure classification (renter/owner).

For a given characteristic, the AHS-MS weighting factor adjustment was a function of the sample size in each survey and the variance associated with each survey's estimates.

The corresponding weighting factor was then applied to the existing weight of each AHS-MS and AHS-National sample unit, and the weights were then combined according to characteristic (i.e., AHS-MS new construction + AHS-National new construction, etc.).

**Combined Sample Ratio Estimation Procedures.** The following ratio estimation procedure was applied only in the Los Angeles-Long Beach, CA, PMSA. It involved the

ratio estimation of the weighted sample of new construction units in the metropolitan area to an independent estimate of units in the metropolitan area completed during the same time period. This ratio estimation factor was equal to the following:

$$\frac{\text{Independent estimate of housing units completed in November 1980 or later}}{\text{Sample estimate of housing units completed in November 1980 or later}}$$

The numerator of this ratio was determined using Survey of Construction data. The denominator of this ratio was obtained from the weighted estimate of the AHS sample housing units using the existing weight (i.e., the starting weight times the combined sample weighting factor).

The computed ratio estimate factor was then applied to the existing weight for all units in the metropolitan area completed in November 1980 or later.

The next ratio estimation procedure was applied only in the Phoenix, AZ, MSA. This procedure involved the ratio estimation of the weighted sample estimate of occupied mobile homes in the metropolitan area to an independent estimate of occupied mobile homes in the metropolitan area. This factor is given by the following:

$$\frac{\text{Independent estimate of occupied mobile homes in the Phoenix, AZ, MSA}}{\text{Sample estimate of occupied mobile homes in the Phoenix, AZ, MSA}}$$

The numerator of this ratio was determined using data from the 1980 census and the 1985 Special Census for the Phoenix, AZ, MSA. The denominator was obtained using the existing weight of AHS sample units (i.e., the starting weight times the combined sample weighting factor).

The computed ratio estimate factor was then applied to the existing weight for all interviewed mobile homes in the metropolitan area.

The next ratio estimation procedure was applied to all metropolitan areas except the Phoenix, AZ, MSA. Each metropolitan area was subdivided into geographic areas consisting of a combination of counties or minor civil divisions. The ratio estimation procedure involved the ratio estimation of the weighted sample estimate of the August 1, 1985, housing inventory in each geographic area for each metropolitan area to an independent estimate of occupied housing units for the corresponding cell. This ratio estimation factor equalled the following:

$$\frac{\text{Independent estimate of the August 1, 1985, occupied housing unit inventory for the corresponding geographic area of the metropolitan area}}{\text{AHS-metropolitan area sample estimate of the occupied housing inventory for the corresponding geographic area of the metropolitan area}}$$

The independent estimates of occupied housing units that were used as the numerator of this ratio are described

below. The denominator of this ratio was obtained from the weighted estimate of the occupied AHS sample housing units using the existing weight.

Independent estimates were derived for the August 1, 1985 occupied housing inventory for each geographic area within each metropolitan area. For all metropolitan areas excluding the Boston, MA-NH, CMSA; Los Angeles-Long Beach, CA, PMSA; and Phoenix, AZ, MSA, the estimates were based on the following ratio:

$$\frac{\text{1985 estimate of population (age 15+ )}}{\text{excluding group quarters in the county}}$$


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$$\frac{\text{1985 estimate of population (age 15+) per household}}{\text{excluding group quarters in the county}}$$

For the Boston, MA-NH, CMSA, the estimate was based on the following ratio:

$$\frac{\text{1985 estimate of total population}}{\text{excluding group quarters in the minor civil division}}$$


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$$\frac{\text{1985 estimate of total population per household}}{\text{excluding group quarters in the minor civil division}}$$

The methodology used to derive the independent estimates for all metropolitan areas excluding the Los Angeles-Long Beach, CA, PMSA and Phoenix, AZ, MSA was based on the population-per-household method as described in the *Proceedings of the Bureau of the Census Second Annual Research Conference, March 23-26, 1986*, pages 83-110. This method is based on the national trend of the adult population per household and assumes that this trend is uniform throughout the country.

For the Los Angeles-Long Beach, CA, PMSA, the independent estimates were obtained from the State of California, Department of Finance. In this metropolitan area, the population-per-household could not be applied since the national population-per-household trend under estimated the true population per household in this metropolitan area.

The AHS sample estimate of the housing inventory for the corresponding geographic area was obtained using the existing weight.

The computed ratio estimation factors were then applied to all housing units (including vacant units) in the corresponding geographic area of each metropolitan area and the resulting product was used as the final weight for tabulation purposes.

The following ratio estimation procedure was applied only in the Phoenix, AZ, MSA. This procedure involved the ratio estimation of the weighted sample of all combined sample units in the metropolitan area to an independent estimate of all units in the metropolitan area. This factor is equal to the following:

$$\frac{\text{Independent estimate of the total housing inventory}}{\text{in the Phoenix, AZ, MSA}}$$


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$$\frac{\text{AHS combined sample estimate of the total housing inventory}}{\text{in the Phoenix, AZ, MSA}}$$

The numerator of this ratio was determined using data from the 1985 Special Census of the Phoenix, AZ, MSA. The denominator was obtained from the weighted estimate of AHS combined sample units using the existing weight.

The computed ratio estimate factor was then applied to the existing weight for all combined sample units in the metropolitan area.

The effect of this ratio estimation procedure, as well as the overall estimation procedures, was to reduce the sampling error for most statistics below what would have been obtained by simply weighting the results of the sample by the inverse of the probability of selection. Since the housing population of the sample differed somewhat, by chance, from the metropolitan area as a whole, it can be expected that the sample estimates will be improved when the sample housing population, or different portions of it, is brought into agreement with known good estimates of the metropolitan area housing population.

## ACCURACY OF THE ESTIMATES

There are two types of possible errors associated with estimates based on data from sample surveys—sampling and nonsampling errors. The following is a description of the sampling and nonsampling errors associated with the AHS sample estimates.

**Nonsampling Errors.** In general, nonsampling errors can be attributed to many sources: inability to obtain information about all cases, definitional difficulties; differences in the interpretation of questions; inability or unwillingness of respondents to provide correct information; mistakes in recording or coding the data; other errors of collection, response, processing, and coverage; and estimation for missing data. Nonsampling errors are not unique to sample surveys since they can, and do, occur in complete censuses as well.

Obtaining a measurement of the total nonsampling error associated with the estimates from a survey is very difficult, considering the number of possible sources of error. However, an attempt was made to measure some of the nonsampling errors associated with the estimates for the 1985 AHS-Metropolitan Area sample.

**AHS-MS Content Errors.** A content reinterview program was done for the 1985 AHS-Metropolitan Area sample units. A sample of these units was revisited and answers to some of the questions on the questionnaire were obtained again. The original interview and reinterview were assumed to be two independent readings and thus were the basis for the measurement of the accuracy of the data collected from interviewed units.

The 1985 content reinterview program served as an interviewer quality check and a quality analysis of particular survey questions. One-fourth of all interviewers were

selected for the quality check which reviewed the interviewers' proficiency in properly evaluating the items listed below. The other portion of the reinterview program was performed to ensure that certain survey questions elicited consistent responses from the interviewed households. These reinterview items and their response variability are discussed below.

The six interviewer items reviewed were: (1) correct unit visited; (2) area segment coverage; (3) living quarters classification; (4) tenure; (5) interview status; and (6) household composition.

The AHS-MS survey items reviewed generally fell into three categories: (1) major repairs; (2) mortgage; and (3) mobility. In the category of major repairs, all but one item showed moderate response variability; "major repairs over \$500 each" had high variability. Of the 11 measurable items in the mortgage category, 7 had low variability, 3 had moderate variability, and "payments the same throughout mortgage" showed high response variability. The mobility category had four items which showed moderate response variability, and two which showed high response variability: "area lived at age 16" and "preferred place to live in 5 years." One final item which had high variability was the size of the lot. It was found that most people did not know their lot size either in square feet, feet by feet, or whole acres.

Low levels of inconsistency indicate that the response error is insignificant relative to the standard error in this report. Moderate levels of inconsistency indicate that the response error is not insignificant compared to the standard error in this report. High levels of inconsistency indicate that the response error is very significant compared to the standard error in this report, and caution should be used when examining estimates of these characteristics.

In this publication, cross-tabulations involving those items which are subject to high levels of inconsistency may also be subject to a large distortion as a consequence and thus are considered to be less reliable than comparable cross-tabulations which do not involve these data. Since the reinterview programs only measured inconsistencies for a sample of the items on the AHS questionnaire, there may be other items with high levels of inconsistency.

For additional information on the content reinterview program, refer to the Census Bureau memorandum, "1985 AHS-MS Reinterview Results."

**AHS-National Content Errors.** A content reinterview program was conducted for the AHS-National households as well. A subsample of the original households was revisited, and certain questions from the original questionnaire were asked again. The original and reinterview were assumed to be two independent readings and thus were the basis for the measurement of the response error of these AHS estimates. The reinterview also served as a check for interviewer evaluation and quality control. The

AHS-National reinterview program performed an interviewer quality check using questions similar to those described above. The reinterview study for survey questions was done for three groups of items. They are units in structure and description of structure, number and type of rooms, and appliances including the age and fuel of the appliances. For reinterview results, refer to the current housing reports H-150-85 series.

Reinterview studies were also conducted in conjunction with previous AHS-National and AHS-MS enumerations. These studies included items dealing primarily with poor housing quality, attitudes about the neighborhood, and certain housing costs. The following table shows the items which had higher levels of inconsistency. While these questions were not included in the 1985 reinterview studies, questions from previous enumerations were not altered enough to lead one to believe that the level of inconsistent responses would change.

Survey items	Level of Inconsistency
Open cracks or holes on inside of building	Moderate to High
Holes in floors	Moderate to High
Broken plaster or peeling paint on ceilings and walls	High
Mice or rats	Moderate
Working electric outlet in all rooms	High
Concealed wiring	High
Blown fuses/tripped circuit breakers	Moderate to High
Neighborhood conditions: street noise; roads in need of repair; crime; trash, litter, junk in streets or on properties; boarded-up/abandoned structures; nonresidential activities; odors, smoke, gas	Moderate to High
Satisfactory neighborhood services: police protection; hospitals/health clinics; public transportation; shopping; elementary schools	Moderate to High
Electricity cost	Moderate to High
Gas cost	High
Oil, coal, kerosene, wood or other fuel cost	Moderate to High
Fire/hazard insurance	Moderate to High
Real estate taxes	Moderate
Cost of real estate taxes	Moderate
Cost of water supply and sewage disposal	High
Cost of garbage collection	Moderate to High
Gross income	High
Prefer to live in same area or somewhere else	Moderate

A possible explanation for the results of the reinterview studies, as well as the surveys themselves, is that respondents may lack precise information. Also, since the results of the reinterview studies are derived from sample surveys, there is sampling error associated with these estimates of nonsampling error. The possibility of such errors should be taken into account when considering the results of these studies.

**Coverage Errors.** In errors of coverage and estimation for missing data, the AHS new construction sample had deficiencies in the representation of conventional (nonmobile home or trailer) new construction, except for the Los

Angeles-Long Beach, CA, PMSA, which was adjusted to an independent estimate. Due to time constraints, only those building permits issued more than 7 months before the survey ended were eligible to be sampled to represent conventional new construction in permit-issuing areas for this metropolitan area. However, those permits issued during the last 7 months of the survey do not necessarily represent missed housing units. Due to the relatively short time span involved, it is possible that construction of these housing units was not completed at the time the survey was conducted, in which case, they would not have been eligible for interview. In addition to these deficiencies, new construction in special places that do not require building permits, such as military bases, is not adequately presented.

AHS misses a significant portion of new mobile homes (i.e., those placed after January 1, 1980). It is believed that most of the difference is due to poor coverage of new mobile home parks in address ED's. The AHS-National survey estimates that as much as 25 percent of new mobile homes are missed due to poor coverage. Note that the mobile home ratio estimation procedure used in the Phoenix, AZ, MSA was an attempt to correct for this deficiency.

Deficiencies also exist in ED's where area sampling methods are used. It had been assumed that all housing units located inside these ED's would be represented in the sample. However, it has been established that the AHS sample missed as much as 2 percent of all housing units in these ED's because they were not listed during the canvassing. It should be noted that since these ED's were recanvassed each time this metropolitan area was surveyed, the number of missed housing units may be considerably less for the 1985 survey.

The final ratio estimation procedure corrects for these deficiencies as far as the count of total housing is concerned (i.e., it adjusts to the best available estimate). However, biases of subtotals would still remain.

**Rounding Errors.** For errors associated with processing, rounding of estimates introduces another source of error in the data, the severity of which depends upon the statistics being measured. The effect of rounding is significant relative to the sampling error only for small percentages or small medians, when these figures are derived from relatively large bases (e.g., median number of persons per household). This means that confidence intervals formed from the standard errors given may be distorted, and this should be taken into account when considering the results of the survey.

**Sampling Errors for the AHS Combined Sample Estimates.** The particular sample used for this survey is one of a large number of possible samples of the same size that could have been selected using the same sample design. Even if the same questionnaires, instructions, and

interviewers were used, estimates from each of the different samples would differ from one another. The sampling error of a survey estimate provides a measure of the variation among the estimates from all possible samples, and thus is a measure of the precision with which an estimate from a sample approximates the average result of all possible samples.

One common measure of the sampling error is the standard error. As calculated for this report, the standard error reflects the variation in the estimates due to sampling and nonsampling errors, but it does not measure as such any systematic biases in the data. Therefore, the accuracy of the estimates depends upon the standard error, biases, and any additional nonsampling errors not measured by the standard error. The sample estimate and its estimated standard error enable one to construct interval estimates in which the interval includes the average result of all possible samples with a known probability. For example, if all possible samples were selected, each of these surveyed under essentially the same general conditions, and an estimate and its estimated standard error were calculated from each sample, then:

Approximately 90 percent of the intervals from 1.6 standard errors below the estimate to 1.6 standard errors above the estimate would include the average result of all possible samples.

The average result of all possible samples either is or is not contained in any particular computed interval. However, for a particular sample, one can say with specified confidence that the average result of all possible samples is included in the constructed interval.

The figures presented in the tables that follow (see page App-68) are approximations to the standard errors of various estimates shown in this report for this metropolitan area. In order to derive standard errors that would be applicable to a wide variety of items and also could be prepared at a moderate cost, a number of approximations were required. As a result, the tables of standard errors provide an indication of the order of magnitude of the standard errors rather than precise standard errors for any specific item.

Table I presents the standard errors applicable to estimates of characteristics of the 1985 housing inventory. Linear interpolation should be used to determine the standard errors for estimates not specifically shown in this table.

The reliability of an estimated percentage, computed by using the sample data for both numerator and denominator, depends upon both the size of the percentage and the size of the total upon which the percentage is based. Estimated percentages are relatively more reliable than the corresponding estimates of the numerators of the percentages, particularly if the percentages are 50 percent or more.

Table II presents the standard errors of estimated percentages for the 1985 housing inventory. Two-way interpolation should be used to determine standard errors for estimated percentages not specifically shown in table II.

Included in tables I and II are estimates of standard errors for estimates of zero and zero percent. These estimates of standard errors are considered as overestimates of the true standard errors and should be used primarily for construction of confidence intervals for characteristics when estimates of zero are obtained.

For ratios,  $100(x/y)$ , where  $x$  is not a subclass of  $y$ , table I underestimates the standard error of the ratio when there is little or no correlation between  $x$  and  $y$ . For this type of ratio, a better approximation of the standard error may be obtained by letting the standard error of the ratio be approximately equal to:

$$(100) \frac{x}{y} \sqrt{\left(\frac{s_x}{x}\right)^2 + \left(\frac{s_y}{y}\right)^2}$$

- where:  $x$  = the numerator of the ratio
- $y$  = the denominator of the ratio
- $s_x$  = the standard error of the numerator
- $s_y$  = the standard error of the denominator

**Illustration of the Use of the Standard Error Tables.**

Table 2-1 of this report shows that in the Boston, MA-NH, metropolitan area, there were 864,200 owner-occupied housing units. Interpolation using table I of this appendix shows that the standard error of an estimate of this size is approximately 12,760. The following interpolation procedure was used.

The information presented in the following table was extracted from table I. The entry for "x" is the one sought.

Size of estimate	Standard error
800,000 .....	12,820
864,200 .....	x
900,000 .....	12,720

The entry for "x" is determined as follows by vertically interpolating between 12,820 and 12,720.

$$864,200 - 800,000 = 64,200$$

$$900,000 - 800,000 = 100,000$$

$$12,820 + \frac{64,200}{100,000}(12,720 - 12,820) = 12,760$$

Consequently, the 90-percent confidence interval, as shown by these data, is from 843,780 to 884,620 housing units. Therefore, a conclusion that the average estimate, derived from all possible samples, of 1985 owner-occupied housing units lies within a range computed in this way would be correct for roughly 90 percent of all possible samples.

Table 2-3 also shows that of the 864,200 owner-occupied housing units, 206,500 or 23.9 percent had two bedrooms. Interpolation using table II of this appendix (i.e., interpolation on both the base and percent) and applying the factor given in the footnote shows that the standard error of the 23.9 percent is approximately 0.9 percentage points. The following interpolation procedure was used.

The information presented in the following table was extracted from table II.

Base of percentage	Estimated percentage		
	10 or 90	23.9	25 or 75
800,000 .....	0.7	a	1.0
864,200 .....		p	
900,000 .....	0.6	b	0.9

1. The entry for cell "a" is determined by horizontal interpolation between 0.7 and 1.0.

$$23.9 - 10.0 = 13.9$$

$$25.0 - 10.0 = 15.0$$

$$0.7 + \frac{13.9}{15.0}(1.0 - 0.7) = 1.0$$

2. The entry for cell "b" is determined by horizontal interpolation between 0.6 and 0.9.

$$23.9 - 10.0 = 13.9$$

$$25.0 - 10.0 = 15.0$$

$$0.6 + \frac{13.9}{15.0}(0.9 - 0.6) = 0.9$$

3. The entry for "p" is then determined by vertical interpolation between 1.0 and 0.9

$$864,200 - 800,000 = 64,200$$

$$900,000 - 800,000 = 100,000$$

$$1.0 + \frac{64,200}{100,000}(0.9 - 1.0) = 0.9$$

Applying a factor of 1.0 according to the footnote from table II gives a standard error of 0.9 percentage points. Consequently, the 90-percent confidence interval, as shown by these data, is from 22.5 to 25.3 percent.

**Differences.** The standard errors shown are not directly applicable to differences between two sample estimates. The standard error of a difference between estimates is approximately equal to the square root of the sum of the squares of the standard error of each estimate considered separately. This formula is quite accurate for the difference between estimates of the same characteristics in two different metropolitan areas or the difference between separate and uncorrelated characteristics in the same metropolitan area. If there is a high positive correlation

between the two characteristics, the formula will overestimate the true standard error; but if there is a high negative correlation, the formula will underestimate the true standard error.

**Illustration of the Computation of the Standard Error of a Difference.** Table 2-3 of this report shows that in the Boston, MA-NH, metropolitan area, there were 398,500 owner-occupied housing units with three bedrooms. Thus, the apparent difference, as shown by these data, between owner-occupied housing units with two bedrooms and owner-occupied housing units with three bedrooms is 192,000. Table I shows that the standard error of 206,500 is approximately 8,600 and the standard error of 398,500 is approximately 11,100. Therefore, the standard error of the estimated difference of 192,000 is about 14,040.

$$14,040 = \sqrt{(8,600)^2 + (11,100)^2}$$

Consequently, the 90-percent confidence interval for the 192,000 difference is from 169,540 to 214,460 housing units. Therefore, a conclusion that the average estimate derived from all possible samples, of this difference, lies within a range computed in this way would be correct for roughly 90 percent of all possible samples. Thus, we can conclude with 90-percent confidence that the number of 1985 owner-occupied housing units with three bedrooms is greater than the number of owner-occupied units with two bedrooms since the 90-percent confidence interval does not include zero or negative values.

**Medians.** For medians presented in certain tables, the sampling error depends on the size of the base and on the distribution upon which the median is based. An approximate method for measuring the reliability of the estimated median is to determine an interval about the estimated median so that there is a stated degree of confidence that the average median from all possible samples lies within the interval. The following procedure may be used to estimate confidence limits of a median based on sample data:

1. From table II, determine the standard error of a 50-percent characteristic on the base of the median.
2. Add to and subtract from 50 percent, the standard error determined in step 1.
3. Using the distribution of the characteristics, determine the confidence interval corresponding to the two points established in step 2. To find the lower endpoint of the confidence interval, it is necessary to know into which interval of the distribution the lower percentage limit

falls. Similarly, to find the upper endpoint of the confidence interval, it is necessary to know into which interval of the distribution the upper percentage limit falls. Note that these distribution intervals could be different, although this will not happen very often.

A 1.6 standard-error confidence interval may be determined by finding the values corresponding to 50 percent plus and minus 1.6 times the standard error determined in step 1. For about 90 out of 100 possible samples, the average median from all possible samples would lie between these two values.

**Illustration of the Computation of the 90-Percent Confidence Interval of a Median.** Table 2-9 of this report shows the median number of persons for owner-occupied housing units is 2.8. The base of the distribution from which this median was determined is 864,200 housing units.

1. Interpolation using table II and the applied factor shows that the standard error of 50 percent on a base of 864,200 is approximately 1.1 percentage points.
2. To obtain a 90-percent confidence interval on the estimated median, initially add to and subtract from 50 percent 1.6 times the standard error determined in step 1. This yields percentage limits of 48.2 and 51.8.
3. From the distribution for "persons" in table 2-9, the interval for owner-occupied housing units with three persons (for purposes of calculating the median, the category of three persons is considered to be from 2.5 to 3.5 persons) contains the 48.2 percent derived in step 2. About 388,700 housing units or 45.0 percent fall below this interval, and 165,200 housing units or 19.1 percent fall within this interval. By linear interpolation, the lower limit of the 90-percent confidence interval is found to be about 2.7.

$$2.5 + (3.5 - 2.5) \frac{48.2 - 45.0}{19.1} = 2.7$$

Similarly, the interval for owner-occupied housing units with three persons contains the 51.8 percent derived in step 2. About 388,700 housing units or 45.0 percent fall below this interval, and 165,200 housing units or 19.1 percent fall within this interval. The upper limit of the 90-percent confidence interval is found to be about 2.9.

$$2.5 + (3.5 - 2.5) \frac{51.8 - 45.0}{19.1} = 2.9$$

Thus, the 90-percent confidence interval ranges from 2.7 to 2.9 persons.

Table I. Standard Errors for Estimated Number of Housing Units in the 1985 Boston, MA-NH, CMSA

Size of estimate	Standard error <sup>1</sup>		
	Combined owner and renter housing units <sup>2</sup>	Owner housing units <sup>3</sup>	Renter housing units <sup>4</sup>
0.....	420	410	410
500.....	460	450	450
700.....	540	540	530
1,000.....	650	640	640
2,500.....	1,020	1,010	1,010
5,000.....	1,440	1,430	1,420
10,000.....	2,030	2,020	2,010
25,000.....	3,200	3,190	3,160
50,000.....	4,490	4,470	4,430
75,000.....	5,460	5,430	5,380
100,000.....	6,250	6,220	6,160
150,000.....	7,530	7,490	7,420
200,000.....	8,540	8,490	8,420
250,000.....	9,370	9,320	9,240
300,000.....	10,070	10,020	9,940
400,000.....	11,170	11,120	11,020
500,000.....	11,960	11,900	11,800
600,000.....	12,490	12,420	12,320
700,000.....	12,790	12,730	12,620
800,000.....	12,890	12,820	-
900,000.....	12,780	12,720	-
1,000,000.....	12,470	-	-
1,100,000.....	11,920	-	-
1,200,000.....	11,120	-	-
1,300,000.....	10,000	-	-
1,400,000.....	8,440	-	-
1,500,000.....	6,090	-	-
1,594,620.....	-	-	-

<sup>1</sup>To compute standard errors for new construction estimates, the standard errors in the table should be multiplied by a factor of 1.1 for owner housing units, 1.1 for renter housing units, and 1.1 for the combined owner and renter housing units.

<sup>2</sup>Some examples that pertain to both owner and renter housing units are: total housing units; all occupied housing units; all year-round housing units; mobile homes or trailers; and total vacant housing units.

<sup>3</sup>The owner housing units pertain to owner-occupied housing units and vacant housing units excluding vacant-for-rent housing units.

<sup>4</sup>The renter housing units pertain to renter-occupied housing units and vacant-for-rent housing units.

**Table II. Standard Errors for Estimated Percentages of Housing Units in the 1985 Housing Inventory of the Boston, MA-NH, CMSA**

Base of percentage	Estimated percentage <sup>1</sup>					
	0 or 100	1 or 99	5 or 95	10 or 90	25 or 75	50
500 .....	44.8	44.8	44.8	44.8	44.8	45.0
700 .....	36.7	36.7	36.7	36.7	36.7	38.1
1,000 .....	28.8	28.8	28.8	28.8	28.8	31.8
2,500 .....	14.0	14.0	14.0	14.0	17.4	20.1
5,000 .....	7.5	7.5	7.5	8.5	12.3	14.2
10,000 .....	3.9	3.9	4.4	6.0	8.7	10.1
25,000 .....	1.6	1.6	2.8	3.8	5.5	6.4
50,000 .....	0.8	0.9	2.0	2.7	3.9	4.5
75,000 .....	0.5	0.7	1.6	2.2	3.2	3.7
100,000 .....	0.4	0.6	1.4	1.9	2.8	3.2
150,000 .....	0.3	0.5	1.1	1.6	2.3	2.6
200,000 .....	0.20	0.4	1.0	1.4	1.9	2.3
250,000 .....	0.16	0.4	0.9	1.2	1.7	2.0
300,000 .....	0.13	0.4	0.8	1.1	1.6	1.8
400,000 .....	0.10	0.3	0.7	1.0	1.4	1.6
500,000 .....	0.08	0.3	0.6	0.9	1.2	1.4
600,000 .....	0.07	0.3	0.6	0.8	1.1	1.3
700,000 .....	0.06	0.2	0.5	0.7	1.0	1.2
800,000 .....	0.05	0.2	0.5	0.7	1.0	1.1
900,000 .....	0.05	0.2	0.5	0.6	0.9	1.1
1,000,000 .....	0.04	0.2	0.4	0.6	0.9	1.0
1,100,000 .....	0.04	0.2	0.4	0.6	0.8	1.0
1,200,000 .....	0.03	0.18	0.4	0.6	0.8	0.9
1,300,000 .....	0.03	0.18	0.4	0.5	0.8	0.9
1,400,000 .....	0.03	0.17	0.4	0.5	0.7	0.9
1,500,000 .....	0.03	0.16	0.4	0.5	0.7	0.8
1,594,620 .....	0.03	0.16	0.3	0.5	0.7	0.8

<sup>1</sup>Standard errors are presented to the nearest one-tenth of one percentage point except when the standard error is less than or equal to fifteen-hundredths of one percentage point; in those cases, the standard error is shown to the nearest one-hundredth of one percentage point. For estimates pertaining to new construction, the standard errors shown in the table should be multiplied by a factor of 1.1

The following factors should be applied to estimates that do not pertain strictly to new construction. For estimates pertaining to both owners and renters, apply a factor of 1.0. For estimates pertaining to owner housing units, apply a factor of 1.0. For estimates pertaining to renter housing units, apply a factor of 1.0.